

ON THE MOMENTS OF SOME ONE-SIDED STOPPING RULES

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1. Introduction. The moments of stopping rules (or stopping times) have been discussed in [1,3,4], and the following results have been proved. Let x_n be independent random variables with $Ex_1 = 0$, $Ex_n^2 = 1$, and $S_n = x_1 + \dots + x_n$. For $c > 0$ and $m = 1, 2, \dots$, define t_m to be the first $n \geq m$ such that $|S_n| > c n^{1/2}$. If $c \geq 1$, then $Et_1 = \infty$. If $P[|x_n| \leq K] = 1$ for some $K < \infty$ and $n = 1, 2, \dots$, then $Et_m < \infty$ for every m if $c < 1$, $Et_m^2 < \infty$ for every m if $c < 3 - \sqrt{6}$, and $Et_m^2 = \infty$ for some large m if $c > 3 - \sqrt{6}$.

In this note, we are interested in the following one-sided stopping rules, instead of the above stated two-sided stopping rules. For $c > 0$ and $1 > p \geq 0$, define

$$s = \text{first } n \geq 1 \text{ such that } S_n \geq c n^p .$$

One of the results states that, if x_n are independent, $Ex_n = \mu > 0$, and $Ex_n^2 - \mu^2 = \sigma^2 < \infty$, then $Es^2 < \infty$ and

$$(1) \quad \lim_{c \rightarrow \infty} \mu^2 Es^2 / (c^2 Es^{2p}) = \lim_{c \rightarrow \infty} \mu Es^2 / (c Es^{1+p}) = 1 .$$

When $p = 0$, $Es^2 < \infty$ implies that $P[S_1 < c, \dots, S_n < c] = P[s > n] = o(n^{-2})$ as $n \rightarrow \infty$, which completes a result of Morimura [8]. Also (1) extends the elementary renewal theorem from first moments to second moments and generalizes some results due to Chow and Robbins [2] and Hatori [6].

2. The first moment.

Let $(\Omega, \mathfrak{F}, P)$ be a probability space and x_n be a sequence of integrable random variables. Let \mathfrak{F}_n be the Borel field generated by x_1, \dots, x_n and $\mathfrak{F}_0 = \{\emptyset, \Omega\}$. Put $S_n = x_1 + \dots + x_n$, $S_0 = 0$, $m_n = E(x_n | \mathfrak{F}_{n-1})$ and $T_n = \sum_{j=1}^n m_j$. Assume that for some constant $\infty > \mu > 0$ and for some null set N ,

$$(2) \quad \lim_{n \rightarrow \infty} T_n/n = \mu, \text{ uniformly on } \Omega - N.$$

For $c > 0$ and $1 > p \geq 0$, define

$$s = \text{first } n \geq 1 \text{ such that } S_n \geq c n^p.$$

Theorem 1. (i) If for some $0 < \delta < \mu/3$, $P[x_n \leq m_n + n\delta] = 1$ for all large n , then $Es < \infty$. (ii) If $E[(x_n - m_n)^+ | \mathfrak{F}_{n-1}] \leq K n^{-\alpha}$ for some $\alpha > 1$, then $Es < \infty$ and

$$(3) \quad \lim_{c \rightarrow \infty} \mu Es / (c Es^p) = 1 = \lim_{c \rightarrow \infty} ES_s / (c Es^p).$$

Proof. (1) Set $t = \min(s, k)$ for $k = 1, 2, \dots, T$.
Then by the Wald identity for martingales [see 5, p. 302; or 3],

$$E T_t = E S_t = E(S_{t-1} + x_t) = \leq c E t^p + E(m_t + \delta t).$$

Let $0 < \epsilon < \delta$. As $k \rightarrow \infty$, by (2)

$$E T_t \geq (\mu - \epsilon) E t + o(1), \quad E m_t = o(1) + o(E t).$$

Hence

$$(\mu - \epsilon) E t \leq c E t^p + \delta E t + o(1) + o(E t),$$

$$\int_{[s \leq k]} s dP + k P[s > k] = E t = o(1),$$

as $k \rightarrow \infty$. Therefore $P[s < \infty] = 1$ and $E s < \infty$.

(ii) For any $0 < \delta < \mu/3$, define $x'_n = \min(x_n, m_n + n\delta)$,
 $m'_n = E(x'_n | \mathcal{F}_{n-1})$, and $T'_n = m'_1 + \dots + m'_n$. Let $I(A)$ be the
indicator function of the set A . Then

$$\begin{aligned} 0 &\leq m_n - m'_n = E((x_n - m_n - n\delta) I[x_n > m_n + n\delta] | \mathcal{F}_{n-1}) \\ (4) \quad &\leq E((x_n - m_n) I[x_n > m_n + n\delta] | \mathcal{F}_{n-1}) \\ &\leq E^{1/\alpha}([(x_n - m_n)^+]^\alpha | \mathcal{F}_{n-1}) P^{1/\alpha'}(x_n - m_n > n\delta | \mathcal{F}_{n-1}) \quad (\alpha + \alpha' = \alpha\alpha') \\ &\leq K(n\delta)^{-\alpha/\alpha'} \end{aligned}$$

Therefore $\lim_{n \rightarrow \infty} T'_n/n = \mu$ uniformly on Ω -N. Define

$$t = \text{first } n \geq 1 \text{ such that } x_1^i + \dots + x_n^i \geq c n^p .$$

Then $s \leq t$. By (1), $Et < \infty$. Therefore $Es < \infty$ and it follows by the Wald identity again [5, p. 302; or 3] that

$$(5) \quad E(c s^p + x_s) \geq ES_s = ET_s \geq c Es^p .$$

Let $Z_n = \sum_1^n [(x_j - m_j)^+]^\alpha$. Then by Lemma 6 of [3],

$$(6) \quad E^\alpha(x_s - m_s)^+ \leq EZ_s = E \sum_1^s E([(x_j - m_j)^+]^\alpha | \mathcal{F}_{j-1}) \leq K Es .$$

Since (2) implies that as $c \rightarrow \infty$, $Em_s = O(1) + o(Es)$ and $ET_s = O(1) + (\mu + O(1)) Es$, we have

$$(7) \quad Ex_s = O(E^{1/\alpha} s) + o(Es) + O(1)$$

from (6), and

$$\lim_{c \rightarrow \infty} \mu Es / (c Es^p) = \lim_{c \rightarrow \infty} ET_s / (c Es^p) = \lim_{c \rightarrow \infty} ES_s / (c Es^p) = 1$$

from (5) and (7), since $\lim_{c \rightarrow \infty} Es = \infty$. The proof is completed.

When $p = 0$, part (ii) of Theorem 1 reduces to an elementary renewal theorem, which was proved in [2], in a slightly restricted form by requiring that $m_n = E(x_n)$ for each n .

3. The second moment.

Assume that $EX_n^2 < \infty$ for each n , let $V_n = \sum_1^n E((x_j - m_j)^2 | \mathcal{F}_{j-1})$ for $n = 1, 2, \dots$, and define s as before. For a random variable y , put $\|y\| = (Ey^2)^{1/2}$.

Theorem 2. If (2) holds and $E((x_n - m_n)^2 | \mathcal{F}_{n-1}) \leq K < \infty$, then $Es^2 < \infty$, $ES_s^2 < \infty$, and as $c \rightarrow \infty$

$$(8) \quad ES_s^2 + ET_s^2 = EV_s + 2ES_s T_s,$$

$$(9) \quad \lim ES_s^2 / ET_s^2 = 1,$$

$$(10) \quad \lim \mu^2 Es^2 / (c^2 Es^{2p}) = 1,$$

$$(11) \quad \lim ES_s^2 / (c^2 Es^{2p}) = 1,$$

$$(12) \quad \lim \mu Es^2 / (c Es^{1+p}) = 1.$$

Proof. (1) First, assume that for some $0 < \delta < \mu/8$ and $0 < M < \infty$, $P[x_n \leq m_n + n\delta + M] = 1$. Set $t = \min(t, k)$ for $k = 1, 2, \dots$. Then by Theorem 1 and Lemma 6 of [3],

$$E(S_t - T_t)^2 = EV_t \leq K Et.$$

Hence by Schwarz inequality

$$(13) \quad ES_t^2 + ET_t^2 \leq K Et + 2\|T_t\| \cdot \|S_t\|.$$

Assume, on the contrary, that $Es^2 = \infty$. Then $\lim_{k \rightarrow \infty} Et^2 = \infty$ and (2) implies that

$$Em_t^2 = O(1) + o(Et^2) = o(Et^2),$$

as $k \rightarrow \infty$. Hence

$$(14) \quad ||S_t|| \leq ||ct^p + m_t + \delta t + M|| \leq c||t^p|| + \delta||t|| + o(||t||) \\ = (\delta + o(1))||t||;$$

and from (2),

$$(15) \quad ET_t^2 = O(1) + (\mu^2 + o(1))Et^2 = (\mu^2 + o(1))Et^2.$$

By (13), (14) and (15), we have

$$1 + ES_t^2/ET_t^2 \leq O(||t||^{-1}) + 2||S_t||/||T_t|| \\ \leq O(||t||^{-1}) + (2\delta + o(1))/\mu = 2\delta/\mu + o(1).$$

Since $\delta < \mu/8$, we have a contradiction when k is large.

Therefore $Es^2 < \infty$. From (14), (13) and Fatou's lemma

$$ES_s^2 < \infty \text{ and } ET_s^2 < \infty.$$

(ii) For the general case, let $x_n' = \min(x_n, m_n + n\delta + M)$ for an arbitrary constant $\infty > M > 0$ and $0 < \delta < \mu/8$. Define m_n' , T_n' and t as in the proof of part (i) of Theorem 1. Then by (4) (for $\alpha=2$), $0 \leq m_n - m_n' \leq K(n\delta)^{-1}$. Hence

$$\lim T'_n/n = \mu \quad \text{uniformly on } \Omega-N .$$

It is not too difficult to see that

$$E((x_n - m_n)^2 | \mathfrak{F}_{n-1}) - E((x'_n - m'_n)^2 | \mathfrak{F}_{n-1}) \geq E([(x_n - m_n - n\delta - M)^+]^2 | \mathfrak{F}_{n-1}) - E^2((x_n - m_n - n\delta - M)^+ | \mathfrak{F}_{n-1}) \geq 0.$$

Therefore $E((x'_n - m'_n)^2 | \mathfrak{F}_{n-1}) \leq K$. Since $t \geq s$ and from part (1) $Et^2 < \infty$, we have that $Es^2 < \infty$. By Theorem 1 and Lemma 6 of [3] again,

$$(16) \quad E(S_s - T_s)^2 = EV_s \leq K Es .$$

For $\epsilon > 0$, (2) implies that there exists a constant $\infty > L > 0$ such that

$$ET_s^2 \leq L + (\mu^2 + \epsilon) Es^2 .$$

Hence $ET_s^2 < \infty$ and from (16), $ES_s^2 < \infty$. Thus (8) follows.

Now by (16),

$$|ES_s^2 - ET_s^2| \leq E|S_s^2 - T_s^2| \leq ||S_s - T_s|| \cdot ||S_s + T_s|| \leq (K Es)^{1/2} ||S_s + T_s|| .$$

Since $ES_s^2 \geq c^2 Es^{2p}$, from (3)

$$|1 - ET_s^2/ES_s^2| \leq (K Es/ES_s^2)^{1/2} (1 + ||T_s||/||S_s||) = o(1) + o(||T_s||/||S_s||)$$

as $c \rightarrow \infty$. Hence (9) follows.

Since (2) implies that $ET_s^2 = O(1) + (\mu^2 + o(1)) Es^2$ as $c \rightarrow \infty$, from (9)

$$(17) \quad \lim_{c \rightarrow \infty} \mu^2 Es^2 / ET_s^2 = 1 = \lim_{c \rightarrow \infty} \mu^2 Es^2 / ES_s^2 .$$

Let $Z_n = \sum_1^n (x_j - m_j)^2$. Applying Lemma 6 of [3], we have

$$E(x_s - m_s)^2 \leq EZ_s = E \sum_1^s E((x_j - m_j)^2 | \mathfrak{F}_{j-1}) \leq KEs .$$

From (2), $Em_s^2 = O(1) + o(Es^2) = o(Es^2)$ as $c \rightarrow \infty$. Hence

$$(18) \quad Ex_s^2 = E(x_s - m_s + m_s)^2 = o(Es^2), \quad ||x_s|| = o(||s||) .$$

Now from (18), as $c \rightarrow \infty$

$$(19) \quad c||s^D|| \leq ||S_s|| \leq ||cs^D + x_s|| \leq c||s^D|| + ||x_s|| = c||s^D|| + o(||s||) .$$

Therefore (10) follows from (17) and (19), and (11) follows from (17) and (10).

Now $ET_s S_s = O(ES_s) + (\mu + o(1)) Es S_s$ as $c \rightarrow \infty$. By the definition of s and (18), as $c \rightarrow \infty$

$$(20) \quad c E s^{1+p} \leq E s S_s \leq c E s^{1+p} + E s x_s \leq c E s^{1+p} + ||s|| \cdot ||x_s|| \\ \leq c E s^{1+p} + o(E s^2) .$$

Since $E V_s \leq K E s$, from (8), (9), (10), and (11), $\lim E S_s T_s / (\mu^2 E s^2) = 1$.

Hence

$$\lim E s S_s / (\mu E s^2) = 1$$

and then (20) implies (12).

4. Corollaries and comments.

In this section we assume that x_n is a sequence of random variables and $p=0$. Define $S_n, m_n, T_n, \mathfrak{F}_n$ and s as in Section 2.

Corollary 1. If (2) holds and if

$$(21) \quad E((x_n - m_n)^2 | \mathfrak{F}_{n-1}) \leq K < \infty ,$$

then $E s^2 < \infty$ and

$$(22) \quad \lim_{c \rightarrow \infty} E s^c / c^\alpha = \mu^{-\alpha} \quad \text{for } 0 \leq \alpha \leq 2 .$$

Proof. Since (21) implies $E(x_n - m_n)^2 \leq K$, from (2) and (21) it follows [7] that $\lim S_n/n = \mu$ a.e. Hence

$$1 \leq \liminf_{c \rightarrow \infty} S_s/c \leq \limsup \mu s/c = \limsup \mu(s-1)/c \\ = \limsup S_{s-1}/c \leq 1 .$$

Therefore $\lim s/c = \mu^{-1}$ a.e. Theorem 2 implies that $E(s/c)^2 \leq M < \infty$ for all $c > 0$. Hence [see 5, p. 629] for every $0 \leq \alpha < 2$, $(s/c)^\alpha$ is uniformly integrable and

$$(23) \quad \lim_{c \rightarrow \infty} E|\mu^{-1} - s/c|^\alpha = 0, \quad \lim E s^\alpha / c^\alpha = \mu^{-\alpha} .$$

Thus (22) follows from (23) and (10).

Corollary 2. Let x_n be a sequence of independent, identically distributed random variables such that $E x_1 > 0$ and $E(x_1 - E x_1)^2 > \infty$. Then for every $c > 0$, as $n \rightarrow \infty$,

$$(24) \quad P[S_1 < c, \dots, S_n < c] = o(n^{-2}) .$$

Proof. Since $[s > n] = [S_1 < c, \dots, S_n < c]$, $E s^2 < \infty$ implies (24) and thus Corollary 2 follows from Corollary 1.

(22) has been proved by Hatori [6] for every $\alpha > 0$, by requiring, in addition to the assumptions of Corollary 1, that x_n be independent, $P[x_n \geq 0] = 1$ and $m_n \geq L > 0$ for each n .

Under the conditions of Corollary 2, Morimura [8] proves that $P[S_1 < c, \dots, S_n < c] = O(n^{-\delta})$ for $0 \leq \delta < (1 + \sqrt{5})/2$ and that there exists an example such that for some $D > 0$ and for each $\epsilon > 0$, $P[S_1 < c, \dots, S_n < c] \geq D n^{-2-\epsilon}$ when n is large enough. Thus (24) is the best possible. Clearly, Corollary 2 completes Morimura's work.

The counter example in [8] satisfies the condition $Es^{2+\epsilon} = \infty$ for every $\epsilon > 0$, since $P[s > n] \neq o(n^{-2-\epsilon})$. Therefore (22) can not be extended to the cases where $\alpha > 2$, without some conditions as $P[x_n \geq 0] = 1$ imposed in [6].

References

- [1] Blackwell, D. and Freedman, D. (1964). A remark on the coin tossing game. Ann. Math. Statist. 35, pp. 1345-1347.
- [2] Chow, Y.S. and Robbins, H. (1963). A renewal theorem for random variables which are dependent or non-identically distributed. Ann. Math. Statist. 34, pp. 390-395.
- [3] Chow, Y.S., Robbins, H. and Teicher, H. (1965). Moments of randomly stopped sums. Ann. Math. Statist. 36
- [4] Chow, Y.S. and Teicher, H. (1965). On the second moments of some two-sided stopping rules. (unpublished)
- [5] Doob, J.L. (1953). Stochastic processes. Wiley, New York.
- [6] Hatori, H. (1959). Some theorems in an extended renewal theory, I. Kodai Math. Sem. Rep. 11, pp. 139-146.
- [7] Loève, M. (1951). On almost sure convergence. Proc. Second Berkeley Symp. Math. Statist. Prob. pp. 279-303.
- [8] Morimura, H. (1961). A note on sums of independent random variables. Kodai Math. Sem. Rep. 13, pp. 255-260.