

A Note on Watson's Paper

by

M. S. Patel

Purdue University

Department of Statistics

Mimeograph Series Number 6

April, 1963

Mimeograph Series No. 6
 May, 1963

A Note on Watson's Paper

M. S. Patel

Purdue University

Watson in his paper [1] mentions on page 381 that "It is hard to be sure than an increase in α increases $E(M_R)$ but the lower bound to $E(M_R)$ certainly increases." In this short note, the author intends to remove some kind of doubt left in the first part of the statement. Following the notations in [1], we have

$$E(M_R) = kp' \sum_{n=0}^g n \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) \binom{g}{n} \pi_1^{*n} (1-\pi_1^*)^{g-n} \quad (1)$$

for shortness, we shall write $b(n, \pi_1^*, g)$ to denote $\binom{g}{n} \pi_1^{*n} (1-\pi_1^*)^{g-n}$.

Differentiating (1) w.r.t. α , we get

$$\frac{\partial}{\partial \alpha} E(M_R) = k \frac{dp'}{d\alpha} \sum_{n=0}^g n \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) b(n, \pi_1^*, g)$$

$$+ kp' \sum_{n=0}^g (n-\pi_1^*) \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) \frac{\partial \pi_1^*}{\partial \alpha} b(n-1, \pi_1^*, g-1) / (1-\pi_1^*) \quad (2)$$

With slight adjustment, (2) becomes

$$\begin{aligned}
 \frac{\partial}{\partial \alpha} E(M_R) &= f \pi_1^* \frac{dp'}{d\alpha} \sum_{n=1}^g \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) b(n-1, \pi_1^*, g-1) \\
 &+ fp' \frac{d\pi_1^*}{d\alpha} \sum_{n=1}^g \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) b(n-1, \pi_1^*, g-1) \\
 &+ fp' \frac{d\pi_1^*}{d\alpha} \sum_{n=1}^g \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) \left[(n-1) - (g-1) \pi_1^* \right] \\
 &\quad \times b(n-1, \pi_1^*, g-1) / (1 - \pi_1^*) . \tag{3}
 \end{aligned}$$

Next, $p' = p\pi_1'/\pi_1^*$ [cf. (4.14), page 379, [1]]. Therefore,

$$\frac{d}{d\alpha} (\pi_1^* p') = \pi_1^* \frac{dp'}{d\alpha} + p' \frac{d\pi_1^*}{d\alpha} = \frac{d}{d\alpha} (p\pi_1') = p \frac{d\pi_1'}{d\alpha} . \tag{4}$$

Using this, (3) reduces to

$$\begin{aligned}
 \frac{\partial}{\partial \alpha} E(M_R) &= fp \frac{d\pi_1'}{d\alpha} \sum_{n=1}^g \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) b(n-1, \pi_1^*, g-1) \\
 &+ fp' \frac{d\pi_1^*}{d\alpha} \sum_{n=1}^g \pi_2 \left(\sqrt{\frac{nk+1}{2}} \frac{\Delta}{\sigma}, \beta \right) \left[(n-1) - (g-1) \pi_1^* \right] \\
 &\quad \times b(n-1, \pi_1^*, g-1) / (1 - \pi_1^*) . \tag{5}
 \end{aligned}$$

The first term in (5) is obviously positive for a non-zero β and with the same restriction on β , the second term is positive since it is the covariance between two non-decreasing functions of n , namely,

$$(n-1) \text{ and } \pi_2 \left(\sqrt{\frac{nk+1}{2}}, \frac{\Delta}{\sigma}, \beta \right)$$

multiplied by a positive term $f_p' \frac{d\pi_1^*}{d\alpha} / 1-\pi_1^*$. Hence,

$$\frac{\partial}{\partial \alpha} E(M_R) > 0.$$

Therefore, it is clear that $E(M_R)$ increases as α increases for a given $\beta \neq 0$.

Reference

- [1] Watson, G. S., "A Study of the Group Screening Method,"
 Technometrics, Vol. 3, No. 3, August 1961., pp. 371-388.