

Exact and Asymptotic Distributions of Some
Statistics in Multivariate Analysis

by

B. N. Nagarsenker

Purdue University

Department of Statistics
Division of Mathematical Science
Mimeograph Series #300

August 1972

$$(2.5) \quad C(p, n, \lambda) = \Gamma_p(n/2) [\Gamma_p(n_1/2) \Gamma_p(n_2/2)]^{-1} |\lambda|^{-n_1/2}.$$

Now using the density (2.3) we get

$$(2.6) \quad E(Y^h) = C(p, n, \lambda) \pi^{p(p-1)/4} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} C_{\kappa}(M)}{k!} \cdot \frac{\prod_{i=1}^p \Gamma[n_1/2 + ah + k_i - (i-1)/2] \prod_{i=1}^p \Gamma[n_2/2 + bh - (i-1)/2]}{\prod_{i=1}^p \Gamma[n/2 + (a+b)h + k_i - (i-1)/2]}.$$

Making use of the inverse Mellin transform, we have the density of Y as

$$(2.7) \quad f(Y) = C(p, n, \lambda) \pi^{p(p-1)/4} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} C_{\kappa}(M)}{k!} Y^{-1} \cdot \frac{1}{2\pi i} \int_C Y^{-h} \frac{\prod_{i=1}^p \Gamma[n_1/2 + ah + k_i - (i-1)/2] \prod_{i=1}^p \Gamma[n_2/2 + bh - (i-1)/2]}{\prod_{i=1}^p \Gamma[n/2 + (a+b)h + k_i - (i-1)/2]} dh.$$

Noting that the integral on the R.H.S. of (2.7) is in the form of the H-function, the non-central density of Y for test (1) can be put in a single general form for different sets of values of a and b as follows:

$$(2.8) \quad f(Y) = C(p, n, \lambda) \alpha \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} \delta C_{\kappa}(M)}{k!} Y^{-1} H_{t,u}^{r,s}(Y | \begin{matrix} (a_i, \alpha_i)_{i=1, \dots, t} \\ (b_i, \beta_i)_{i=1, \dots, u} \end{matrix})$$

where $C(p, n, \lambda)$ is as in (2.5) and the constants are given in the following table.

$$(2.14) \quad f(Y) = C_2(p, n, q, \lambda^2) \alpha \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} (n/2)_{\kappa} \delta C_{\kappa}(\lambda^2)}{(q/2)_{\kappa} k!}$$

$$\cdot H_{t,u}^{r,s}(Y | \begin{matrix} (a_i, \alpha_i)_{i=1, \dots, t} \\ (b_i, \beta_i)_{i=1, \dots, u} \end{matrix}),$$

where the constants α , δ , r , s , t , u , (a_i, α_i) and (b_i, β_i) are as in Table 1, in which n_1 is to be replaced by q throughout.

3. SPECIAL CASES

(i) Wilks' Λ criterion. Taking $a = 0$ and $b = 1$ in (2.8) and using the relation between the H-function and the G-function, we find that the non-central density of Wilks' $\Lambda = \prod_{i=1}^p (1 - \theta_i)$ is as obtained by Pillai, Al-Ani, Jouris in the three cases [25]

(ii) Wilks-Lawley U-criterion. If $a = 1$ and $b = 0$ in (2.8), we obtain the non-central density of Wilks-Lawley U-statistic,

$U = \prod_{i=1}^p \theta_i$ for test (1), in the form

$$(3.1) \quad f(u) = C(p, n, \lambda^2) \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} C_{\kappa}(M)}{k!} \Gamma_p(n_2/2)$$

$$\cdot Y^{-1} H_{p,p}^{p,0}(Y | \begin{matrix} (a_i, \alpha_i)_{i=1, \dots, p} \\ (b_i, \beta_i)_{i=1, \dots, p} \end{matrix}),$$

where $C(p, n, \lambda^2)$ is as in (2.4), $(a_i, \alpha_i) = (n/2 + k_i - (i-1)/2, 1)$ and $(b_i, \beta_i) = (n_1/2 + k_i - (i-1)/2, 1)$ $i=1, \dots, p$. (3.1) can also be expressed in terms of the G-function. The density of U for the Manova and Canonical correlation cases can be written down using the substitution (2.10) and (2.13) respectively.

(iii) Taking $a = n_1/2$ and $b = n_2/2$ in (2.8) we obtain the non-central density of the modified likelihood ratio criterion for testing $\Sigma_1 = \Sigma_2$ i.e. of the statistic

$$\lambda = \prod_{i=1}^p \theta_i^{n_1/2} (1-\theta_i)^{n_2/2} = |\Sigma_1|^{n_1/2} |\Sigma|^{-n/2} |\Sigma_2|^{n_2/2} \text{ where } \Sigma = \Sigma_1 + \Sigma_2,$$

in the form

$$f(\lambda) = C(p, n, \lambda) \pi^{p(p-1)/4} \sum_{k=0}^{\infty} \frac{(n/2)_{\kappa} C_{\kappa}^{(M)}}{k!} \lambda^{-1} H_{p, 2p}^{2p, 0}(\lambda | \begin{matrix} (a_i, \alpha_i)_{i=1, \dots, p} \\ (b_i, \beta_i)_{i=1, \dots, p} \end{matrix}),$$

where $(a_i, \alpha_i) = (n/2 + k_i - (i-1)/2, n/2)$ and $(b_i, \beta_i) =$

$\{(n_1/2 + k_i - (i-1)/2, n_1/2), (n_2/2 - (i-1)/2, n_2/2) \mid i = 1, \dots, p\}$.

The densities in the other two cases can be written down using (2.10) and (2.13).

(iv) Taking $a = 1$ and $b = -1$ in (2.8) we obtain the noncentral density of the statistic $W = \prod_{i=1}^p \theta_i (1-\theta_i)^{-1} = |\Sigma_1 \Sigma_2^{-1}|$ for test (1)

in the form

$$(3.2) \quad f(Y) = C(p, n, \lambda) \pi^{p(p-1)/2} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} C_{\kappa}^{(M)}}{k! \Gamma_p(n/2, \kappa)} Y^{-1}$$

$$H_{p, p}^{p, p}(Y | \begin{matrix} (a_i, 1) \quad i = 1, \dots, p \\ (b_i, 1) \quad i = 1, \dots, p \end{matrix})$$

where $a_i = 1 - n_2/2 + (i-1)/2$ and $b_i = n_1/2 + k_i - (i-1)/2$. The density in (3.2) can be easily written down in terms of the G-function. The non-central densities of W for the Manova and Canonical correlation cases can be written down using (2.10) and (2.13).

4. NON-CENTRAL DISTRIBUTION OF Y IN THE COMPLEX CASE

The non-central density of Y in the complex case can be obtained in a similar manner and is noted below. (a') The general form of the density of Y for test (1) can be written down from (2.8) by making the following substitutions.

$$(4.1) \quad (\pi^d, n_1/2, n_2/2, n/2, (i-1)/2, \Gamma_p(\cdot), \Gamma_p(\cdot, \kappa), C_\kappa(\cdot), (\cdot)_\kappa) \\ \rightarrow (\pi^{2d}, n_1, n_2, n, (i-1), \tilde{\Gamma}_p(\cdot), \tilde{\Gamma}_p(\cdot, \kappa), \tilde{C}_\kappa(\cdot), [\cdot]_\kappa)$$

where $\tilde{\Gamma}_p(\cdot)$, $\tilde{\Gamma}_p(\cdot, \kappa)$, $\tilde{C}_\kappa(\cdot)$ and $[\cdot]_\kappa$ are as defined in James [12].

(b') For the Manova case the general form of the density of Y is obtained from (2.11) by making the substitutions as in (4.1).

(c') In the case of Canonical correlation also, the general form of the density of Y can be written down from (2.14) using (4.1).

5. ASYMPTOTIC EXPANSION OF THE DISTRIBUTION OF Y, $a=n_1/2$ AND $b=n_2/2$

First we give some preliminaries.

(a) Preliminaries. For this case, putting $a = n_1/2$ and $b = n_2/2$ in (2.6) we have,

$$(5.1) \quad E(Y^h) = \frac{\Gamma_p(n/2)}{\Gamma_p(n_1/2)\Gamma_p(n_2/2)} |\Lambda|^{-n_1/2} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_\kappa \Gamma_p\left\{\frac{n_1(1+h)}{2}, \kappa\right\}}{k! \Gamma_p\left\{\frac{n(1+h)}{2}, \kappa\right\}} \\ \cdot \Gamma_p\{n_2(1+h)/2\} C_\kappa(M).$$

This can be easily written in the form

$$(5.2) \quad E(Y^h) = \left\{ \Gamma_p(n/2) \Gamma_p[n_1(1+h)/2] \Gamma_p[n_2(1+h)/2] / \Gamma_p[n(1+h)/2] \right. \\ \left. \Gamma_p(n_1/2) \Gamma_p(n_2/2) \right\} |\Lambda|^{-n_1/2} \cdot {}_2F_1(n/2, n_1(1+h)/2; n(1+h)/2, M).$$

We shall assume that

$$(5.3) \quad n_i = \tau_i n, \quad (i = 1, 2) \text{ where } \tau_1 + \tau_2 = 1.$$

The asymptotic expansion of the distribution of Y will be derived in terms of n increasing and also in terms of $m = \rho n$ increasing where $0 < \rho < 1$ and is defined later, with τ_1 and τ_2 fixed. (See Anderson [1], p. 254). The h^{th} moment of

$$(5.4) \quad W = [n^{(1/2)pn} / n_1^{(1/2)pn_1} n_2^{(1/2)pn_2}] \cdot Y$$

is given by

$$(5.5) \quad E(W^h) = n^{(1/2)pnh} n_1^{-(1/2)pn_1h} n_2^{-(1/2)pn_2h} \left\{ \Gamma_p(n/2) \Gamma_p[n_1(1+h)/2] / \right. \\ \left. \Gamma_p[n(1+h)/2] \Gamma_p(n_1/2) \right\} \left\{ \Gamma_p[n_2(1+h)/2] / \Gamma_p(n_2/2) \right\} \\ |\Lambda|^{-n_1/2} \cdot {}_2F_1(n/2, n_1(1+h)/2; n(1+h)/2, M).$$

We shall obtain the asymptotic expansion for (i) $-2 \log W$ in terms of n increasing and assuming M to be of the form $M = (2/n) P$ where P is a fixed matrix and (ii) $-2\rho \log W$ in terms of $m = \rho n$ increasing instead of n and assuming $M = (2/m) P$ where P is a fixed matrix and the correction factor ρ is given by (see Anderson [1], p. 255)

$$(5.6) \quad m = pn = n - 2\alpha \text{ where } \alpha = (\tau_1^{-1} + \tau_2^{-1} - 1)(2p^2 + 3p - 1)/12(p+1).$$

We will need the following lemmas proved in [28].

Lemma 5.1. Let $C_\kappa(Z)$ be a zonal polynomial corresponding to the partition $\kappa = \{k_1, k_2, \dots, k_p\}$ with $k_1 + k_2 + \dots + k_p = k$ and $k_1 \geq k_2 \geq k_3 \dots \geq k_p \geq 0$. Putting

$$(5.7) \quad a_1(\kappa) = \sum_{i=1}^p k_i(k_i - 1), a_2(\kappa) = \sum_{i=1}^p k_i(4k_i^2 - 6ik_i + 3i^2)$$

Then the following equalities hold:

$$(5.8) \quad \sum_{k=0}^{\infty} \sum_{\kappa} x^k C_{\kappa}(Z) a_1(\kappa)/k! = (x^2 \operatorname{tr} Z^2) e^{\operatorname{tr}(xZ)},$$

$$(5.9) \quad \sum_{k=1}^{\infty} \sum_{\kappa} x^k C_{\kappa}(Z) a_1(\kappa)/(k-1)! = (2x^2 \operatorname{tr} Z^2 + x^3 \operatorname{tr} Z^2 \operatorname{tr} Z) e^{\operatorname{tr}(xZ)},$$

$$(5.10) \quad \sum_{k=0}^{\infty} \sum_{\kappa} x^k (a_1(\kappa))^2 C_{\kappa}(Z)/k! = \{x^4 (\operatorname{tr} Z^2)^2 + 4x^3 \operatorname{tr} Z^3 + x^2 \operatorname{tr} Z^2 + x^2 (\operatorname{tr} Z)^2\} e^{\operatorname{tr}(xZ)},$$

$$(5.11) \quad \sum_{k=0}^{\infty} \sum_{\kappa} x^k C_{\kappa}(Z) a_2(\kappa)/k! = \{4x^3 \operatorname{tr} Z^3 + 3x^2 \operatorname{tr} Z^2 + 3x^2 (\operatorname{tr} Z)^2 + x \operatorname{tr} Z\} e^{\operatorname{tr}(xZ)},$$

$$(5.12) \quad \sum_{k=2}^{\infty} \sum_{\kappa} C_{\kappa}(Z)/(k-1)! = (\operatorname{tr} Z) e^{\operatorname{tr} Z},$$

and

$$(5.13) \quad \sum_{k=2}^{\infty} \sum_{\kappa} C_{\kappa}(Z)/(k-2)! = (\operatorname{tr} Z)^2 e^{\operatorname{tr} Z}.$$

Lemma 5.2. With the notations of the lemma 5.1, for large n

$$(5.14) \quad (n/2)_{\kappa} = (n/2)^k [1 + n^{-1} a_1(\kappa) + (1/6n^2) \{k - a_2(\kappa) + 3(a_1(\kappa))^2\} + O(n^{-3})],$$

and

$$(5.15) \quad (n/2+a)_{\kappa} = (n/2)^k [1 + (1/2n) \{4ak + 2a_1(\kappa)\} + (1/24n^2) \{4k + 48a^2 k(k-1) + 48a(k-1)a_1(\kappa) - 4a_2(\kappa) + 12(a_1(\kappa))^2\} + O(n^{-3})].$$

(b) Derivation of Asymptotic Expansions. We consider below asymptotic expansions of the distributions of (i) and (ii) above.

(i) Asymptotic expansion of the distribution of $-2 \log W$. Let $\phi(t)$ be the characteristic function of $-2 \log W$. Then from (5.2) we have

$$(5.16) \quad \phi(t) = E(e^{-2it \log W}) = E(W^{-2it}) = C_1(t)C_2(t)C_3(t) |A|^{-n_1/2}$$

where

$$(5.17) \quad C_1(t) = n^{-itpn} n_1^{itpn_1} n_2^{itpn_2},$$

$$(5.18) \quad C_2(t) = \Gamma_p(n/2) \Gamma_p(n_1 g/2) \Gamma_p(n_2 g/2) [\Gamma_p(n g/2) \Gamma_p(n_1/2) \Gamma_p(n_2/2)]^{-1}$$

and

$$(5.19) \quad g = (1-2it), \quad C_3(t) = {}_2F_1(n/2, n_1 g/2; n g/2, M).$$

We shall use the following asymptotic formula for the gamma function as in Anderson ([1], p. 204)

$$(5.20) \quad \log \Gamma(x+h) = \log \sqrt{2\pi} + (x+h - \frac{1}{2}) \log x - x - \sum_{r=1}^m \frac{(-1)^r B_{r+1}(h)}{r(r+1)x^r} + O(|x|^{-m-1}),$$

which holds for large $|x|$ and fixed h . The Bernoulli polynomial $B_r(h)$ of degree r is given by $(t e^{ht})/(e^t-1) = \sum_{r=0}^{\infty} \frac{t^r}{r!} B_r(h)$. Some of these which we shall need in the sequel, are listed below.

$$B_1(h) = h-1/2, \quad B_2(h) = h^2-h+1/6, \quad (5.21)$$

$$B_3(h) = h^3-3h^2/2+h/2 \quad \text{and} \quad B_4(h) = h^4-2h^3+h^2-1/30.$$

Applying the formula (5.20) to each gamma function in $C_2(t)$, we have

$$(5.22) \quad \log C_2(t) = it \, p n \log(n/2) - it \, p n_2 \log(n_2/2) - it \, p n_1 \log(n_1/2) \\ - f \log(g)/2 + (r/n)(g^{-1}-1) + (s/n^2)(1-g^{-2}) + o(n^{-3}),$$

where

$$(5.23) \quad f = p(p+1)/2, \quad r = p(2p^2+3p-1)(\tau_1^{-1} + \tau_2^{-2} - 1)/24$$

and

$$s = p(p+1)(2-p^2-p)(\tau_1^{-2} + \tau_2^{-2} - 1)/48.$$

It therefore follows that

$$(5.24) \quad C_1(t)C_2(t) = g^{-f/2} \exp[(r/n)(g^{-1}-1) + (s/n^2)(1-g^{-2}) + o(n^{-3})] \\ = g^{-f/2} [1 + (r/n)(g^{-1}-1) + n^{-2}\{s(1-g^{-2}) + (r^2/2)(g^{-1}-1)^2\} \\ + o(n^{-3})].$$

Let $M = [I - \Lambda^{-1}] = (2/n) P$ where P is a fixed matrix. Then

$$(5.25) \quad |\Lambda|^{-n_1/2} = |\Lambda - \frac{2}{n} P| \tau_1^{n/2}.$$

Now using the expansion

$$(5.26) \quad \log |\Lambda - \frac{2}{n} P| = -(2/n) \text{tr } P - (2/n^2) \text{tr}(P^2) - (8/3n^3) \text{tr}(P^3) + O(n^{-4})$$

we obtain

$$(5.27) \quad \begin{aligned} |\Lambda - \frac{2}{n} P| \tau_1^{n/2} &= \exp[(\tau_1 n/2) \log |\Lambda - \frac{2}{n} P|] \\ &= e^{(\tau_1 n/2) [-(2/n) \text{tr } P - (2/n^2) \text{tr}(P^2) - (8/3n^3) \text{tr}(P^3) + O(n^{-4})]} \\ &= e^{-\tau_1 \text{tr } P} [1 - n^{-1} A_1 - n^{-2} A_2 + O(n^{-3})] \end{aligned}$$

where

$$(5.28) \quad A_1 = \tau_1 \text{tr}(P^2) \text{ and } A_2 = (4/3) \tau_1 \text{tr}(P^3) - \tau_1^2 (\text{tr } P^2)^2 / 2.$$

Applying asymptotic formula (5.14) to $(n/2)_\kappa$, $(n_1 g/2)_\kappa$ and $(ng/2)_\kappa$ we have after some algebraic simplification,

$$(5.29) \quad \begin{aligned} &(n/2)_\kappa (n_1 g/2)_\kappa / (ng/2)_\kappa \\ &= (n \tau_1 / 2)^k [1 + n^{-1} a_1(\kappa) B(t) + (1/6n^2) \{((k - a_2(\kappa) + 3(a_1(\kappa)))^2) \\ &\quad A(t) - g^{-2} (k - a_2(\kappa) - 3(a_1(\kappa)))^2 - D(t) (a_1(\kappa))^2\} + O(n^{-3})] \end{aligned}$$

where

$$(5.30) \quad A(t) = 1 + (\tau_1 g)^{-2}, \quad B(t) = 1 + (\tau_1^{-1} - 1)/g \text{ and}$$

$$D(t) = 6[g^{-1} + (\tau_1 g^2)^{-1} - (\tau_1 g)^{-1}].$$

Using (5.29) and the lemma 5.1, we have on simplification,

$$(5.31) \quad C_3(t) = \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(n/2)_{\kappa} (n_1 g/2)_{\kappa} C_{\kappa}(\frac{2}{n} R)}{(ng/2)_{\kappa} k!}$$

$$= e^{\tau_1 \text{tr } R} [1 + (K/n)B(t) + (1/6n^2)\{LA(t) - Mg^{-2} - ND(t)\} + O(n^{-3})]$$

where

$$K = \tau_1^2 \text{tr } R^2, \quad L = 8\tau_1^3 \text{tr } R^3 + 3\tau_1^4 (\text{tr } R^2)^2,$$

$$(5.32) \quad M = -3\tau_1^4 (\text{tr } R^2)^2 - 16\tau_1^3 \text{tr } R^3 - 6\tau_1^2 \{\text{tr } R^2 + (\text{tr } R)^2\},$$

and

$$N = \tau_1^4 (\text{tr } R^2)^2 + 4\tau_1^3 \text{tr } R^3 + \tau_1^2 \{\text{tr } R^2 + (\text{tr } R)^2\}.$$

From (5.16), (5.24), (5.27) and (5.31), we have

$$(5.33) \quad \phi(t) = g^{-f/2} [1 + n^{-1} \{\alpha_0 + g^{-1} \alpha_1\} + n^{-2} \{\alpha_2 + g^{-1} \alpha_3 + g^{-2} \alpha_4\} + O(n^{-3})],$$

where the coefficients α_i 's are given by

$$\alpha_0 = K - A_1 - r, \quad \alpha_1 = K(\tau_1^{-1} - 1) + r,$$

$$\alpha_2 = L/6 - A_2 - KA_1 + s + r^2/2 - Kr + A_1 r,$$

$$(5.34) \quad \alpha_3 = r(K - A_1) - r^2 + (\tau_1^{-1} - 1)(N - rK - A_1 K),$$

and

$$\alpha_4 = (L \tau_1^{-2} - M)/6 - N \tau_1^{-1} - s + r^2/2 + rK(\tau_1^{-1} - 1).$$

By inverting the characteristic function in (5.33), using the fact that $(g)^{-f/2}$ is the characteristic function of χ_f^2 , a chi square variable with f degrees of freedom, we obtain the following asymptotic expansion for the distribution of $-2 \log W$.

$$\begin{aligned}
 (5.35) \quad P(-2 \log W \leq z) &= P(\chi_{\xi}^2 \leq z) + n^{-1} (\alpha_0 P(\chi_{\xi}^2 \leq z) \\
 &+ \alpha_1 P(\chi_{\xi+2}^2 \leq z)) + n^{-2} (\alpha_2 P(\chi_{\xi}^2 \leq z) + \alpha_3 P(\chi_{\xi+2}^2 \leq z)) \\
 &+ \alpha_4 P(\chi_{\xi+4}^2 \leq z) + O(n^{-3}),
 \end{aligned}$$

where α_1, α_3 are defined in (5.35).

(ii) Asymptotic expansion of the distribution of $-2\rho \log W$. Here we shall derive the asymptotic expansion for $-2\rho \log W$ where ρ is given by (5.6). Put $m = \rho n$ and let m tend to infinity instead of n . From (5.2), the characteristic function $f(t)$ of $-2\rho \log W$ can be written as

$$(5.36) \quad f(t) = E(e^{-2\rho i t \log W}) = C_4(t)C_5(t),$$

where $C_4(t)$ and $C_5(t)$ are given by

$$(5.37) \quad C_4(t) = \frac{n^{-pn_1 t \rho}}{n_1} \frac{n^{-pn_2 t \rho}}{n_2} \frac{\Gamma_p(n/2) \Gamma_p\left[\frac{n_1(1-2\rho i t)}{2}\right] \Gamma_p\left[\frac{n_2(1-2i t \rho)}{2}\right]}{\Gamma_p\left[\frac{n(1-2\rho i t)}{2}\right] \Gamma_p(n_1/2) \Gamma_p(n_2/2)}$$

and

$$(5.38) \quad C_5(t) = |\Lambda|^{-\frac{\tau_1}{2}(m+2\alpha)} {}_2F_1\left(\frac{m}{2} + \alpha, \frac{mg\tau_1}{2} + \alpha\tau_1; \frac{mg}{2} + \alpha, M\right),$$

g and α being as defined in (5.19) and (5.6) respectively. Now the first factor $C_4(t)$ in (5.36) can be expanded asymptotically (See Anderson [1], p. 255) as follows.

$$(5.39) \quad C_4(t) = g^{-f/2} [1 + (A/m^2) (g^{-2} - 1) + O(m^{-3})]$$

where

$$(5.40) \quad f = p(p+1)/2, \quad A = [p(p+1)/48] [(p-1)(p+2)(\tau_1^{-2} + \tau_2^{-2} - 1) - 6\gamma]$$

$$\text{and } \gamma = (\tau_1^{-1} + \tau_2^{-1} - 1)^2 (2p^2 + 3p - 1)^2 / 36(p+1)^2 = 4\alpha^2.$$

Now as stated before, let

$$\mathbb{I} - \mathbb{A}^{-1} = (2/m) \mathbb{P},$$

where \mathbb{P} is a fixed matrix. We then have

$$(5.41) \quad C_5(t) = \left| \mathbb{I} - \frac{2}{m} \mathbb{P} \right|^{\frac{\tau_1(m+2\alpha)}{2}} {}_2F_1\left(\frac{m}{2} + \alpha, \frac{m\tau_1 g}{2} + \alpha\tau_1; \frac{mg}{2} + \alpha, \frac{2}{m} \mathbb{P}\right).$$

Using the asymptotic expansion (5.15) to $\left(\frac{m}{2} + \alpha\right)_\kappa$, $(m\tau_1 g/2 + \alpha\tau_1)_\kappa$ and $(mg/2 + \alpha)_\kappa$, we have

$$(5.42) \quad \begin{aligned} & (m/2 + \alpha)_\kappa (m\tau_1 g/2 + \alpha\tau_1)_\kappa / (mg/2 + \alpha)_\kappa \\ &= (m\tau_1/2)^k [1 + m^{-1} \{2\alpha k + \delta_1 a_1(\kappa)\} + m^{-2} \{k\delta_2 + k^2\delta_3 + \delta_4 a_1(\kappa) \\ &+ \delta_5 k a_1(\kappa) + \delta_6 a_2(\kappa) + \delta_7 (a_1(\kappa))^2\} + 0(m^{-3})], \end{aligned}$$

where

$$\delta_1 = 1 + (\tau_1^{-1} - 1)g^{-1}, \quad \delta_2 = -2\alpha^2 + [1 + (\tau_1^{-2} - 1)g^{-2}]/6,$$

$$\delta_3 = 2\alpha^2,$$

$$(5.43) \quad \delta_4 = -2\alpha + 2\alpha g^{-2}(1 - \tau_1^{-1}), \quad \delta_5 = 2\alpha + 2\alpha(\tau_1^{-1} - 1)g^{-1},$$

$$\delta_6 = \{(1 - \tau_1^{-2})g^{-2} - 1\}/6 \text{ and } \delta_7 = \{1 + (\tau_1^{-1} - 1)g^{-1}\}^2/2.$$

From (5.42) and the lemma 5.1, it then easily follows that,

$$(5.44) \quad {}_2F_1\left(\frac{m}{2} + \alpha, \frac{m\tau_1 g}{2} + \alpha\tau_1; \frac{mg}{2} + \alpha, \frac{2}{m} p\right)$$

$$= e^{\tau_1 \operatorname{tr} p} \lambda_{[1+m^{-1}(2a\alpha+b\delta_1)+m^{-2}(a\delta_2+a(a+1)\delta_3+b\delta_4+c\delta_5+d\delta_6+e\delta_7)+0(m^{-3})]},$$

where the constants a , b , c , d and e are given by

$$a = \tau_1 \operatorname{tr} p; \quad b = \tau_1^2 \operatorname{tr} p^2; \quad c = 2\tau_1^2 \operatorname{tr} p^2 + \tau_1^3 \operatorname{tr} p^2 \operatorname{tr} p,$$

$$(5.45) \quad d = 4 \tau_1^3 \operatorname{tr} p^3 + 3 \tau_1^2 \operatorname{tr} p^2 + 3 \tau_1^2 (\operatorname{tr} p)^2 + \tau_1 \operatorname{tr} p$$

and

$$e = \tau_1^4 (\operatorname{tr} p^2)^2 + 4 \tau_1^3 \operatorname{tr} p^3 + \tau_1^2 [\operatorname{tr} p^2 + (\operatorname{tr} p)^2].$$

Also

$$(5.46) \quad \left| \lambda - \frac{2}{m} p \right|^{\frac{\tau_1 m}{2} + \alpha\tau_1} = \left| \lambda - \frac{2}{m} p \right|^{\tau_1 m/2} \left| \lambda - \frac{2}{m} p \right|^{\alpha\tau_1}$$

and using (5.26) and (5.27) to the factor on the right hand side of (5.46), it can be easily checked that

$$(5.47) \quad \left| \lambda - \frac{2}{m} p \right|^{\frac{\tau_1 m}{2} + \alpha\tau_1} = e^{-\tau_1 \operatorname{tr} p} \lambda_{[1-m^{-1}\delta_8+m^{-2}\delta_9+0(m^{-3})]},$$

where

$$\delta_8 = \tau_1 \{ \operatorname{tr} p^2 + 2\alpha \operatorname{tr} p \}$$

and

$$\delta_9 = \delta_8^2/2 - \tau_1 \{ 4 \operatorname{tr} p^3/3 + 2\alpha \operatorname{tr} p^2 \}.$$

We therefore have

$$(5.48) \quad C_5(t) = [1+m^{-1}(2a\alpha+b\delta_1-\delta_8)+m^{-2}(a\delta_2+a(a+1)\delta_3+b\delta_4+c\delta_5$$

$$+d\delta_6+e\delta_7+\delta_9-2a\alpha\delta_8-b\delta_1\delta_8)+0(m^{-3})]$$

and finally from (5.36), (5.39) and (5.48) we have

$$(5.49) \quad f(t) = g^{-f/2} [1 + m^{-1} \{\alpha_0 + \alpha_1 g^{-1}\} + m^{-2} \{\beta_0 + \beta_1 g^{-1} + \beta_2 g^{-2}\} + O(m^{-3})],$$

where

$$(5.50) \quad \begin{aligned} \alpha_0 &= 2a\alpha + b - \delta_g, \quad \alpha_1 = b(\tau_1^{-1} - 1), \\ \beta_0 &= -A + 2\alpha^2 a^2 + (a-d)/6 - 2\alpha(b-c+a\delta_g) + e/2 + \delta_g - b\delta_g, \\ \beta_1 &= (\tau_1^{-1} - 1) (2c\alpha + e - b\delta_g) \end{aligned}$$

$$\text{and} \quad \beta_2 = A + (\tau_1^{-2} - 1) (a-d)/6 + 2b\alpha(1 - \tau_1^{-1}) + e(\tau_1^{-1} - 1)^2/2.$$

Inverting the characteristic function in (5.49), we have the asymptotic expansion of the distribution of $-2\rho \log W$ in the form,

$$(5.51) \quad \begin{aligned} P(-2\rho \log W \leq z) &= P_r(X_f^2 \leq z) + m^{-1} \{\alpha_0 P(X_f^2 \leq z) \\ &+ \alpha_1 P(X_{f+2}^2 \leq z)\} + m^{-2} \{\beta_0 P(X_f^2 \leq z) + \beta_1 P(X_{f+2}^2 \leq z) \\ &+ \beta_2 P(X_{f+4}^2 \leq z)\} + O(m^{-3}). \end{aligned}$$

6. ASYMPTOTIC EXPANSION OF THE DISTRIBUTION OF Y , $a=1$ AND $b=0$

For this case putting $a = 1$ and $b = 0$ in (2.6) we can easily see that

$$(6.1) \quad E(Y^h) = \frac{\Gamma_p(n/2) \Gamma_p(n_1/2+h)}{\Gamma_p(n/2+h) \Gamma_p(n_1/2)} |\Lambda|^{-n_1/2} {}_2F_1(n/2, n_1/2+h; n/2+h, M).$$

We assume (5.3) and obtain the asymptotic expansion of L_1 where

$$(6.2) \quad L_1 = \sqrt{n} \log (Y/\tau_1^p)$$

in terms of n increasing with τ_1 and τ_2 fixed assuming that $M = [I - \Lambda^{-1}] = (2/n) P$ where P is a fixed matrix. Let $\chi(t)$ be the characteristic function of L_1 . Then

$$(6.3) \quad \chi(t) = E(e^{it L_1}) = C_6(t)C_7(t),$$

where

$$(6.4) \quad C_6(t) = (1/\tau_1)^{it\sqrt{n}p} \Gamma_p(n/2) \Gamma_p(n_1/2+it\sqrt{n}) [\Gamma_p(n/2+it\sqrt{n}) \Gamma_p(n_1/2)]^{-1}$$

and

$$(6.5) \quad C_7(t) = |\Lambda|^{-n_1/2} {}_2F_1(n/2, n_1/2+it\sqrt{n}; n/2+it\sqrt{n}, M).$$

Using the formula (5.18) to each gamma function on the right hand side of (6.4), we have

$$(6.6) \quad C_6(t) = e^{-pT_1 t^2} [1-n^{-1/2}\{fT_1(it)+2pT_2(it)^3/3\} + n^{-1}\{(fT_2+f^2T_1^2/2)(it)^2 + 2p(T_3+fT_1T_2)(it)^4/3+2p^2T_2^2(it)^6/9\}+0(n^{-3/2})],$$

where

$$(6.7) \quad T_1 = (\tau_1^{-1}-1), T_2 = \tau_1^{-2}-1, T_3 = \tau_1^{-3}-1 \text{ and } f = p(p+1)/2.$$

Now using lemma 5.2 to $(n/2)_\kappa$, $(n/2 + it\sqrt{n})_\kappa$ and $(n_1/2 + it\sqrt{n})_\kappa$ we have

$$(6.8) \quad (n/2)_\kappa (n_1/2+it\sqrt{n})_\kappa / (n/2 + it\sqrt{n})_\kappa = (n_1/2)^\kappa [1+n^{-1/2} 2 it^\kappa T_1 + n^{-1}\{\tau_1^{-1} a_1(\kappa)+2(it)^2(k^2 T_1^2 - k T_2)\} + 0(n^{-3/2})].$$

As before let $M = (2/n)P$ where P is a fixed matrix. Then from (6.8) and the lemma 5.1, we have after a little simplification,

$$(6.9) \quad {}_2F_1(n/2, n_1/2 + it\sqrt{n}; n/2 + it\sqrt{n}, 2/n P) \\ = e^{\tau_1 \text{tr} P} [1-n^{-1/2} (it) A_1 + n^{-1} \{(it)^2 A_2 + q\} + O(n^{-3/2})],$$

where

$$(6.10) \quad A_1 = -2T_1 \text{tr}(\tau_1 P), \quad q = \tau_1 \text{tr} P^2$$

$$\text{and} \quad A_2 = 2\{T_1^2[(\text{tr} \tau_1 P)^2 + \text{tr}(\tau_1 P)] - T_2(\tau_1 \text{tr} P)\}.$$

Also from (5.27) we have

$$(6.11) \quad |A|^{-n_1/2} = |I - 2/n P|^{-\tau_1 n/2} = e^{-\tau_1 \text{tr} P} [1-n^{-1} q + O(n^{-3/2})]$$

and thus

$$(6.12) \quad C_7(t) = [1-n^{-1/2} (it) A_1 + n^{-1} (it)^2 A_2 + O(n^{-3/2})].$$

From (6.3), (6.6) and (6.12), we obtain the following asymptotic expansion for $\chi(t)$.

$$(6.13) \quad \chi(t/\sqrt{2pT_1}) = e^{-t^2/2} [1-n^{-1/2} D_1 + n^{-1} D_2 + O(n^{-3/2})],$$

where the coefficients D_1 and D_2 are given by

$$(6.14) \quad D_1 = (2pT_1)^{-\frac{1}{2}} [(it)(A_1 + fT_1) + (3T_1)^{-1} T_2 (it)^3] \text{ and} \\ D_2 = (2pT_1)^{-3} 4p^2 [(it)^2 (fT_2 + f^2 T_1^2/2 + fT_1 A_1 + A_2) T_1^2 \\ + (it)^4 T_1 (T_3 + fT_1 T_2 + A_1 T_2)/3 + (it)^6 T_2^2/18].$$

$$(7.5) \quad C_9(t) = |\lambda|^{-n_1/2} {}_2F_1(n/2, n_1/2; \frac{n}{2} + it\sqrt{n}, \frac{M}{\lambda}).$$

Using the formula (5.17) to each gamma function on the right-hand side of (7.4) we get

$$(7.6) \quad C_8(t) = e^{-pR_1 t^2} [1 - n^{-\frac{1}{2}} \{fR_1(it) + 2pR_2(it)^3/3\} \\ + n^{-1} \{(fR_2 + f^2 R_1^2/2)(it)^2 \\ + 2p(R_3 + fR_1 R_2)(it)^4/3 + 2p^2 R_2^2(it)^6/9\} + O(n^{-3/2})],$$

where coefficients R_1, R_2, R_3 and R_4 are given by

$$(7.7) \quad R_1 = \tau_2^{-1} - 1, \quad R_2 = \tau_2^{-2} - 1, \quad R_3 = \tau_2^{-3} - 1 \quad \text{and} \quad f = p(p+1)/2.$$

Using lemma 5.1 and 5.2, we have proceeding as in section 6

$$(7.8) \quad {}_2F_1(n/2, n_1/2; \frac{n}{2} + it\sqrt{n}, \frac{2}{n} \frac{p}{\lambda}) \\ = e^{\tau_1 \text{tr} \frac{p}{\lambda}} [1 - n^{-\frac{1}{2}} (it)B_1 + n^{-1} \{B_2 + (it)^2 B_3\} + O(n^{-3/2})],$$

where $B_1 = 2(\text{tr} \tau_1 \frac{p}{\lambda})$, $B_2 = \tau_1 \text{tr} \frac{p^2}{\lambda}$ and $B_3 = B_1(4 + B_1)/2$.

Using (7.8) and (6.11), we can write $C_9(t)$ as

$$(7.9) \quad C_9(t) = [1 - n^{-\frac{1}{2}} (it)B_1 + n^{-1} (it)^2 B_3 + O(n^{-3/2})]$$

and thus we have the following asymptotic expansion for $H(t)$.

$$(7.10) \quad H(t/\sqrt{2R_1 p}) = e^{-t^2/2} [1 - n^{-\frac{1}{2}} \beta_1 + n^{-1} \beta_2 + O(n^{-3/2})],$$

CHAPTER II
ON THE DISTRIBUTION OF THE SPHERICITY TEST CRITERION
IN CLASSICAL AND COMPLEX NORMAL POPULATIONS HAVING
UNKNOWN COVARIANCE MATRICES

1. INTRODUCTION AND SUMMARY

Let $x: px1$ be distributed $N(\mu, \Sigma)$ where μ and Σ are both unknown. Let S_N be the sum of product matrix of a sample of size N . To test the hypothesis of sphericity, namely,
 $H_0: \Sigma = \sigma^2 I_p$, where $\sigma^2 > 0$ is unknown, against $H_1: \Sigma \neq \sigma^2 I_p$,
 Mauchly [20] obtained the likelihood ratio test criterion for H_0 in the form $W = |S_N| / [(tr S_N)/p]^p$. Thus the criterion W is a power of the ratio of the geometric mean and the arithmetic mean of the roots $\theta_1, \theta_2, \dots, \theta_p$ of $|S_N - \theta I_p| = 0$ (see Anderson [1]). For $p = 2$,
 Mauchly [20] showed that the density of W is

$$(1.1) \quad f(w) = \frac{1}{2} (n-1) w^{\frac{1}{2}(n-3)}, \quad 0 \leq w \leq 1,$$

where $n = N-1$. The exact distribution in the null case was obtained by Consul [5], [6], in the form

$$(1.2) \quad f(w) = k(p, n) w^{\frac{1}{2}(n-p-1)} G_{p,0}^{p,p} \left(w \left| \begin{matrix} \frac{1}{2}(p-1) + (p-1)/p, \dots, \frac{1}{2}(p-1) \\ 0, \frac{1}{2}, 1, \dots, \frac{1}{2}(p-1) \end{matrix} \right. \right),$$

where

$$k(p,n) = (2\pi)^{\frac{1}{2}(p-1)} \frac{1}{p^{\frac{1}{2} - \frac{1}{2}pn}} \Gamma\left(\frac{1}{2}pn\right) / \prod_{j=1}^p \Gamma\left[\frac{1}{2}(n-j-1)\right],$$

and $G_{p,q}^{m,n}(x \mid \begin{smallmatrix} a_1, \dots, a_p \\ b_1, \dots, b_q \end{smallmatrix})$ is the G-function defined in the next

section.

In this chapter we have obtained the general moments of W both in real and complex cases for arbitrary covariance matrices and also the corresponding distributions of W in terms of G-function.

2. SOME DEFINITIONS AND RESULTS

In this section we give a few definitions and some lemmas which are needed in the sequel. First we define Meijer's G-function by [21]

$$(2.1) \quad G_{p,q}^{m,n}(x \mid \begin{smallmatrix} a_1, \dots, a_p \\ b_1, \dots, b_q \end{smallmatrix}) = (2\pi i)^{-1} \int_C \frac{\prod_{j=1}^m \Gamma(b_j - s) \prod_{j=1}^n \Gamma(1 - a_j + s)}{\prod_{j=m+1}^q \Gamma(1 - b_j + s) \prod_{j=n+1}^p \Gamma(a_j - s)} x^s ds,$$

where an empty product is interpreted as unity and C is a curve separating the singularities of $\prod_{j=1}^m \Gamma(b_j - s)$ from those of $\prod_{j=1}^n \Gamma(1 - a_j + s)$, $q \geq 1$, $0 \leq n \leq p \leq q$, $0 \leq m \leq q$; $x \neq 0$ and $|x| < 1$ if $q = p$; $x \neq 0$ if $q > p$.

The G-function of (2.1) can be expressed as a finite number of generalized hypergeometric functions (see Pillai, Al-Ani and Jouris [25] and Luke [18]) and in particular we have

$$(2.3) \quad G_{2,2}^{2,0}(x |_{b_1, b_2}^{a_1, a_2}) =$$

$$\frac{x^{b_1-1} (1-x)^{a_1+a_2-b_1-b_2-1}}{\Gamma(a_1+a_2-b_1-b_2)} {}_2F_1(a_2-b_2, a_1-b_2, a_1+a_2-b_1-b_2, 1-x)$$

$$0 < x < 1$$

where ${}_2F_1$ here is the Gauss hypergeometric function.

Now we state the Gauss and Legendre's multiplication formula for gamma functions as

$$(2.4) \quad \prod_{r=1}^n \Gamma[z+(r-1)/n] = (2\pi)^{\frac{1}{2}(n-1)} \frac{1}{n^{\frac{1}{2}} - nz} \Gamma(nz).$$

Further, the hypergeometric function of matrix variates is defined by

$${}_pF_q(a_1, \dots, a_p; b_1, \dots, b_q; S, T) = \sum_{k=0}^{\infty} \sum_{\kappa} \frac{(a_1)_{\kappa} \dots (a_p)_{\kappa} C_{\kappa}(S) C_{\kappa}(T)}{(b_1)_{\kappa} \dots (b_q)_{\kappa} C_{\kappa}(I_m) k!}$$

where the zonal polynomials, $C_{\kappa}(\cdot)$, and $(\cdot)_{\kappa}$ are defined in [12].

Lemma 2.1. Let $Z: mxm$ be a complex symmetric matrix whose real part is p.d. and let $T: mxm$ be an arbitrary complex symmetric matrix.

Let $S: mxm$ be a real p.d. matrix. Then

$$(2.5) \quad \int_{S>0} \exp(-\text{tr } Z S) |S|^{t-\frac{1}{2}(m+1)} C_{\kappa}(T S) d S = \Gamma_m(t, \kappa) |Z|^{-t} C_{\kappa}(T Z^{-1})$$

where $\Gamma_m(t, \kappa)$ is defined in (15) of Constantine [4] and $R(t) > \frac{1}{2}(m-1)$.

(See Constantine [4]).

Lemma 2.2. Let T be as in lemma 2.1. Then

$$(2.6) \quad \int_{S > 0} \exp\left(-\frac{1}{2} \text{tr } S\right) |S|^{t-\frac{1}{2}(m+1)} (\text{tr } S)^q C_{\kappa}(T, S) dS \\ = \Gamma_m(t, \kappa) 2^{tm+k+q} \Gamma(mt+k+q) C_{\kappa}(T) / \Gamma(mt+k).$$

Proof. We shall consider the cases when (1) $q \geq 0$ and (2) $q < 0$.

(1) $q \geq 0$. From (2.5), for $u > 0$ we have

$$(2.7) \quad \int_{S > 0} \exp\left(-\frac{1}{2} u \text{tr } S\right) |S|^{t-\frac{1}{2}(m+1)} C_{\kappa}(T, S) dS \\ = 2^{tm+k} u^{-tm-k} \Gamma_m(t, \kappa) C_{\kappa}(T).$$

To prove this case we differentiate (2.7) q times w.r.t. u under the integral sign and let $u = 1$ to obtain

$$(2.8) \quad \int_{S > 0} \exp\left(-\frac{1}{2} \text{tr } S\right) |S|^{t-\frac{1}{2}(m+1)} (\text{tr } S)^q C_{\kappa}(T, S) dS \\ = 2^{tm+k+q} \Gamma_m(t, \kappa) \Gamma(tm+k+q) C_{\kappa}(T) / \Gamma(mt+k).$$

which is also (19) of Khatri [15].

(2) $q < 0$. To prove this case, we integrate (2.7) successively r times w.r.t. u , change the order of integration and let $u = 1$, yielding

$$(2.9) \quad \int_{S > 0} \exp\left(-\frac{1}{2} \text{tr } S\right) |S|^{t-\frac{1}{2}(m+1)} (\text{tr } S)^{-r} C_{\kappa}(T, S) dS \\ = 2^{tm+k-r} \Gamma(tm+k-r) \Gamma_m(t, \kappa) C_{\kappa}(T) / \Gamma(tm+k).$$

Since $\Gamma(tm+k-r)/\Gamma(tm+k) = \prod_{j=1}^n (tm+k-j)^{-1}$, (2.9) holds if

$tm+k-r > 0$. This proves the lemma.

Lemma 2.3. Let Z : $m \times m$ be a complex symmetric matrix whose real part is p.d. and let T : $m \times m$ be an arbitrary complex symmetric matrix and S : $m \times m$ be a Hermitian matrix. Then

$$(2.10) \quad \int_{\tilde{S}'=S>0} \exp(-\text{tr } Z S) |S|^{a-m} \hat{C}_\kappa(T S) dS \\ = \hat{\Gamma}_m^{\nu}(a, \kappa) |Z|^{-a} \hat{C}_\kappa(T),$$

where $\hat{\Gamma}_m^{\nu}(a, \kappa)$ is defined in [12].

Lemma 2.4. Let T and S be as in lemma 2.3. Then

$$(2.11) \quad \int_{\tilde{S}'=S>0} \exp(-\text{tr } S) |S|^{a-m} (\text{tr } S)^j \hat{C}_\kappa(T S) dS \\ = \hat{\Gamma}_m^{\nu}(a, \kappa) \Gamma(am+k+j) \hat{C}_\kappa(T) / \Gamma(am+k).$$

Proof. The proof is exactly similar to the proof of lemma 2.2 and hence omitted.

Lemma 2.5. If s is any complex variate and $f(x)$ is a function of a real variable x , such that

$$F(s) = \int_0^{\infty} x^{s-1} f(x) dx$$

exists, then under certain regularity conditions

$$f(x) = (2\pi i)^{-1} \int_{C-i\infty}^{C+i\infty} x^{-s} F(s) ds.$$

$F(s)$ is called the Mellin transform of $f(x)$, and $f(x)$ is the inverse Mellin transform of $F(s)$. (See Titchmarsh [29])

3. DISTRIBUTION OF W IN THE REAL CASE

Let ξ : $p \times p$ be distributed as Wishart (n, p, ξ) . Then the distribution of the latent roots g_1, g_2, \dots, g_p of ξ has been shown by James [12] to depend only on the latent roots of ξ and is given by

$$(3.1) \quad k(p, n, \xi) |\xi|^{-\frac{1}{2}n} {}_0F_0\left(-\frac{1}{2} \xi^{-1}, G\right) |G|^{\frac{1}{2}(n-p-1)} \prod_{i < j} (g_i - g_j) \prod_{i=1}^p dg_i,$$

where

$$k(p, n, \xi) = |\xi|^{-\frac{1}{2}n} \frac{1}{\pi^{\frac{1}{2}p^2}} / \{2^{\frac{1}{2}pn} \Gamma_p\left(\frac{1}{2}n\right) \Gamma\left(\frac{1}{2}p\right)\},$$

$$G = \text{diag}(g_1, g_2, \dots, g_p), \quad \infty > g_1 \geq g_2 \geq \dots \geq g_p > 0.$$

The distribution (3.1) is not convenient for further development and the convergence of the series is slow. But the convergence may be improved by writing (3.1) in the form suggested by Pillai, Al-Ani and Jouris [25]),

$$(3.2) \quad k(p, n, \xi) |G|^{\frac{1}{2}(n-p-1)} \exp\left(-\frac{1}{2} \text{tr } G\right) \prod_{i < j} (g_i - g_j) {}_0F_0(M, G)$$

where

$$M = \frac{1}{2} (I - \xi^{-1}).$$

Theorem 3.1. Let ξ be distributed as in (3.2) and let $W = |\xi|/\{(\text{tr } \xi)/p\}^p$ be the sphericity criterion. Then the h -th moment of W is given by

$$(3.3) \quad E(W^h) = \frac{p^{ph} |\xi|^{-\frac{1}{2}n}}{2^{\frac{1}{2}pn} \Gamma_p(\frac{1}{2}n)} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{C_{\kappa}(M)}{k!} 2^k \frac{\Gamma_p(\frac{1}{2}n+h, k) \Gamma(\frac{1}{2}pn+k)}{\Gamma(\frac{1}{2}pn+ph+k)}.$$

Proof. To find $E(W^h)$ we multiply (3.2) by $|\xi|/\{(\text{tr } \xi)/p\}^p$, transform $\xi \rightarrow H \chi H'$ where H is an orthogonal and χ a symmetric matrix, integrate out H and χ using (44) and (22) of Constantine [4]. We get

$$(3.4) \quad E(W^h) = p^{ph} k(p, n, \xi) \Gamma_p(\frac{1}{2}p) \pi^{-\frac{1}{2}p^2} \cdot \sum_{k=0}^{\infty} \sum_{\kappa} [C_{\kappa}(M)/C_{\kappa}(I_p)k!]$$

$$\int_{\chi > 0} \exp(-\frac{1}{2} \text{tr } \chi) |\chi|^{\frac{1}{2}(n+h) - \frac{1}{2}(p+1)} (\text{tr } \chi)^{-ph} C_{\kappa}(\chi) d\chi.$$

Applying lemma (2.2) to the integral on the R.H.S. of (3.4) we get (3.3).

Theorem 3.2. For any finite p , the p.d.f. of W is

$$(3.5) \quad f(w) = C(p, n, \xi) \sum_{k=0}^{\infty} \sum_{\kappa} \frac{2^k C_{\kappa}(M)}{k!} p^{\frac{1}{2} - \frac{1}{2}pn-k} \Gamma(\frac{1}{2}pn+k)$$

$$\cdot w^{\frac{1}{2}(n-p-1)} G_{p,p}^{p,0}(w | a_1, \dots, a_p; b_1, \dots, b_p),$$

where

$$C(p, n, \xi) = \pi^{\frac{1}{4}} p(p-1) |\xi|^{-\frac{1}{2}n} (2\pi)^{\frac{1}{2}(p-1)} / \Gamma_p(\frac{1}{2}n),$$

$$a_j = (k+j-1)/p + \frac{1}{2}(p-1); \quad b_j = k_j + \frac{1}{2}(p-j).$$

For $p = 2$, (3.5) reduces to

$$(3.6) \quad f(w) = \frac{|\xi|^{-\frac{1}{2}n}}{2\Gamma(n-1)} w^{\frac{1}{2}(n-3)} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{\Gamma(n+k)}{k!} C_{\kappa}(\mathcal{M}) w^{k_1 + \frac{1}{2}} \cdot {}_2F_1(a_2 - b_2, a_1 - b_2; a_1 + a_2 - b_1 - b_2, 1-w).$$

Proof. Applying (2.4) on $\Gamma[p(\frac{1}{2}n+h+k/p)]$ we have from (3.3)

$$E(W^h) = C(p, n, \xi) \sum_{k=0}^{\infty} \sum_{\kappa} [2^k C_{\kappa}(\mathcal{M}) p^{\frac{1}{2} - \frac{1}{2}pn-k} \Gamma(\frac{1}{2}pn+k)] / k! \prod_{j=1}^p [\Gamma(\frac{1}{2}n+h+k_j - \frac{1}{2}(i-j)) / \Gamma(\frac{1}{2}n + ((k+j-1)/p) + h)].$$

Using Lemma 2.5, the density of W has the form

$$(3.7) \quad f(w) = C(p, n, \xi) \sum_{k=0}^{\infty} \sum_{\kappa} \frac{2^k C_{\kappa}(\mathcal{M})}{k!} p^{\frac{1}{2} - \frac{1}{2}pn-k} \Gamma(\frac{1}{2}pn+k) w^{\frac{1}{2}(n-p-1)}$$

$$\cdot (2\pi i)^{-1} \int_{C-i\infty}^{C+i\infty} w^{-r} \frac{\prod_{i=1}^p \Gamma(r+b_i)}{\prod_{i=1}^p \Gamma(r+a_i)} dr,$$

where

$$r = \frac{1}{2}n+h - \frac{1}{2}(p-1), \quad b_j = k_j + \frac{1}{2}(p-j), \quad a_j = (k+j-1)/p + \frac{1}{2}(p-1).$$

Noting that the integral in (3.7) is in the form of Meijer's G-function, we can write the density of W as in (3.5).

(3.6) can be obtained easily from (3.5) by putting $p = 2$ in (3.5) and using (2.3).

Remark. Putting $\xi = \sigma^2 I_p$ in (3.5) and (3.6), we can easily deduce the result of Consul in (1.2) [5], [6], and Mauchly in (1.1), [20].

4. DISTRIBUTION OF W IN THE COMPLEX CASE

Let ξ : $p \times p$ be distributed as a Complex Wishart (n, p, ξ) (see Goodman [11]). Then the distribution of the latent roots g_1, g_2, \dots, g_p of ξ is (James [12])

$$(4.1) \quad k(p, n, \xi) {}_0F_0(-\frac{\nu-1}{\xi}, \xi) |G|^{n-p} \prod_{i < j} (g_i - g_j)^2 \prod_{i=1}^p dg_i$$

where

$$k(p, n, \xi) = \frac{|\xi|^{-n} \pi^{p(p-1)}}{\Gamma_p(n) \Gamma_p(p)} ; \Gamma_p(n) \text{ and}$$

${}_pF_q(a_1, \dots, a_p; b_1, \dots, b_q; \xi, \tau)$ are defined in (83) and (88) of James [12].

As in the real case, the convergence of (4.1) may be improved by writing it in the form

$$(4.2) \quad k(p, n, \xi) {}_0F_0(\hat{M}_1, \xi) \exp(-\text{tr} \xi) |G|^{n-p} \prod_{i < j} (g_i - g_j)^2 \prod_{i=1}^p dg_i,$$

where $\hat{M}_1 = I_p - \xi^{-1}$.

Theorem 5.1. Let \mathcal{G} be distributed as in (4.2) and let $W = |\mathcal{G}|/[(\text{tr } \mathcal{G})/p]^P$. Then h-th moment of W is

$$(4.3) \quad \frac{p^{ph}}{\tilde{\Gamma}_p(n)} |\Sigma|^{-n} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{\tilde{C}_{\kappa}(\tilde{M}_1)}{k!} \Gamma(np+k) \tilde{\Gamma}_p(n+h, \kappa) / \Gamma(np+k+ph).$$

Proof: Multiplying (4.2) by $|\mathcal{G}|/[(\text{tr } \mathcal{G})/p]^P$, using the transformation $\mathcal{G} \rightarrow \mathcal{U} \mathcal{X} \mathcal{U}'$ where \mathcal{U} is unitary and \mathcal{X} is hermitian p.d. we have on integrating out \mathcal{U} and using the results (see Khatri [14]) that the Jacobian of transformation is

$$J(\mathcal{G}, \mathcal{U} \mathcal{X}) = \prod_{i < j} (g_i - g_j)^2 h_2(\mathcal{U})$$

and that $\int_{\mathcal{U} \mathcal{U}' = I} h_2(\mathcal{U}) = \frac{\pi^p (p-1)}{\tilde{\Gamma}_p(p)}$, we have

$$E(W^h) = \frac{p^{ph} |\tilde{\Sigma}|^{-n}}{\tilde{\Gamma}_p(n)} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{\tilde{C}_{\kappa}(\tilde{M}_1)}{\tilde{C}_{\kappa}(I_p) k!} \int_{V > 0} \exp(-\text{tr } \mathcal{X}) |\mathcal{X}|^{n+h-p} (\text{tr } \mathcal{X})^{-ph} \tilde{C}_{\kappa}(\mathcal{X}) d\mathcal{X}$$

Using lemma (2.4) to the integral on the right, we get (4.3).

Theorem 5.2. The density of W is

$$f(w) = \frac{\pi^{\frac{1}{2} p(p-1)} |\tilde{\Sigma}|^{-n} (2\pi)^{\frac{1}{2} (p-1)}}{\tilde{\Gamma}_p(n)} \sum_{k=0}^{\infty} \sum_{\kappa} \frac{\tilde{C}_{\kappa}(\tilde{M}_1)}{k!} \Gamma(pn+k)$$

$$p^{\frac{1}{2} - pn - k} w^{n-p} G_{p,p}^{p,0} \left(w \mid \begin{matrix} a_1, a_2, \dots, a_p \\ b_1, \dots, b_p \end{matrix} \right)$$

where $a_j = (k/p) + (j-1)/p + (p-1)$, and $b_j = k_j - j + p$.

Proof. The proof is exactly similar to that of theorem 3.2 and hence is omitted.

CHAPTER III

THE DISTRIBUTION OF THE SPHERICITY TEST CRITERION
UNDER THE NULL HYPOTHESIS1. INTRODUCTION

1 Let \underline{x} : $p \times 1$ be distributed $N(\underline{\mu}, \underline{\Sigma})$ where $\underline{\mu}$ and $\underline{\Sigma}$ are both unknown. Let \underline{S} be the sum of product matrix of a sample of size N . To test the hypothesis of sphericity, namely, $H_0: \underline{\Sigma} = \sigma^2 \underline{I}_p$, where $\sigma^2 > 0$ is unknown, against $\underline{\Sigma} \neq \sigma^2 \underline{I}_p$, Mauchly [20] obtained the likelihood ratio test criterion for H_0 in the form $W = |\underline{S}| / [(\text{tr } \underline{S})/p]^p$. For $p = 2$, Mauchly [20], showed that the density of W is

$$(1.1) \quad f(w) = \frac{1}{2} \binom{n-1}{n-3} w^{\frac{1}{2}(n-3)}, \quad 0 \leq w \leq 1,$$

where $n = N-1$. The exact distribution in the null case was given by Consul in [5] for some special values of p and in the closed form in [6] in terms of Meijer's G-function, while its non-null distribution is obtained in Chapter II. However the forms of the distribution of W obtained by Consul and later by Mathai and Rathie [17] are not quite suited for computational purposes. No systematic attempt seems to have been made so far to compute the exact percentage points of W . The approximate percentage points for

$p = 3$ have been obtained by Mauchly [20] by fitting a Pearson curve of the form

$$(1.2) \quad y = k x^{p-1} (1-x)^{q-1}$$

and more recently by Davis [8] for $p = 3, 6, 10$ and $n = 4(1)8, 10, 12, 15, 20$ using a Cornish-Fisher inversion of Box's Series.

The object of the present paper is to develop methods similar to the ones used by Box [2] and U. S. Nair [23], [24], in order to obtain the exact distribution of W in series form and to compute exact percentage points of W . In particular, methods have been given which yield facility in computation for the cases when the sample size is small as well as when sample size is large. Tabulations of percentage points for $p = 2(1)10$ for various significance levels are given and comparisons made with approximate values using (1) Box's series, [Anderson [1], page 263], (2) Mauchly [20], (3) Tukey and Wilks [30], and (4) Davis [8].

2. EXACT DISTRIBUTION OF THE SPHERICITY CRITERION W .

For ease in computation, it is desirable to give methods which are particularly suited for extremely small values of N , the sample size, and those which are suited for larger values of N . Thus exact percentage points of W can then be computed for all values of N . In parts (a) and (b) of this section, we shall consider two methods which will achieve the former objective

while in part (c) we shall give a method which will achieve the latter objective. The first method which we shall now consider makes use of Mellin transform and Contour integration.

(a) Exact distribution of W through Contour Integration

It has been shown by Mauchly [20], that the h -th moment of the sphericity criterion $W = |\xi|/[(\text{tr } \xi)/p]^P$ is given by

$$(2.1) \quad E(W^h) = p^{ph} \prod_{i=1}^p \left[\frac{\Gamma\{\frac{1}{2}(N-i)+h\}}{\Gamma\{\frac{1}{2}(N-i)\}} \frac{\Gamma\{\frac{1}{2} p(N-1)\}}{\Gamma\{\frac{1}{2} p(N-1)+ph\}} \right].$$

Using Mellin's transform, the density of W is given by

$$(2.2) \quad f(w) = K(p,n) \frac{1}{2\pi i} \int_{-i\infty}^{i\infty} \frac{w^{-h-1} p^{ph} \prod_{i=1}^p \Gamma\{\frac{1}{2}(N-i)+h\}}{\Gamma\{\frac{1}{2} pn+ph\}} dh,$$

where $n = N-1$ and $K(p,n) = \Gamma(\frac{1}{2} pn) / \prod_{i=1}^p \Gamma(N-i)/2$. Putting $\frac{1}{2}(N-p)+h=s$

in (2.2), we have

$$(2.3) \quad f(w) = K(p,n) p^{-\frac{1}{2} p(N-p)} w^{\frac{1}{2}(N-p)-1} \cdot p(w),$$

where

$$(2.4) \quad p(w) = \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} (w/p^p)^{-s} \left[\prod_{i=1}^p \Gamma(s + \frac{p-i}{2}) / \Gamma_p(s + \frac{p-1}{2}) \right] ds,$$

and $c = \frac{1}{2}(N-p)$.

Special Cases. (i) $p = 2$. For $p = 2$, we have

$$(2.5) \quad p(w) = \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} [(w/4)^{-s} \Gamma(s) \Gamma(s + \frac{1}{2}) / \Gamma(2s+1)] ds.$$

Using the following duplicating formula for the gamma function

$$(2.6) \quad \Gamma(s)\Gamma\left(s + \frac{1}{2}\right) = (\pi)^{\frac{1}{2}} \Gamma(2s)/2^{2s-1},$$

in (2.5) we have

$$(2.7) \quad p(w) = \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} (\pi)^{\frac{1}{2}} [w^{-s}/s] ds .$$

The pole of the integrand is at $s = 0$ and the corresponding residue is then $(\pi)^{1/2}$. Hence from (2.3) we have for $p = 2$, the density of W as in (1.1) obtained by Mauchly.

(ii) $p = 3$. In this case, we have from (2.4)

$$(2.8) \quad p(w) = \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} (w/27)^{-s} [\Gamma(s+1)\Gamma\left(s+\frac{1}{2}\right)\Gamma(s)/\Gamma(3s+3)] ds.$$

Using the duplication formula (2.6), we have

$$(2.9) \quad p(w) = 2(\pi)^{1/2} \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} (4w/27)^{-s} [\Gamma(2s)\Gamma(s+1)/\Gamma(3s+3)] ds.$$

We shall again evaluate the integral in (2.9) by Contour integration.

The poles of the integrand are at the points

$$(2.10) \quad s = -m/2, \quad m = 0, -1, -2, \dots,$$

and the residue at these points can be found by putting $s = t - \frac{1}{2}m$ in the integrand and taking the residue of the integrand at $t = 0$.

Thus substituting $s = t - \frac{m}{2}$, the integrand in (2.9) becomes

$$(2.11) \quad (4w/27)^{-t+\frac{m}{2}} \Gamma(2t-m)\Gamma\left(t+\frac{m}{2}\right)/\Gamma\left\{3t-\frac{3}{2}m+3\right\} .$$

To evaluate the integral we need to consider separately the cases when (i) m is even and (ii) m is odd.

First let m be even and say $m = 2r$. Then the integrand reduces to

$$(2.12) \quad (4w/27)^{-t+r} \Gamma(2t-2r)\Gamma(t-r+1)/\Gamma(3t-3r+3),$$

which by expanding the gamma functions becomes

$$(2.13) \quad \frac{3}{2}(4w/27)^{-t+r} \frac{\Gamma(2t+1)\Gamma(t+1) \left[\prod_{i=1}^{3r-3} (3t-i) \right]}{t\Gamma(3t+1) \left[\prod_{i=1}^{2r} (2t-i) \prod_{i=1}^{r-1} (t-i) \right]},$$

valid for $r > 1$ and the cases when $r = 0$ and $r = 1$ have to be considered separately. The expression in (2.13) has a simple pole of first order at $t = 0$ and its residue at this point is clearly given by

$$(2.14) \quad \frac{3}{2}(4w/27)^r (3r-3)!/(2r)!(r-1)! .$$

For $r = 0$, the integrand reduces to

$$(4w/27)^{-t} \Gamma(2t)\Gamma(t+1)/\Gamma(3t+3),$$

which can be written as

$$(2.15) \quad (4w/27)^{-t} \Gamma(2t+1)\Gamma(t+1)/(2t)\Gamma(3t+3)$$

and has a simple pole at $t = 0$, the residue at this point being $\frac{1}{4}$.

For $r = 1$, the integrand after a little simplification becomes,

$$(2.16) \quad \frac{3}{2}(4w/27)^{-t+1} \Gamma(2t+1) \Gamma(t+1) / (t)(2t-1)(t-1) \Gamma(3t+1)$$

and this has a simple pole at $t = 0$, the corresponding residue being equal to $\frac{3}{2}(4w/27)$.

Now if m is odd say equal to $2q+1$, where q is an integer or zero, then as before, the integrand (2.11) can be easily written down as in (2.13) in the form

$$(2.17) \quad (4w/27)^{-t+q+\frac{1}{2}} \frac{\Gamma(2t+1) \Gamma(t+\frac{1}{2}) \prod_{i=1}^{3q-1} (3t+\frac{1}{2}-i)}{(2t) \left[\prod_{i=1}^{2q+1} (2t-i) \prod_{i=1}^q (t+\frac{1}{2}-i) \right] \Gamma(3t+\frac{1}{2})},$$

which clearly holds for $q > 0$; the case $q = 0$ has to be treated separately.

For $q > 0$, (2.17) has a pole of first order at $t = 0$ and its residue at this point is

$$(4w/27)^{q+\frac{1}{2}} \frac{\prod_{i=1}^{3q-1} (\frac{1}{2}-i)}{[2 \prod_{i=1}^{2q+1} (-i) \prod_{i=1}^q (\frac{1}{2}-i)]},$$

which can be written in an alternate form as

$$(2.18) \quad (4w/27)^{q+\frac{1}{2}} \Gamma(3q-\frac{1}{2}) / 2 \Gamma(2q+2) \Gamma(q+\frac{1}{2}).$$

For $q = 0$, the integrand reduces after a little simplification to

$$(2.19) \quad (4w/27)^{-t+\frac{1}{2}} \Gamma(2t+1) \Gamma(t+\frac{1}{2}) / (2t)(2t-1) \Gamma(3t+\frac{3}{2}),$$

and this has a pole of first order at $t = 0$, the corresponding residue being equal to $-(4w/27)^{1/2}$. Hence finally using Cauchy's Residue Theorem, the integral in (2.8) is seen to be equal to

$$(2.20) \quad p(w) = 2(\pi)^{1/2} \left[\frac{1}{4} - (4w/27)^{1/2} + \sum_{r=1}^{\infty} \frac{\frac{3}{2}(4w/27)^r \Gamma(3r-2)}{\Gamma(2r+1)\Gamma(r)} \right. \\ \left. + \sum_{q=1}^{\infty} \frac{(4w/27)^{\frac{1}{2}(2q+1)} \Gamma(3q-\frac{1}{2})}{2\Gamma(2q+2)\Gamma(q+\frac{1}{2})} \right].$$

From (2.3), the density of W for $p = 3$ is therefore

$$(2.21) \quad f(w) = \left[\frac{\Gamma\{3(N-1)/2\}}{\prod_{i=1}^3 \Gamma(\frac{N-i}{2})} \right] w^{\frac{1}{2}(N-3)-1} p(w)^{\frac{3}{2}(N-3)}$$

where $p(w)$ is as in (2.20). John [13] has recently given an explicit form for the density of W for $p = 3$ but not in a very convenient form for use.

(iii) $p \geq 4$. The cases for values of $p \geq 4$ can be treated in almost a similar way but the method involves psi functions and their derivatives and makes use of the following lemma due to Nair [23].

Lemma 2.1. Let (a_i) be a sequence of numbers, finite or infinite, and let

$$(2.22) \quad F(x;t;a_2, a_3, \dots) \equiv e^{xt+a_2 \frac{t^2}{2!} + a_3 \frac{t^3}{3!} + \dots}$$

Then the n th derivative of $F(x;t;a_2, a_3, \dots)$ at $t = 0$ is

$$(2.23) \quad D_n(x, a) = \begin{vmatrix} x & -1 & 0 & 0 & 0 & \dots & 0 \\ a_2 & x & -1 & 0 & \dots & \dots & 0 \\ a_3 & 2a_2 & x & -1 & \dots & \dots & 0 \\ a_4 & 3a_3 & 3a_2 & x & -1 & \dots & 0 \\ \dots & \dots & \dots & \dots & \dots & \dots & \dots \\ a_n & \binom{n-1}{1}a_{n-1} & \binom{n-1}{2}a_{n-2} & \dots & \dots & \dots & x \end{vmatrix}.$$

Case (1). p odd. First let $p = 2k+1$, ($k > 1$) be odd and let us denote the integrand in (2.4) by $G(s)$. Then it is easy to see that

$$(2.24) \quad G(s) = A(W_1)^{-s} \Gamma(s+k) \Gamma(2s) \prod_{i=1}^{k-1} \Gamma(2s+2i) / \Gamma\{p(s+k)\},$$

where

$$(2.25) \quad A = (\sqrt{\pi})^k 2^{-k(k-2)} \quad \text{and} \quad W_1 = 2^{2k} w/p^p.$$

The poles of $G(s)$ are at the points given in (2.10) and the residue at these points is equal to the residue of $G(t - \frac{m}{2})$ at $t = 0$. Now

$$(2.26) \quad G(t - \frac{m}{2}) = A(W_1)^{-t + \frac{m}{2}} \Gamma(t - \frac{m}{2} + k) \Gamma(2t - m) \prod_{i=1}^{k-1} \Gamma(2t - m + 2i) \\ \cdot [\Gamma\{p(t - \frac{m}{2}) + kp\}]^{-1}.$$

We have to consider the cases when (1) m is even and (2) m is odd.

First let $m = 2r$ be even. Then we have

$$(2.27) \quad G(t-r) = A p W_1^r \cdot C(t) / (2t)^k$$

where

$$(2.28) \quad C(t) = \frac{(W_1)^{-t} \Gamma(t+1) [\Gamma(2t+1)]^k \prod_{i=1}^{p(r-k)} (j-pt) / \prod_{j=1}^{2r} (j-2t)}{\prod_{i=1}^{r-k} (j-t) \Gamma(pt+1) \prod_{i=1}^{k-1} \prod_{j=1}^{2r-2i} (j-2t)}$$

Thus for $r > k$, the pole of $G(t-r)$ is of order k and the residue R_r at $t = 0$ is

$$(2.29) \quad R_r = [A_p W_1^r 2^{-k} / \Gamma(k)] \left(\frac{d}{dr} \right)_{t=0}^{k-1} e^{\log C(t)}$$

Using the formula (Erdélyi; [10])

$$(2.30) \quad \log \Gamma(x+a) = \log \Gamma(a) + x \psi(a) + \frac{x^2}{2!} \psi_1(a) + \frac{x^3}{3!} \psi_2(a) + \dots$$

where

$$(2.31) \quad \psi(a) = \frac{d}{dx} \log \Gamma(x) \Big|_{x=a} \quad \text{and} \quad \psi_j(a) = \left(\frac{d}{dx} \right)^j \psi(x) \Big|_{x=a}$$

$\log C(t)$ can be written as

$$(2.32) \quad \log C(t) = b_0 + b_1 t + b_2 \frac{t^2}{2!} + \dots$$

where

$$(2.33) \quad b_0 = \log \left[\frac{p(r-k)!}{(r-k)! \prod_{i=0}^{k-1} (2r-2i)!} \right],$$

$$b_1 = (1+2k-p)\psi(1) + p \sum_{j=1}^{p(r-k)} (1/j) - \sum_{j=1}^{r-k} (1/j) \\ - 2 \sum_{i=0}^{k-1} \sum_{j=1}^{2r-2i} (1/j) - \log W_1$$

and

$$b_q = (1+k2^{q-p^q})\psi_{q-1}(1)+\Gamma(q) \left[\sum_{j=1}^{p(r-k)} (p/j)^q - \sum_{i=0}^{k-1} \sum_{j=1}^{2r-2i} (2/j)^q - \sum_{j=1}^{r-k} (1/j)^q \right], \quad q = 2, 3, \dots$$

Using (2.32) in (2.29) and then applying lemma 2.1, we have

$$(2.34) \quad R_r = \frac{A_p(W_1^r) [p(r-k)!]}{2^k \Gamma(k) (r-k)! \prod_{i=0}^{k-1} (2r-2i)!} D_{k-1}(W_1, b), \quad (r > k)$$

where

$$(2.35) \quad D_{k-1}(W_1, b) = \begin{vmatrix} b_1 & -1 & 0 & 0 & 0 & 0 \\ b_2 & b_1 & -1 & 0 & \dots & 0 \\ b_3 & 2b_2 & b_1 & -1 & \dots & 0 \\ b_4 & 3b_3 & 3b_2 & b_1 & -1 & 0 \\ \dots & \dots & \dots & \dots & \dots & \dots \\ b_{k-1} & \binom{k-2}{1} b_{k-2} & \binom{k-2}{2} b_{k-3} & \dots & \dots & b_1 \end{vmatrix}$$

and b_q 's are as in (2.33).

If $r = 0$, then it can be shown that $G(t)$ has a simple pole at $t = 0$ and the residue R at this point is

$$(2.36) \quad R = A \Gamma(k) \prod_{i=1}^{k-1} \Gamma(2i) / \{2\Gamma(kp)\}.$$

For $r = \ell$ where $\ell = 1, 2, \dots, k-1$, we have from (2.27)

$$(2.37) \quad G(t-\ell) = \frac{A W_1^{-t+\ell} \Gamma\{t+(k-\ell)\} [\Gamma(2t+1)]^{(\ell+1)} \prod_{i=\ell+1}^{k-1} \Gamma\{2t+(2i-2\ell)\}}{(2t)^{\ell+1} \prod_{j=1}^{2\ell} (2t-j) \Gamma\{pt+p(k-\ell)\} \prod_{i=1}^{\ell-1} \prod_{j=1}^{2\ell-2i} (2t-j)}$$

where $\prod_{i=n}^m (\cdot)$ is interpreted as unity if $n > m$.

Thus for $r = \ell$, $\ell = 1, 2, \dots, k-1$, $G(t-\ell)$ has a pole of order $(\ell+1)$ at $t = 0$ and its residue B_ℓ , on using the lemma and proceeding as above is easily seen to be

$$(2.38) \quad B_\ell = \frac{A(W_1)^\ell \Gamma(k-\ell) \prod_{i=\ell+1}^{k-1} \Gamma(2i-2\ell)}{2^{\ell+1} \Gamma(\ell+1) \Gamma\{p(k-\ell)\} \left[\prod_{i=0}^{\ell-1} (2\ell-2i)! \right]} D_\ell(W_1, c), \ell = 1, 2, \dots, k-1,$$

where $D_\ell(W_1, c)$ is equal to the determinant on the right hand side of (2.23) with x replaced by c_1 , n by ℓ and a_q 's by c_q 's, $q = 2, 3, \dots, \ell$. The coefficients c_q 's are given by

$$(2.39) \quad c_1 = \psi(k-\ell) + 2[(\ell+1)\psi(1) + \sum_{i=\ell+1}^{k-1} \psi(2i-2\ell) - \sum_{i=0}^{\ell-1} \sum_{j=1}^{2\ell-2i} (1/j)] \\ - p\psi(pk-p\ell) - \log W_1$$

and

$$c_q = \psi_{q-1}(k-\ell) + 2^q [(\ell+1)\psi_{q-1}(1) + \sum_{i=\ell+1}^{k-1} \psi_{q-1}(2i-2\ell) \\ - \sum_{i=0}^{\ell-1} \sum_{j=1}^{2\ell-2i} (\Gamma(q)/j^q)] - p^q \psi_{q-1}(pk-p\ell), \quad q = 2, 3, \dots,$$

where $\sum_{i=n}^m (\cdot)$ is interpreted as zero if $n > m$.

Similarly for $r = k$, $G(t-k)$ has a pole of order k at $t = 0$ and the residue R_k is

$$(2.40) \quad R_k = \frac{A_p(W_1/2)^k}{\Gamma(k) \prod_{i=0}^{k-1} (2k-2i)!} D_{k-1}(W_1, d)$$

where $D_{k-1}(W_1, d)$ can be obtained from (2.35) by replacing b_q 's by d_q 's where d_q 's are given by

$$(2.41) \quad d_1 = (1+2k-p)\psi(1) - \log W_1 - \sum_{i=0}^{k-1} \sum_{j=1}^{2k+2i} (2/j) \quad \text{and}$$

$$d_q = (1+2^q k - p^q)\psi_{q-1}(1) - \sum_{i=0}^{k-1} \sum_{j=1}^{2k+2i} (2^q \Gamma(q)/j^q), q=2,3,\dots$$

Now let $m = 2q+1$ be odd. Then it can be easily checked that for $q > 0$, $G(t - q - \frac{1}{2})$ has a pole of order k at $t = 0$ and its residue G_q is given by

$$(2.42) \quad G_q = \frac{(-1)^k A W_1^{q + \frac{1}{2}} \Gamma(k - \frac{1}{2}) \prod_{j=1}^{pq} (j + \frac{p}{2} - pk)}{2^k \Gamma(k) \Gamma(kp - \frac{p}{2}) \prod_{j=1}^q (j - k + \frac{1}{2})} D_{k-1}(W_1, f), \quad q > 0,$$

where $D_{k-1}(W_1, f)$ is the determinant equal to the right hand side of (2.35) with b_n 's replaced by f_n 's where f_n 's are given by

$$(2.43) \quad f_1 = -\log W_1 + \psi(k - \frac{1}{2}) + 2k\psi(1) + \sum_{j=1}^{pq} \{p/(j + \frac{p}{2} - pk)\} \\ - p\psi(kp - \frac{p}{2}) - \sum_{\ell=0}^{k-1} \sum_{j=1}^{2q-2\ell+1} (2/j), \quad \text{and}$$

$$f_n = \psi_{n-1}(k - \frac{1}{2}) + 2k\psi_{n-1}(1) - p^n \psi_{n-1}(pk - \frac{k}{2}) + \sum_{j=1}^{pq} [p^n \Gamma(n) / (j + \frac{p}{2} - pk)^n] \\ - \sum_{i=0}^{k-1} \sum_{j=1}^{2q-2i+1} (\frac{2^n \Gamma(n)}{j^n}), \quad n = 2, 3, \dots$$

Similarly if $q = 0$, then $G(t - \frac{1}{2})$ has a simple pole at $t = 0$ and its residue B is

$$(2.44) \quad B = -A(W_1)^{1/2} \Gamma(k - \frac{1}{2}) \prod_{i=1}^{k-1} (2i-1)! / \{2\Gamma p(k - \frac{1}{2})\} ,$$

Thus if p is odd, we have from (2.3) and Cauchy's Residue theorem that the density of W is

$$(2.45) \quad f(w) = k(p,n)p^{-\frac{1}{2}} w^{p(N-p)-1} \left[\sum_{r=k+1}^{\infty} R_r + R + \sum_{\ell=1}^{k-1} B_{\ell} + R_k + \sum_{q=1}^{\infty} G_q + B \right] .$$

Case (2). p is even. Let $p = 2k$ ($k > 1$) be even. Then the integrand $H(s)$ in (2.4) can be written as

$$(2.46) \quad H(s) = A(W_1)^{-s} \Gamma(2s) \prod_{i=1}^{k-1} \Gamma(2s+2i) / \Gamma(ps+pk-k) ,$$

where A and W_1 are as in (2.25). The poles of $H(s)$ are at the points given in (2.10) and the residue of $H(s)$ at these points is equal to the residue of $H(t - \frac{m}{2})$ at $t = 0$. Now

$$(2.47) \quad H(t - \frac{m}{2}) = A(W_1)^{-t + \frac{m}{2}} \Gamma(2t-m) \prod_{i=1}^{k-1} \Gamma(2t-m+2i) / \Gamma\{pt - \frac{pm}{2} + k(p-1)\} .$$

We have to consider separately the cases when m is even and m is odd. When $m = 2r$, then proceeding as before it is seen that for $r \geq k$, $H(t-r)$ has a pole of order $k-1$ and the residue D_r at $t = 0$ is given by

$$(2.48) \quad D_r = \frac{(-1)^{k(p-1)} A_p(W_1)^r \Gamma\{pr-k(p-1)+1\}}{2^k \Gamma(k-1) \prod_{i=0}^{k-1} (2r-2i)!} V_{k-2}(W_1, g), \quad r \geq k,$$

when the determinant $V_{k-2}(W_1, g)$ is similar to the determinant on the right hand side of (2.35) having $(k-2)$ rows and the elements b_q 's being replaced by g_q 's where g_q 's are given by

$$(2.49) \quad g_1 = (2k-p)\psi(1)-p \sum_{j=1}^{pr-k(p-1)} (1/j) - \sum_{r=0}^{k-1} \sum_{j=1}^{2r-2i} (2/j) - \log W_1$$

$$g_q = (k2^q - p^q)\psi_{q-1}(1) + \Gamma(q) \left[\sum_{j=1}^{pr-k(p-1)} (p/j)^q - \sum_{i=0}^{k-1} \sum_{j=1}^{2r-2i} (2/j)^q \right],$$

$$q = 2, 3, \dots, \dots$$

For $r = 0$, $H(t)$ has a simple pole at $t = 0$ and the residue D at this point is

$$(2.50) \quad D = A \prod_{i=1}^{k-1} \Gamma(2i)/2\Gamma k(p-1).$$

For $r = \ell$ where $\ell = 1, 2, \dots, k-1$, $H(t-\ell)$ has a pole of order $\ell+1$ at $t = 0$ and the residue E_ℓ is given by

$$(2.51) \quad E_\ell = \frac{A(W_1)^\ell \prod_{i=\ell+1}^{k-1} \Gamma(2i-2\ell) / \prod_{i=0}^{\ell-1} (2\ell-2i)!}{2^{\ell+1} \Gamma(\ell+1) \Gamma\{p(k-\ell)-k\}} V_\ell(W_1, h) \quad \ell=1, 2, \dots, k-1,$$

where $V_\ell(W_1, h)$ is the determinant of ℓ th order similar to that in (2.35) with b_q 's being replaced by h_q 's where h_q 's are given by

$$(2.52) \quad h_1 = -\log W_1 + 2[(\ell+1)\psi(1) + \sum_{i=\ell+1}^{k-1} \psi(2i-2\ell) - \sum_{i=0}^{\ell-1} \sum_{j=1}^{2\ell-2i} (1/j)] \\ - p\psi(pk-p\ell-k)$$

and

$$h_q = 2^q [(\ell+1)\psi_{q-1}(1) + \sum_{i=\ell+1}^{k-1} \psi_q(2i-2\ell) - \sum_{i=0}^{\ell-1} \sum_{j=1}^{2\ell-2i} (\Gamma(q)/j^q)] \\ - p^q \psi_{q-1}(pk-p\ell-k), \quad q = 2, 3, \dots$$

Now let $m = 2q + 1$ be odd. Then for $q \geq k-1$, $H(t-q-\frac{1}{2})$ has a pole of order $(k-1)$ at $t = 0$ and the residue F_q is

$$(2.53) \quad F_q = \frac{A_p(W_1)^{q+\frac{1}{2}} [p(q-k+1)!]}{2^k \Gamma(k-1) \sum_{i=0}^{k-1} (2q+1-2i)!} V_{k-2}(W_1, K), \quad q \geq k-1,$$

where $V_{k-2}(W_1, K)$ is the determinant of order $k-2$ similar to that in (2.35) with the elements given by

$$(2.54) \quad k_1 = -\log W_1 + (2k-p)\psi(1) + \sum_{j=1}^{p(q-k+1)} (p/j) - \sum_{i=0}^{k-1} \sum_{j=1}^{2q+1-2i} (2/j),$$

and

$$k_n = (k2^n - p^n) \psi_{n-1}(1) + \Gamma(n) \left[\sum_{j=1}^{p(q-k+1)} (p/j)^n - \sum_{i=0}^{k-1} \sum_{j=1}^{2q+1-2i} (2/j)^n \right],$$

$$n = 2, 3, \dots$$

where

$\sum_{j=n}^m (\cdot)$ is interpreted as zero if $m < n$.

For $q = \ell$, $\ell = 1, 2, \dots, k-2$, $H(t - \ell - \frac{1}{2})$ has a pole of order $(\ell+1)$ at $t = 0$ and the corresponding residue G_ℓ is

$$(2.55) \quad G_\ell = \frac{A(W_1)^{\ell + \frac{1}{2}} \prod_{i=\ell+1}^{k-1} \Gamma(2i-1+2\ell)}{2^{\ell+1} \Gamma(\ell+1) \prod_{i=0}^{\ell} (2\ell+1-2i)! \Gamma p(k-1-\ell)} V_\ell(W_1, m),$$

where as before $V_\ell(W, m)$ is the determinant of order ℓ similar to the one in (2.35), with elements given by

$$(2.56) \quad m_1 = -\log W_1 + 2(\ell+1)\psi(1) + 2 \sum_{i=\ell+1}^{k-1} \psi(2i-1-2\ell) \\ - \sum_{i=0}^{\ell} \sum_{j=1}^{2\ell+1-2i} (2/j) - p\psi(pk-p-p\ell)$$

and

$$m_n = 2^n [(\ell+1)\psi_{n-1}(1) + \sum_{i=\ell+1}^{k-1} \psi_{n-1}(2i-1-2\ell) - \sum_{i=0}^{\ell} \sum_{j=1}^{2\ell+1-2i} (\Gamma(n)/j^{n0})] \\ - p^n \psi_{n-1}(pk-p-p\ell) \quad n = 2, 3, \dots$$

Finally for $q = 0$, $H(t - \frac{1}{2})$ has a simple pole at $t = 0$ and the residue C is

$$(2.57) \quad C = -A(W_1)^{\frac{1}{2}} \prod_{i=1}^{k-1} \Gamma(2i-1) / 2\Gamma p(k-1).$$

Thus when p is even, the density of W is given by

$$(2.58) \quad f(w) = k(p,n)p^{-\frac{1}{2}p(N-p)} w^{\frac{1}{2}(N-p)-1} [D+C$$

$$+ \sum_{r=k}^{\infty} D_r + \sum_{\ell=1}^{k-1} E_{\ell} + \sum_{q=k-1}^{\infty} F_q + \sum_{\ell=1}^{k-2} G_{\ell}] .$$

(b) Distribution of W as a gamma series

We shall now obtain the distribution of W in gamma series form. For this let

$$(2.59) \quad L = W^{\frac{n}{2}} .$$

Then from (2.1) we have,

$$(2.60) \quad E(L^h) = K(p,n)p^{\frac{np}{2}} \prod_{\alpha=1}^p \frac{\Gamma\{\frac{n}{2}(1+h)+\frac{1-\alpha}{2}\}}{\Gamma\{\frac{1}{2}pn(1+h)\}} .$$

Now let $\lambda = -2q \log L$ where q is an adjustable constant which can be chosen so as to govern the rate of convergence of the resulting gamma series and $0 < q < \infty$. If $\phi(t)$ is the characteristic function λ then

$$(2.61) \quad \phi(t) = K(p,n) \cdot C(t)$$

where

$$(2.62) \quad C(t) = p^{-np} e^{it} q \prod_{\alpha=1}^p \frac{\Gamma\{\frac{n}{2}(1-2q it)+\frac{1-\alpha}{2}\}}{\Gamma\{\frac{pn}{2}(1-2q it)\}}$$

and therefore

$$(2.63) \quad \log \phi(t) = \log\{K(p,n)\} - np it q \log p - \log \Gamma\{\frac{pn}{2}(1-2q it)\}$$

$$+ \sum_{\alpha=1}^p \log \Gamma\{\frac{n}{2}(1-2q it)+\frac{1-\alpha}{2}\} .$$

The expansion of $\log \phi(t)$ will be based on the following expansion for the gamma function:

$$(2.64) \quad \log \Gamma(x+h) = \frac{1}{2} \log(2\pi) + (x+h-\frac{1}{2}) \log x - x - \sum_{r=1}^m \frac{(-1)^r B_{r+1}(h)}{r(r+1)x^r} + R_{m+1}(x),$$

where $R_m(x)$ is the remainder such that $|R_m(x)| \leq \theta/|x^m|$, θ a constant independent of x , and $B_r(h)$ the Bernoulli polynomial of degree r and order one defined by

$$\frac{\tau e^{h\tau}}{e^\tau - 1} = \sum_{r=0}^{\infty} \frac{\tau^r}{r!} B_r(h).$$

Explicitly the polynomials are

$$B_0(h) = 1, \quad B_1(h) = h - \frac{1}{2}, \quad B_2(h) = h^2 - h + \frac{1}{6}$$

and in general

$$B_r(h) = h^r - \frac{1}{2} r h^{r-1} + {}_r C_2 B_1 h^{r-2} - {}_r C_4 B_2 h^{r-4} + \dots$$

(last term is x or a constant)

$$= h^r - \frac{1}{2} r h^{r-1} + \sum_{m=1}^{\leq r/2} (-1)^{m-1} {}_r C_{2m} B_m h^{r-2m}.$$

where ${}_n C_m = n!/(n-m)!m!$, B_m are the Bernoulli numbers and have been tabulated extensively. Using (2.64) we obtain

$$\begin{aligned}
(2.65) \quad \log \phi(t) &= \log\{K(p,n)\} + \left(\frac{p-1}{2}\right) \log(2\pi) \\
&- \left(\frac{p^2+p-2}{4}\right) \log \{n(1-2q \text{ it})/2\} \\
&- \left(\frac{pn-1}{2}\right) \log p + \sum_{r=1}^m (Q_r/n^r) \left(\frac{1-2q \text{ it}}{2}\right)^{-r} \\
&+ R'_{m+1}(n,t) \quad ,
\end{aligned}$$

where the coefficients Q_r 's are given by

$$Q_r = (-1)^{r-1} \left[\left(\sum_{\alpha=1}^p B_{r+1} \left(\frac{1-\alpha}{2}\right) \right) - B_{r+1}(0)/p^r \right] / r(r+1) \quad .$$

The characteristic function of L can then be obtained from (2.65) as

$$\begin{aligned}
(2.66) \quad \phi(t) &= K_1(p,n) [n(1-2q \text{ it})/2]^{-v} \left(\sum_{j=0}^{\infty} (B_j/n^j) \left(\frac{1-2q \text{ it}}{2}\right)^{-j} \right) \\
&+ R''_{m+1}(n,t) \quad ,
\end{aligned}$$

where

$$\begin{aligned}
K_1(p,n) &= K(p,n) (2\pi)^{\left(\frac{p-1}{2}\right)} p^{-(pn-1)/2} \quad , \\
v &= (p^2 + p-2)/4 \quad ,
\end{aligned}$$

and the coefficients B_j 's which we need in our computations are listed below:

$$\begin{aligned}
B_0 &= 1, & B_1 &= Q_1, & B_2 &= Q_1^2/2 + Q_2, \\
B_3 &= Q_1 Q_2 + Q_3 + Q_1^3/6, \\
B_4 &= Q_2 Q_1^2/2 + Q_1^4/24 + Q_4 + Q_3 Q_1 + Q_2^2/2, \\
B_5 &= Q_1^3 Q_2/6 + Q_2^2 Q_1/2 + Q_1^5/120 + Q_3 Q_1^2/2 + Q_5 + Q_3 Q_2 + Q_4 Q_1, \\
B_6 &= Q_1 Q_2 Q_3 + Q_1^6/720 + Q_4 Q_1^2/2 + Q_1^4 Q_2/24 + Q_3^2/2 + Q_1^2 Q_2^2/4 + Q_4 Q_2 \\
&\quad + Q_1^3 Q_3/6 + Q_5 Q_1 + Q_2^3/6 + Q_6, \\
B_7 &= Q_1^4 Q_3/24 + Q_3^2 Q_1/2 + Q_4 Q_2 Q_1 + Q_1 Q_2^3/6 + Q_1^2 Q_2 Q_3/2 + Q_5 Q_1^2/2 \\
&\quad + Q_1^3 Q_4/6 + Q_1^7/5040 + Q_5 Q_2 + Q_2^2 Q_3/2 + Q_1^3 Q_2^2/12 + Q_7 \\
&\quad + Q_6 Q_1 + Q_1^5 Q_2/120 + Q_4 Q_3, \\
B_8 &= Q_8 + 3Q_3 B_5/8 + Q_6 Q_2/4 + Q_1 B_7/8 + Q_5 Q_1 Q_2/4 + Q_4 B_4/2 + Q_4 Q_2^2/4 \\
&\quad + 5Q_5 B_3/8 + Q_2 Q_3^2/8 + 3Q_6 B_2/4 + Q_4 Q_2 Q_1^2/8 + Q_1 Q_2^2 Q_3/4 + Q_2^4/24 \\
&\quad + Q_1^3 Q_3 Q_2/24 + Q_1^2 Q_3^3/16 + Q_1^4 Q_2^2/96 + 7Q_7 Q_1/8 + Q_1^6/720, \\
B_9 &= Q_9 + 7Q_7 B_2/9 + Q_1 Q_2 Q_3^2/3 + Q_1^6 Q_3/2160 + 24Q_1^2 Q_3/6 + 8Q_8 Q_1/9 \\
&\quad + 2Q_6 B_3/3 + Q_1^4 Q_2 Q_3/72 + Q_3^3/6 + 5Q_5 B_4/9 + Q_1^2 Q_2^2 Q_3/12 \\
&\quad + Q_4 Q_2 Q_3/3 + Q_1^3 Q_3^2/18 + 4Q_4 B_5/9 + Q_5 Q_1 Q_3/3 + Q_1 B_8/9 \\
&\quad + Q_2^3 Q_3/18 + 2Q_2 B_7/9 + Q_6 Q_3/3,
\end{aligned}$$

and

$$\begin{aligned}
B_{10} &= Q_{10} + 7Q_7 Q_1 Q_2/10 + 3Q_6 B_4/5 + 7Q_7 Q_3/10 + 4Q_8 B_2/5 + 9Q_9 Q_1/10 \\
&\quad + 7Q_7 Q_1^3/60 + 3Q_3 B_7/10 + Q_5 Q_1^3 Q_2/12 + Q_2^2 Q_5 Q_1/4 + Q_1^5 Q_5/240 \\
&\quad + Q_2 B_8/5 + Q_3 Q_1^2 Q_5/4 + Q_1 B_9/10 + Q_5^2/2 + Q_1 Q_4 Q_5/2 + Q_2 Q_3 Q_5/2.
\end{aligned}$$

The rest of the coefficients are given in appendix A.

(c) Distribution of W as a Beta Series

In the sequel we shall need the following theorem, restated from Nair [24].

Theorem 2.1. Let

$$\phi(t) = \int x^t p(x) dx$$

be the moment function of a random variable x with distribution law p(x). If

$$\phi(t) = O(t^{-k})$$

with real part of t tending to ∞ , then $\phi(t)$ can be expanded as a factorial series of the form

$$\phi(t) = \sum_{n=0}^{\infty} a_n \Gamma(t+a)/\Gamma(t+k+n+a)$$

a being an arbitrary non-negative constant.

We shall now obtain the distribution of W. Putting $\frac{1}{2}(N-\lambda)+h=s$ in (2.2), where λ is an adjustable constant which can be chosen to govern the rate of convergence, we have

$$(2.72) \quad f(w) = k(p, N, w, \lambda) \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} (w/p^p)^{-s} \frac{\prod_{i=1}^p \Gamma(s + \frac{\lambda-i}{2})}{\Gamma_p(s + \frac{\lambda-1}{2})} ds,$$

where

$$(2.73) \quad k(p, N, w, \lambda) = \frac{\Gamma\{\frac{1}{2} p(n-1)\}}{\prod_{i=1}^p \Gamma(\frac{N-i}{2})} p^{-\frac{1}{2} p(N-\lambda)} w^{\frac{1}{2}(N-\lambda)-1}$$

Using the expansion (2.64) to each gamma functions involved in the integrand of the right hand side of (2.72), we have after some simplification

$$(2.74) \quad \log \left[\prod_{i=1}^p \Gamma\left(s + \frac{\lambda-i}{2}\right) / \Gamma_p\left(s + \frac{\lambda-1}{2}\right) \right] \\ = \log \left[(2\pi)^{(p-1)/2} s^{-v} / p^{ps + (p\lambda-p-1)/2} \right] + \sum_{r=1}^{\infty} (q_r / s^r)$$

where the coefficients q_i 's are given by

$$(2.75) \quad q_r = (-1)^{r-1} \left[\prod_{i=1}^p B_{r+1}\left(\frac{\lambda-i}{2}\right) - p^{-r} B_{r+1}\left(p(\lambda-1)/2\right) \right] / r(r+1)$$

and

$$v = (p^2 + p - 2) / 4 .$$

From (2.74), we deduce that

$$(2.76) \quad \prod_{i=1}^p \Gamma\left(s + \frac{\lambda-i}{2}\right) / \Gamma_p\left(s + \frac{\lambda-1}{2}\right) \\ = (2\pi)^{(p-1)/2} s^{-v} \left[1 + \sum_{r=1}^{\infty} (B_r / s^r) \right] / p^{ps + (p\lambda-p-1)/2},$$

where the coefficients B_r 's are as given in (2.67) with Q_r 's on the right hand side of (2.67) being replaced by q_r 's. Now from (2.72) and (2.76) we have the density of W as

$$(2.77) \quad f(w) = K_1(p, n) w^{\frac{1}{2}(N-\lambda)-1} \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} w^{-s} s^{-v} \left[1 + \sum_{r=1}^{\infty} B_r / s^r \right] ds$$

where $K_1(p, n)$ is as given in (2.66).

The integral on the right hand side of (2.77) can be easily computed if v is an integer and its value is by Cauchy's theorem of Residues, the residue of $w^{-s} s^{-v} [1 + \sum_{r=1}^{\infty} B_r / s^r]$ at $s = 0$. This is easily seen to be equal to

$$(2.78) \quad \sum_{r=0}^{\infty} [(-\log w)^{v+r-1} B_r / \Gamma(v+r)], \quad B_0 = 1$$

and thus from (2.77), the density of W is

$$(2.79) \quad f(w) = K_1(p, n) \sum_{r=0}^{\infty} (B_r) w^{\frac{1}{2}(N-\lambda)-1} (-\log w)^{v+r-1} / \Gamma(v+r).$$

The probability that W is less than any value, say x_0 , is

$$(2.80) \quad P(W \leq x_0) = K_1(p, n) \sum_{r=0}^{\infty} B_r \int_0^{x_0} w^{\frac{1}{2}(N-\lambda)-1} (-\log w)^{v+r-1} dw / \Gamma(v+r).$$

For computational purposes, we let

$$(2.81) \quad I_{v+r-1, u} = \int_0^{x_0} w^u (-\log w)^{v+r-1} dw / \Gamma(v+r),$$

where $u = \frac{1}{2}(N-\lambda)-1$. Then integrating by parts the R.H.S. of (2.81) it can be easily checked that the following extremely useful recurrence relation holds:

$$(2.82) \quad I_{v+r-1, u}(x_0) = [x_0^{(u+1)} (-\log x_0)^{v+r-1} / \Gamma(v+r) + I_{v+r-2, u} / (u+1)]$$

and
$$I_{0, u}(x_0) = x_0^{(u+1)} / (u+1).$$

With this notation, (2.80) can be rewritten as

$$(2.83) \quad P(W \leq x_0) = K_1(p, n) \sum_{r=0}^{\infty} B_r I_{v+r-1, u}(x_0)$$

where $I_{v+r-1, u}(x_0)$ satisfies the recurrence relation (2.82). It is to be noted that (2.83) holds only if v is an integer. However if v is not an integer, we can appeal to Theorem 2.1, since in this case

$$(2.84) \quad \phi(s) = s^{-v} [1 + \sum_{r=1}^{\infty} B_r / s^r] = O(s^{-v}).$$

Thus according to Theorem 2.1, we can expand $\phi(s)$ in the factorial series as

$$(2.85) \quad s^{-v} [1 + \sum_{r=1}^{\infty} B_r / s^r] = \sum_{i=0}^{\infty} R_i \Gamma(s) / \Gamma(s+v+i)$$

where the coefficients R_i 's can be determined explicitly as is done below.

Using the formula (2.64) to each gamma function on the right hand side of (2.85) we have

$$(2.86) \quad \log \{ \Gamma(s) / \Gamma(s+v+i) \} = [-(v+i) \log s] + \sum_{j=1}^{\infty} (C_{ij} / s^j)$$

where the first few coefficients C_{ij} 's which we need for our computations are given below. Let $t_i = v + i$.

$$\begin{aligned}
C_{i1} &= -\frac{1}{2}(v+i)(v+i-1) = -t_i(t_i-1)/2 \\
C_{i2} &= (t_i^3 - 3t_i^2/2 + t_i/2)/6 \\
C_{i3} &= -(t_i^4 - 2t_i^3 + t_i^2)/12 \\
C_{i4} &= (t_i^5 - 5t_i^4/2 + 5t_i^3/3 - t_i/6)/20 \\
(2.87) \quad C_{i5} &= -(t_i^6 - 3t_i^5 + 5t_i^4/2 - t_i^2/2)/30 \\
C_{i6} &= (t_i^7 - 7t_i^6/2 + 7t_i^5/2 - 7t_i^3/6 + t_i/6)/42 \\
C_{i7} &= -(t_i^8 - 4t_i^7 + 14t_i^6/3 - 7t_i^4/3 + 2t_i^2/3)/56 \\
C_{i8} &= (t_i^9 - 9t_i^8/2 + 6t_i^7 - 21t_i^5/5 + 2t_i^3 - 3t_i/10)/72 \\
C_{i9} &= -(t_i^{10} - 5t_i^9 + 15t_i^8/2 - 7t_i^6 + 5t_i^4 - 3t_i^2/2)/90
\end{aligned}$$

and

$$C_{i10} = (t_i^{11} - 11t_i^{10}/2 + 55t_i^9/6 - 11t_i^7 + 11t_i^5 - 11t_i^3/2 + 5t_i/6)/110 .$$

The remaining coefficients are given in appendix A.

Thus from (2.86) we have

$$(2.88) \quad \Gamma(s)/\Gamma(s+v+i) = s^{-(v+i)} \left(1 + \sum_{j=1}^{\infty} (d_{ij}/s^j) \right) ,$$

where the coefficients d_{ij} 's can be obtained in terms of C_{ij} from (2.67) where we replace B_j by d_{ij} and Q_j by C_{ij} . Using (2.88) on the right hand side of (2.85) we have finally

$$(2.89) \quad s^{-v} \left[1 + \sum_{r=1}^{\infty} (B_r/s^r) \right] = \sum_{i=0}^{\infty} R_i s^{-(v+i)} \left[1 + \sum_{j=1}^{\infty} d_{ij}/s^j \right] .$$

Equating the coefficients of s on both sides of (2.89), it is easy to check that we have the following explicit relations to determine the coefficients R_i 's.

$$\begin{aligned}
 R_0 &= 1, \quad R_1 + R_0 d_{01} = B_1, \\
 R_2 + R_1 d_{11} + R_0 d_{02} &= B_2, \\
 R_3 + R_2 d_{21} + R_1 d_{12} + R_0 d_{03} &= B_3, \\
 R_4 + R_3 d_{31} + R_2 d_{22} + R_1 d_{13} + R_0 d_{04} &= B_4, \\
 R_5 + R_4 d_{41} + R_3 d_{32} + R_2 d_{23} + R_1 d_{14} + R_0 d_{05} &= B_5, \\
 R_6 + R_5 d_{51} + R_4 d_{42} + R_3 d_{33} + R_2 d_{24} + R_1 d_{15} + R_0 d_{06} &= B_6, \\
 (2.90) \quad R_7 + R_6 d_{61} + R_5 d_{52} + R_4 d_{43} + R_3 d_{34} + R_2 d_{25} + R_1 d_{16} + R_0 d_{07} &= B_7, \\
 R_8 + R_7 d_{71} + R_6 d_{62} + R_5 d_{53} + R_4 d_{44} + R_3 d_{35} + R_2 d_{26} + R_1 d_{17} \\
 &\quad + R_0 d_{08} = B_8, \\
 R_9 + R_8 d_{81} + R_7 d_{72} + R_6 d_{63} + R_5 d_{54} + R_4 d_{45} + R_3 d_{36} + R_2 d_{27} \\
 &\quad + R_1 d_{18} + R_0 d_{09} = B_9,
 \end{aligned}$$

and

$$\begin{aligned}
 R_{10} &= R_9 d_{91} + R_8 d_{82} + R_7 d_{73} + R_6 d_{64} + R_5 d_{55} + R_4 d_{46} + R_3 d_{37} \\
 &\quad + R_2 d_{28} + R_1 d_{19} + R_0 d_{010} = B_{10}.
 \end{aligned}$$

Now using (2.85) in (2.77) and noting that term by term integration is valid since a factorial series is uniformly convergent in a half-plane (see Doetch [9]) we have the density of W in the case that v is not an integer in the form

$$(2.91) \quad f(w) = K_1(p, n) \sum_{i=0}^{\infty} R_i \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} w^{-s} [\Gamma(s)/\Gamma(s+v+i)] ds$$

and on using the well known integral (Titchmarsh [29])

$$(2.92) \quad \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} x^{-s} \{\Gamma(s)/\Gamma(s+v)\} ds = (1-x)^{v-1}/\Gamma(v) \quad 0 \leq x \leq 1$$

$$c > 0$$

the exact distribution of $f(w)$ is in the form of a beta series given by

$$(2.93) \quad f(w) = K_1(p,n) \sum_{i=0}^{\infty} R_i w^{\frac{1}{2}(N-\lambda)-1} (1-w)^{v+i-1}/\Gamma(v+i),$$

where the coefficients R_i 's are as given in (2.90).

The distribution of W is then given by

$$(2.94) \quad P(W \leq x_0) = K_1(p,n) \sum_{i=0}^{\infty} R_i I_{x_0}^{\left(\frac{1}{2}(N-\lambda), v+i\right)}/\Gamma(v+i),$$

where $I_{x_0}^{(p,q)}$ is the incomplete beta function $\int_0^{x_0} x^{p-1}(1-x)^{q-1} dx$.

3. APPROXIMATIONS TO THE DISTRIBUTIONS OF W

This section is devoted to approximations.

Approximation using the method of Wilks and Tukey.

Let $(a)_h = a(a+1)(a+2)\dots(a+h-1)$. Then we have

$$\Gamma(a+h) = (a)_h \Gamma(a),$$

$$\text{and} \quad \Gamma(a+rh) = (a)_{rh} \Gamma(a) = \Gamma(a) r^{rh} \prod_{i=1}^r \left(\frac{a+i-1}{r}\right)_h,$$

where r is a positive integer. With these notations, (2.1) can be written as

$$(3.1) \quad E(W^h) = \prod_{i=1}^p \left(\frac{n}{2} + \frac{i-p}{2}\right)_h / \left(\frac{n}{2} + \frac{i-1}{p}\right)_h$$

and this can be put in the form (see eq. (6) of Wilks and Tukey [30])

$$(3.2) \quad E(W^h) = \prod_{i=1}^p \left(\frac{1}{x} - A_i + 1 \right)_h / \left(\frac{1}{x} - B_i + 1 \right)_h$$

where $x = \frac{2}{n}$, $A_i = 1 - \frac{i-p}{2}$ and $B_i = 1 - (i-1)/p$. We can therefore apply the method of Wilks and Tukey [30], to find a fractional power of the test criterion which is approximately distributed according to an incomplete beta distribution function (Pearson Type I)

$$(3.3) \quad dF(u) = \Gamma(\alpha+\beta) [\Gamma(\alpha)\Gamma(\beta)]^{-1} u^{\alpha-1} (1-u)^{\beta-1} du .$$

According to this method, the appropriate values of α , β and the exponent r of the criterion are given by the solutions of the following equations:

$$(3.4) \quad (\alpha+\beta)/r = \frac{1}{x} = \frac{n}{2}, \quad C_1 = \beta \quad \text{and} \quad r = \beta(\beta+1)/(C_2-\beta) ,$$

where

$$C_m = \sum_{i=1}^p (A_i^m - B_i^m) .$$

Using the values of A_i and B_i from (3.2), we have

$$C_1 = (p^2+p-2)/4 \quad \text{and} \quad C_2 = (2p^3+9p^2+5p-12)/24-1/6p .$$

We have the following table of values of α , β and r for various values of p (α being calculated from rounded values of r):

Table 2. Values of α , β and γ for various values of p

p	r	r (rounded)	α	β
2	2	2	$n-1$	1
3	2.74	3	$(3n-5)/2$	2.5
4	3.47	3	$(3n-9)/2$	4.5
5	4.21	4	$2n-7$	7
6	4.95	5	$5(n-4)/2$	10
7	5.69	6	$(6(n-27))/2$	13.5
8	6.43	6	$(6n-35)/2$	17.5
9	7.17	7	$(7n-44)/2$	22
10	7.92	8	$4n-27$	27

Approximation using Box's Series

Following the procedure of Box [2], we have (see Anderson [1], page 263) the following asymptotic expansion for W :

$$(3.5) \quad \Pr\{-np \log W \leq z\} = \Pr\{x_f^2 \leq z\} + w_2 (\Pr\{x_{f+4}^2 < z\} - \Pr\{x_f^2 \leq z\}) + O(n^{-3})$$

where $\rho = 1 - (2p^2 + p + 2) / (6pn)$

and $w_2 = (p+2)(p-1)(p-2)(2p^3 + 6p^2 + 3p + 2) / (288p^2 n^2 \rho^2)$.

Mauchly's Approximation

Mauchly [20] has computed approximate percentage points of W , only for $p = 3$ by fitting a Pearson curve of the type

$$y = K x^{p-1} (1-x)^{q-1}$$

by adjusting p and q so as to obtain agreement with the first two moments of the actual distribution.

Davis Approximation

Davis [8] has obtained percentile approximations for $-\log W$, then ρ is the same as in (3.5), by means of a Cornish-Fisher Inversion of Box's series, expressing the percentage points of the distribution in terms of chi-squared percentiles.

Comparisons of the accuracy of these four approximations are carried out in the next section.

4. COMPUTATIONS USING SERIES FORMS AND THE APPROXIMATIONS

Some of the cases studied are summarized in Tables 3 - 5. Table 3 gives 5% and 1% points for the exact distribution of W , together with the percentage points as approximated by Mauchly's, Box's series and Wilks-Tukey's approximations for various values of p and N . Table 4 gives the .005, .01, .025, .05, .1 and .25 significance points of the exact distribution of W for $p = 3(1)10$ and various N . Table 5 gives comparison with Davis' approximations.

Table 3 reveals that even for moderate sample size N , the approximations given by Mauchly for $p = 3$ is extremely poor. Box's series approximation is reasonably good for small values of p and even moderate values of N . Davis' results are generally correct to the decimal he has given but his table is incomplete in regard to small values of N for the values of p he has considered, i.e. $p = 3, 6$ and 10 .

Table 3. Percentage Points for W from the Exact Distribution
and Approximations: Mauchly, Wilks & Tukey and Box Series

p = 3						
	N = 8		N = 10		N = 12	
	5%	1%	5%	1%	5%	1%
Mauchly	.172	.083	.278	.165	.366	.243
Wilks & Tukey	.14780	.07355	.24194	.14434	.32344	.21458
Box Series	.14050	.06843	.23576	.13912	.31842	.20989
Exact	.14026	.06815	.23564	.13898	.31836	.20981
p = 4						
	N = 6		N = 10		N = 15	
Wilks & Tukey	.0 ² ₂ 274	.0 ³ ₃ 46	.09040	.04563	.24680	.16636
Box Series	.0 ² ₂ 471	.0 ² ₂ 102	.09789	.05061	.25365	.17229
Exact	.0 ² ₂ 38662	.0 ³ ₃ 69040	.097393	.050095	.25352	.17211
p = 5						
	N = 7		N = 10		N = 14	
Wilks & Tukey	.0 ² ₂ 159	.0 ³ ₃ 33	.03103	.01373	.11354	.06852
Box Series	.0 ² ₂ 186	.0 ³ ₃ 44	.03192	.01428	.11497	.06957
Exact	.0 ² ₂ 12621	.0 ³ ₃ 21839	.031104	.013613	.11460	.069151
p = 6						
	N = 8		N = 12		N = 14	
Wilks & Tukey	.0 ³ ₃ 86	.0 ³ ₃ 20	.02610	.01264	.05238	.02922
Box Series	.0 ³ ₃ 77	.0 ³ ₃ 19	.02502	.01206	.05106	.02835
Exact	.0 ³ ₃ 42669	.0 ³ ₃ 71870	.024325	.011478	.050510	.027821
p = 7						
	N = 9		N = 11		N = 15	
Wilks & Tukey	.0 ³ ₃ 45	.0 ³ ₃ 11	.0 ² ₂ 415	.0 ² ₂ 159	.02972	.01620
Box Series	.0 ³ ₃ 33	.0 ⁴ ₄ 9	.0 ² ₂ 345	.0 ² ₂ 132	.02765	.01491
Exact	.0 ³ ₃ 14730	.0 ⁴ ₄ 24239	.0 ² ₂ 29501	.0 ² ₂ 10165	.027115	.014444
p = 8						
	N = 10		N = 12		N = 14	
Wilks & Tukey	.0 ⁴ ₄ 9	.0 ⁴ ₄ 2	.0 ² ₂ 129	.0 ³ ₃ 45	.0 ² ₂ 551	.0 ² ₂ 246
Box Series	.0 ³ ₃ 15	.0 ⁴ ₄ 4	.0 ² ₂ 155	.0 ³ ₃ 59	.0 ² ₂ 606	.0 ² ₂ 279
Exact	.0 ⁴ ₄ 51489	.0 ⁵ ₅ 83064	.0 ² ₂ 12329	.0 ³ ₃ 41204	.0 ² ₂ 56126	.0 ² ₂ 24756
p = 9						
	N = 12		N = 14		N = 16	
Wilks & Tukey	.0 ³ ₃ 22	.0 ⁴ ₄ 7	.0 ² ₂ 151	.0 ³ ₃ 61	.0 ² ₂ 493	.0 ² ₂ 236
Box Series	.0 ³ ₃ 25	.0 ⁴ ₄ 8	.0 ² ₂ 155	.0 ³ ₃ 65	.0 ² ₂ 496	.0 ² ₂ 241
Exact	.0 ³ ₃ 13971	.0 ⁴ ₄ 35438	.0 ² ₂ 12945	.0 ³ ₃ 49428	.0 ² ₂ 46163	.0 ² ₂ 21595
p = 10						
	N = 12		N = 14		N = 16	
Wilks & Tukey	.0 ⁴ ₄ 28	.0 ⁵ ₅ 7	.0 ³ ₃ 34	.0 ³ ₃ 12	.0 ² ₂ 151	.0 ³ ₃ 67
Box Series	.0 ⁴ ₄ 29	.0 ⁵ ₅ 9	.0 ³ ₃ 31	.0 ³ ₃ 12	.0 ² ₂ 141	.0 ³ ₃ 63
Exact	.0 ⁵ ₅ 64552	.0 ⁵ ₅ 10036	.0 ³ ₃ 21066	.0 ⁴ ₄ 66814	.0 ² ₂ 12140	.0 ³ ₃ 50647

Table 4. Percentage Points of Sphericity Criterion W

p = 2

N \ α	.005	.01	.025	.05	.1	.25
3	.0 ⁴ ₂ 25000	.0 ³ ₁ 0000	.0 ³ ₆ 2500	.0 ² ₂ 25000	.010000	.062500
4	.0 ² ₅ 0000	.010000	.025000	.050000	.10000	.25000
5	.029240	.046416	.08550	.13572	.21544	.39685
6	.070711	.10000	.15811	.22361	.31623	.50000
7	.12011	.15849	.22865	.30171	.39811	.57435
8	.17100	.21544	.29240	.36840	.46416	.62996
9	.22007	.26827	.34855	.42489	.51795	.67295
10	.26591	.31623	.39764	.47287	.56234	.70711
11	.30808	.35938	.44054	.51390	.59948	.73487
12	.34657	.39811	.47818	.54928	.63096	.75786
13	.38162	.43288	.51135	.58003	.65793	.77720
14	.41352	.46416	.54074	.60696	.68129	.79370
15	.44258	.49239	.56693	.63073	.70170	.80793
16	.46912	.51795	.59038	.65184	.71969	.82034
17	.49340	.54117	.61149	.67070	.73564	.83124
18	.51567	.56234	.63058	.68766	.74989	.84090
19	.53616	.58171	.64792	.70297	.76270	.84951
20	.55505	.59945	.66373	.71687	.77526	.85724
22	.58870	.63096	.69150	.74113	.79433	.87055
24	.61775	.65793	.71509	.76160	.81113	.88159
26	.64305	.68129	.73535	.77908	.82540	.89090
28	.66527	.70170	.75295	.79418	.83768	.89885
30	.68492	.71969	.76836	.80736	.84834	.90572
34	.71810	.74989	.79409	.82925	.86596	.91700
38	.74501	.77426	.81470	.84668	.87992	.92587
42	.76727	.79433	.83157	.86089	.89125	.93303
46	.78597	.81113	.84563	.87269	.90063	.93893
50	.80191	.82540	.85753	.88265	.90852	.94387
60	.83302	.85317	.88056	.90186	.92367	.95332
70	.85570	.87333	.89718	.91566	.93452	.96005
80	.87297	.88862	.90975	.92606	.94267	.96508
90	.88655	.90063	.91958	.93418	.94901	.96898
100	.89751	.91030	.92748	.94069	.95410	.97210
120	.91411	.92491	.93939	.95049	.96172	.97678
140	.92609	.93544	.94794	.95751	.96718	.98011
160	.93513	.94337	.95438	.96279	.97127	.98261
180	.94221	.94957	.95940	.96690	.97446	.98454
200	.94789	.95455	.96342	.97019	.97701	.98609
250	.95817	.96354	.97069	.97613	.98160	.98888
300	.96507	.96957	.97555	.98010	.98467	.99074

Table 4 (Continued)

p = 3

$N \backslash \alpha$.005	.01	.025	.05	.1	.25
4	.0 ⁵ 39305	.0 ⁴ 15228	.0 ⁴ 99478	.0 ³ 40104	.0 ² 16700	.011603
5	.0 ² 11700	.0 ² 23667	.0 ² 61070	.012679	.026853	.076732
6	.0 ² 88748	.014398	.027585	.045683	.076928	.16044
7	.025882	.037466	.061687	.090921	.13590	.24004
8	.050467	.068151	.10225	.14026	.19471	.31002
9	.079827	.10285	.14486	.18921	.24970	.37019
10	.11161	.13898	.18696	.23564	.29971	.42176
11	.14418	.17494	.22726	.27876	.34471	.46613
12	.17647	.20981	.26516	.31836	.38503	.50453
13	.20786	.24391	.30048	.35457	.42118	.53800
14	.23799	.27457	.33321	.38762	.45365	.56738
15	.26666	.30417	.36350	.41779	.48290	.59335
16	.29383	.33192	.39149	.44538	.50934	.61645
17	.31948	.35789	.41737	.47065	.53332	.63712
18	.34366	.38219	.44133	.49386	.55516	.65571
19	.36644	.40492	.46355	.51522	.57511	.67251
20	.38789	.42619	.48417	.53493	.59340	.68778
22	.42713	.46482	.52124	.57006	.62573	.71444
24	.46203	.49889	.55354	.60040	.65338	.73694
26	.49319	.52908	.58190	.62684	.67729	.75618
28	.52111	.55598	.60696	.65006	.69816	.77281
30	.54624	.58007	.62926	.67060	.71651	.78732
34	.58958	.62136	.66715	.70529	.74730	.81144
38	.62556	.65540	.69811	.73343	.77210	.83066
42	.65584	.68391	.72386	.75670	.79248	.84634
46	.68166	.70811	.74559	.77626	.80953	.85736
50	.70393	.72891	.76417	.79293	.82400	.87035
60	.74809	.76997	.80064	.82546	.85211	.89155
70	.78086	.80028	.82737	.84918	.87249	.90679
80	.80612	.82356	.84779	.86723	.88794	.91828
90	.82617	.84199	.86390	.88143	.90006	.92725
100	.84247	.85694	.87693	.89289	.90981	.93444
120	.86737	.87972	.89672	.91024	.92454	.94527
140	.88548	.89624	.91103	.92276	.93513	.95303
160	.89925	.90878	.92186	.93221	.94312	.95886
180	.91006	.91861	.93034	.93961	.94936	.96340
200	.91877	.92654	.93716	.94555	.95436	.96704
250	.93462	.94092	.94952	.95629	.96340	.97361
300	.94529	.95059	.95781	.96350	.96945	.97799

Table 4 (Continued)

p = 4

N	α					
	.005	.01	.025	.05	.1	.25
5	.0 ⁶ 91162	.0 ⁵ 36645	.0 ⁴ 23265	.0 ⁴ 95283	.0 ³ 40030	.0 ² 29305
6	.0 ³ 33678	.0 ³ 69040	.0 ² 18194	.0 ² 38662	.0 ² 84730	.026147
7	.0 ² 30556	.0 ² 50312	.0 ² 99040	.016868	.029512	.066529
8	.010209	.015033	.025485	.038664	.060019	.11410
9	.022162	.030463	.047058	.066398	.095554	.16287
10	.038208	.050095	.072584	.097393	.1396	.20994
11	.057311	.072583	.10033	.12972	.17030	.25404
12	.078477	.096785	.12902	.16211	.20651	.29477
13	.10089	.12183	.15780	.19381	.24102	.33213
14	.12391	.14708	.18610	.22435	.27358	.36631
15	.14708	.17211	.21356	.25352	.30412	.39756
16	.17006	.19663	.23999	.28119	.33269	.42615
17	.19263	.22044	.26528	.30736	.35936	.45236
18	.21462	.24343	.28938	.33205	.38425	.47643
19	.23595	.26553	.31230	.35332	.40749	.49860
20	.25655	.28673	.33406	.37723	.42920	.51905
22	.29546	.32641	.37429	.41734	.46850	.55550
24	.33132	.36261	.41046	.45301	.50304	.58698
26	.36428	.39559	.44305	.48484	.53356	.61440
28	.39455	.42567	.47247	.51337	.56068	.63847
30	.42235	.45313	.49912	.53903	.58492	.65977
34	.47149	.50132	.54542	.58326	.62635	.69571
38	.51337	.54207	.58415	.61995	.66039	.72486
42	.54938	.57689	.61695	.65082	.68883	.74894
46	.58059	.60692	.64506	.67712	.71293	.76918
50	.60788	.63307	.66939	.69978	.73359	.78641
60	.66298	.68558	.71790	.74471	.77429	.82004
70	.70468	.72509	.75409	.77801	.80425	.84454
80	.73729	.75584	.78211	.80366	.82721	.86318
90	.76346	.78045	.80441	.82401	.84536	.87784
100	.78491	.80057	.82259	.84055	.86007	.88966
120	.81798	.83149	.85042	.86580	.88244	.90756
140	.84225	.85413	.87072	.88415	.89865	.92046
160	.86083	.87141	.88617	.89809	.91094	.93021
180	.87550	.88504	.89832	.90904	.92057	.93782
200	.88737	.89606	.90814	.91787	.92832	.94394
250	.90906	.91615	.92599	.93390	.94238	.95501
300	.92375	.92974	.93804	.94470	.95183	.96243

Table 4 (Continued)

p = 5

N	α					
	.005	.01	.025	.05	.1	.25
6	.0 ⁶ 24579	.0 ⁶ 98368	.0 ⁵ 72524	.0 ⁴ 25776	.0 ³ 10959	.0 ³ 83762
7	.0 ³ 10563	.0 ³ 21839	.0 ³ 58374	.0 ² 12621	.0 ² 28373	.0 ² 92522
8	.0 ² 10968	.0 ² 18281	.0 ² 36768	.0 ² 64001	.011530	.027554
9	.0 ² 40994	.0 ² 61227	.010628	.016501	.026388	.053105
10	.0 ² 97579	.013613	.021543	.031104	.046080	.082916
11	.018156	.024161	.035852	.049192	.069047	.11473
12	.0290262	.037303	.052770	.069704	.093963	.14705
13	.041953	.052479	.071536	.091741	.11983	.17893
14	.056485	.069151	.091503	.11460	.14594	.20983
15	.072206	.086848	.11215	.13775	.17180	.23944
16	.088751	.10518	.13309	.16082	.19710	.26760
17	.10582	.12385	.15402	.18354	.22163	.29428
18	.12317	.14261	.17473	.20575	.24527	.31947
19	.14061	.16129	.19507	.22731	.26797	.34324
20	.15799	.17974	.21492	.24817	.28969	.36563
22	.19215	.21560	.25292	.28761	.33025	.40663
24	.22503	.24971	.28847	.32400	.36713	.44311
26	.25634	.28186	.32151	.35746	.40063	.47566
28	.28593	.31200	.35214	.38818	.43110	.50482
30	.31379	.34018	.38049	.41641	.45885	.53106
34	.36449	.39103	.43108	.46628	.50740	.57625
38	.40909	.43536	.47461	.50878	.54831	.61370
42	.44838	.47413	.51231	.54529	.58315	.64520
46	.48312	.50821	.54519	.57692	.61314	.67203
50	.51397	.53834	.57407	.60456	.63919	.69513
60	.57761	.60010	.63275	.66033	.69137	.74089
70	.62690	.64760	.67745	.70250	.73049	.77478
80	.66607	.68517	.71257	.73544	.76088	.80086
90	.69790	.71558	.74086	.76186	.78514	.82155
100	.72425	.74069	.76411	.78351	.80495	.83836
120	.76529	.77966	.80005	.81686	.83535	.86399
140	.79575	.80850	.82652	.84133	.85757	.88262
160	.81924	.83068	.84682	.86004	.87451	.89676
180	.83790	.84827	.86287	.87481	.88786	.90787
200	.85307	.86255	.87588	.88677	.89864	.91682
250	.88095	.8875	.89969	.90860	.91828	.93307
300	.89995	.90656	.91584	.92337	.93155	.94401

Table 4 (Continued)

p = 6

N	α					
	.005	.01	.025	.05	.1	.25
7	.0 ⁷ 70557	.0 ⁶ 29697	.0 ⁵ 18030	.0 ⁵ 74790	.0 ⁴ 31547	.0 ³ 24844
8	.0 ⁴ 34541	.0 ⁴ 71870	.0 ³ 19456	.0 ³ 42669	.0 ³ 97879	.0 ² 33335
9	.0 ³ 40126	.0 ³ 67578	.0 ² 13837	.0 ² 25527	.0 ² 45255	.011336
10	.0 ² 16522	.0 ² 24979	.0 ² 44243	.0 ² 70038	.011482	.024216
11	.0 ² 42686	.0 ² 60326	.0 ² 97479	.014353	.021791	.041033
12	.0 ² 85127	.011478	.017390	.024325	.034966	.060679
13	.014444	.018800	.027141	.036529	.050383	.082172
14	.021960	.027821	.038682	.050510	.067439	.10473
15	.030903	.038302	.051661	.065830	.085610	.12776
16	.041061	.049980	.065743	.082100	.104468	.15085
17	.052219	.062606	.080634	.098998	.12368	.17368
18	.064174	.075950	.096078	.11626	.14298	.19605
19	.076743	.089816	.11187	.13367	.16217	.21782
20	.089763	.10403	.12783	.15107	.18112	.23890
22	.11662	.13299	.15975	.18538	.21788	.27885
24	.14388	.16196	.19107	.21850	.25277	.31577
26	.17096	.19041	.22132	.25008	.28556	.34973
28	.19745	.21798	.25027	.27997	.31624	.38093
30	.22313	.24449	.27778	.30812	.34485	.40960
34	.27153	.29397	.32844	.35939	.39633	.46025
38	.31571	.33866	.37355	.40450	.44105	.50339
42	.35576	.37885	.41364	.44424	.48005	.54044
46	.39199	.41497	.44935	.47936	.51425	.57254
50	.42477	.44748	.48125	.51055	.54441	.60056
60	.49406	.51571	.54754	.57485	.60606	.65708
70	.54917	.56955	.59930	.62460	.65331	.69978
80	.59381	.61292	.64067	.66413	.69059	.73311
90	.63060	.64852	.67442	.69622	.72071	.75983
100	.66139	.67822	.70246	.72278	.74553	.78171
120	.70994	.72488	.74628	.76412	.78400	.81540
140	.74642	.75981	.77891	.79479	.81241	.84011
160	.77480	.78691	.80415	.81843	.83423	.85899
180	.79750	.80854	.82423	.83720	.85152	.87389
200	.81605	.82620	.84058	.85246	.86555	.88595
250	.85037	.85879	.87069	.88048	.89124	.90796
300	.87391	.88110	.89124	.89957	.90870	.92285

Table 4 (Continued)

p = 7

N	α					
	.005	.01	.025	.05	.1	.25
8	.0 ⁷ 21289	.0 ⁷ 86044	.0 ⁶ 55120	.0 ⁵ 22835	.0 ⁵ 95942	.0 ⁴ 72580
9	.0 ⁴ 115809	.0 ⁴ 24239	.0 ⁴ 66388	.0 ³ 14730	.0 ³ 34311	.0 ² 12149
10	.0 ³ 14825	.0 ³ 25197	.0 ³ 52402	.0 ³ 94336	.0 ² 17761	.0 ² 46290
11	.0 ³ 66517	.0 ² 10165	.0 ² 18324	.0 ² 29501	.0 ² 49404	.010845
12	.0 ² 18510	.0 ² 26462	.0 ² 43552	.0 ² 65237	.010119	.019815
13	.0 ² 39356	.0 ² 53692	.0 ² 82864	.011790	.017307	.031195
14	.0 ² 70567	.0 ² 92955	.013667	.018704	.026327	.044531
15	.011263	.014435	.020431	.027115	.036919	.059370
16	.016537	.020729	.028448	.036821	.048798	.075298
17	.022812	.028074	.037553	.047610	.061690	.091967
18	.029994	.036345	.047578	.059270	.075347	.10909
19	.037977	.045412	.058355	.071609	.089554	.12644
20	.046648	.055143	.069730	.084457	.10413	.14384
22	.065631	.076124	.093740	.11111	.13380	.17825
24	.086164	.098448	.11870	.13831	.16346	.21158
26	.10761	.12146	.14396	.16541	.19254	.24343
28	.12949	.14467	.16905	.19200	.22067	.27360
30	.15142	.16774	.19367	.21781	.24767	.30205
34	.19449	.21253	.24073	.26653	.29794	.35389
38	.23555	.25472	.28433	.31106	.34320	.39952
42	.27402	.29390	.32429	.35146	.38380	.43972
46	.30974	.33002	.36076	.38801	.42019	.47524
50	.34277	.36321	.39399	.42108	.45287	.50678
60	.41458	.43479	.46486	.49099	.52126	.57177
70	.47344	.49296	.52177	.54656	.57505	.62203
80	.52215	.54082	.56817	.59156	.61826	.66192
90	.56296	.58071	.60661	.62863	.65366	.69431
100	.59755	.61441	.63890	.65965	.68314	.72109
120	.65283	.66805	.69003	.70854	.72937	.76277
140	.69496	.70876	.72861	.74526	.76391	.79366
160	.72806	.74066	.75872	.77381	.79067	.81747
180	.75474	.76630	.78284	.79664	.81201	.83636
200	.77668	.78735	.80260	.81529	.82941	.85172
250	.81754	.82649	.83923	.84979	.86149	.87991
300	.84580	.85349	.86441	.87344	.88343	.89910

Table 4 (Continued)

p = 8

N	α					
	.005	.01	.025	.05	.1	.25
9	.0 ⁸ 71788	.0 ⁷ 27598	.0 ⁶ 17155	.0 ⁶ 72189	.0 ⁵ 31412	.0 ⁴ 20202
10	.0 ⁵ 39473	.0 ⁵ 83064	.0 ⁴ 22961	.0 ⁴ 51489	.0 ³ 12180	.0 ³ 44134
11	.0 ⁴ 55082	.0 ⁴ 94377	.0 ³ 19900	.0 ³ 36314	.0 ³ 69598	.0 ² 18771
12	.0 ³ 26703	.0 ³ 41204	.0 ³ 75447	.0 ² 12329	.0 ² 21036	.0 ² 47822
13	.0 ³ 79550	.0 ² 11491	.0 ² 19226	.0 ² 29243	.0 ² 46224	.0 ² 93706
14	.0 ² 17954	.0 ² 24756	.0 ² 38847	.0 ² 56126	.0 ² 83944	.015649
15	.0 ² 33920	.0 ² 45162	.0 ² 67510	.0 ² 93791	.013445	.023497
16	.0 ² 56696	.0 ² 73433	.010564	.014227	.019719	.032724
17	.0 ² 86711	.010982	.015313	.020106	.027108	.043115
18	.012404	.015421	.020950	.026931	.035479	.054455
19	.016850	.020620	.027401	.034597	.044691	.066544
20	.021969	.026523	.034584	.042993	.054605	.079202
22	.034018	.040171	.050778	.061544	.076025	.10563
24	.048080	.055801	.068835	.081781	.098838	.13274
26	.063675	.072873	.088141	.10304	.12235	.15986
28	.080369	.090921	.10820	.12482	.14605	.18654
30	.097792	.10956	.12861	.14671	.16957	.21248
34	.13368	.14749	.16941	.18984	.21517	.26153
38	.16963	.18497	.20900	.23105	.25801	.30642
42	.20463	.22109	.24659	.26972	.29769	.34717
46	.23812	.25537	.28186	.30567	.33419	.38404
50	.26985	.28764	.31474	.33892	.36766	.41741
60	.34107	.35942	.38699	.41123	.43966	.48794
70	.40149	.41973	.44689	.47052	.49797	.54398
80	.45272	.47053	.49686	.51959	.54581	.58935
90	.49642	.51364	.53895	.56068	.58561	.62671
100	.53399	.55054	.57479	.59551	.61917	.65795
120	.59497	.61020	.63235	.65116	.67250	.70718
140	.64217	.65616	.67642	.69354	.71289	.74412
160	.67967	.79256	.71117	.72684	.74448	.77285
180	.71015	.72207	.73925	.75367	.76986	.79580
200	.73538	.74646	.76238	.77572	.79067	.81456
250	.78277	.79216	.80559	.81680	.82932	.84922
300	.81582	.82394	.83554	.84520	.85595	.87299

Table 4 (Continued)

p = 9

N	α					
	.005	.01	.025	.05	.1	.25
10	.0 ⁸ 25350	.0 ⁸ 92163	.0 ⁷ 56095	.0 ⁶ 23259	.0 ⁵ 10164	.0 ⁵ 61014
11	.0 ⁵ 13612	.0 ⁵ 28789	.0 ⁵ 80284	.0 ⁴ 18169	.0 ⁴ 42570	.0 ³ 16718
12	.0 ⁴ 20532	.0 ⁴ 35438	.0 ⁴ 75647	.0 ³ 13971	.0 ³ 27192	.0 ³ 75487
13	.0 ³ 10683	.0 ³ 16629	.0 ³ 30880	.0 ³ 51137	.0 ³ 88727	.0 ² 20826
14	.0 ³ 33897	.0 ³ 49428	.0 ³ 83940	.0 ² 12945	.0 ² 20813	.0 ² 43517
15	.0 ³ 80903	.0 ² 11264	.0 ² 17946	.0 ² 26291	.0 ² 39930	.0 ² 76857
16	.0 ² 16061	.0 ² 21595	.0 ² 32774	.0 ² 46163	.0 ² 67287	.012112
17	.0 ² 28055	.0 ² 36693	.0 ² 53582	.0 ² 73143	.010304	.017596
18	.0 ² 44630	.0 ² 57070	.0 ² 80753	.010744	.014715	.024063
19	.0 ² 66139	.0 ² 8300	.011439	.014894	.019925	.031413
20	.0 ² 92748	.011455	.015437	.019734	.025874	.039535
22	.016116	.019398	.025210	.031285	.039702	.057658
24	.024859	.029330	.037060	.044940	.055599	.077608
26	.035265	.040948	.050587	.060218	.072995	.098694
28	.047052	.053924	.065397	.076672	.091395	.12038
30	.059939	.067945	.081137	.093923	.11040	.14225
34	.088001	.098046	.11426	.12963	.14902	.18545
38	.11775	.12949	.14914	.16553	.18713	.22683
42	.14799	.16109	.18165	.20057	.22377	.26572
46	.17794	.19211	.21411	.23416	.25850	.30189
50	.20710	.22208	.24517	.26601	.29111	.33536
60	.27498	.29117	.31575	.33759	.36349	.40820
70	.33480	.35143	.37641	.39836	.42410	.46791
80	.38693	.40335	.42832	.44991	.47503	.51732
90	.43229	.44866	.47289	.49388	.51815	.55869
100	.47190	.48787	.51140	.53167	.55501	.59374
120	.53733	.55232	.57426	.59302	.61448	.64973
140	.58885	.60282	.62318	.64050	.66020	.69235
160	.63029	.64331	.66220	.67821	.69636	.72583
180	.66429	.67643	.69401	.70886	.72564	.75279
200	.69264	.70400	.72040	.73422	.74981	.77495
250	.74638	.75611	.77011	.78186	.79505	.81621
300	.78422	.79271	.80488	.81507	.82648	.84472

Table 4 (Continued)

p = 10

α	.005	.010	.025	.05	.1	.25
N						
11	.0 ⁸ 11246	.0 ⁸ 35733	.0 ⁷ 19324	.0 ⁷ 77218	.0 ⁶ 33526	.0 ⁵ 21386
12	.0 ⁶ 47154	.0 ⁵ 10036	.0 ⁵ 28256	.0 ⁵ 64552	.0 ⁴ 15891	.0 ⁴ 60952
13	.0 ⁵ 76669	.0 ⁴ 13324	.0 ⁴ 28760	.0 ⁴ 53699	.0 ³ 10610	.0 ³ 30376
14	.0 ⁴ 42583	.0 ⁴ 66814	.0 ³ 12568	.0 ³ 21066	.0 ³ 37102	.0 ³ 89399
15	.0 ³ 14331	.0 ³ 21078	.0 ³ 36286	.0 ³ 56666	.0 ³ 92523	.0 ² 19892
16	.0 ³ 36052	.0 ³ 50647	.0 ³ 81824	.0 ² 12140	.0 ² 18755	.0 ² 37055
17	.0 ³ 75024	.0 ² 10179	.0 ² 15665	.0 ² 22346	.0 ² 33072	.0 ² 61162
18	.0 ² 13670	.0 ² 18040	.0 ² 26712	.0 ² 36920	.0 ² 52683	.0 ² 92661
19	.0 ² 22587	.0 ² 29142	.0 ² 41802	.0 ² 56300	.0 ² 78244	.013126
20	.0 ² 34640	.0 ² 43855	.0 ² 61253	.0 ² 80714	.010952	.017701
22	.0 ² 69174	.0 ² 84980	.011376	.014478	.018907	.028781
24	.011834	.014208	.018414	.022817	.028932	.042056
26	.018193	.021449	.027092	.032865	.040709	.057042
28	.025886	.030071	.037194	.044348	.053894	.073283
30	.034761	.039886	.048484	.056983	.068156	.090387
34	.055364	.062312	.073700	.084683	.098783	.12595
38	.078662	.087261	.10111	.11422	.13076	.16184
42	.10353	.11356	.12949	.14436	.16285	.19695
46	.12913	.14036	.15800	.17428	.19429	.23066
50	.15483	.16705	.18607	.20346	.22464	.26265
60	.21714	.23106	.25238	.27151	.29443	.33462
70	.27444	.28925	.31169	.33157	.35512	.39574
80	.32589	.34109	.36390	.38395	.40748	.44759
90	.37168	.38694	.40970	.42955	.45270	.49182
100	.41235	.42747	.44991	.46937	.49195	.52983
120	.48083	.49536	.51674	.53515	.55634	.59151
140	.53578	.54954	.56969	.58695	.60669	.63923
160	.58059	.59356	.61249	.62861	.64700	.67713
180	.61773	.62994	.64770	.66279	.67994	.70791
200	.64895	.66046	.67715	.69129	.70732	.73339
250	.70872	.71871	.73313	.74529	.75903	.78122
300	.75125	.76003	.77268	.78332	.79529	.81455

Table 5. Correction factors for Mauchly's sphericity test - $2 \rho \log W$.

N	p = 3			p = 6			p = 10			
	5%		1%	5%		1%	5%		1%	
	Exact	Davis	Exact	Davis	Exact	Davis	Exact	Davis	Exact	Davis
5	1.07404	1.074	1.09101	1.091						
6	1.03761	1.0376	1.04629	1.0463						
7	1.02279	1.0228	1.02804	1.0280	0.80918	-	0.86151	-		
8	1.01529	1.0153	1.01881	1.0188	0.74526	-	0.76618	-		
9	1.01097	1.0110	1.01349	1.0135	1.10560	1.105	1.12270	1.12		
11	1.00644	1.0064	1.00790	1.0079	1.05084	1.0508	1.05810	1.058	1.46774	-
13	1.00423	1.0042	1.00519	1.0052	1.03026	1.0305	1.03434	1.0343	1.15373	1.15
16	1.00257	1.0026	1.00315	1.0031	1.01693	1.0169	1.01910	1.0191	1.06697	1.067
21	1.00137	1.0014	1.00167	1.0017	1.00844	1.0084	1.00949	1.0095	1.02875	1.0287
									1.03112	1.0311

CHAPTER IV

DISTRIBUTION OF THE LIKELIHOOD
RATIO CRITERION FOR TESTING $\Sigma = \Sigma_0$ 1. INTRODUCTION

Let $p \times 1$ vectors x_1, x_2, \dots, x_N be a random sample from a p -variate normal distribution with unknown mean vector μ and positive definite covariance matrix Σ . The likelihood ratio criterion for testing the hypothesis $H_0: \Sigma = \Sigma_0$ against the alternatives $H_1: \Sigma \neq \Sigma_0$, for some given positive definite matrix Σ_0 , is given by (Anderson [1]),

$$(1.1) \quad \lambda = (e/N)^{\frac{Np}{2}} \frac{|\Sigma \Sigma_0^{-1}|^{N/2}}{|\Sigma_0|^{N/2}} e^{-\frac{1}{2} \text{tr} \Sigma_0^{-1} S}$$

where $S = \sum_{i=1}^N (x_i - \bar{x})(x_i - \bar{x})'$ and $\bar{x} = \sum_{i=1}^N x_i / N$.

This likelihood ratio is not unbiased. However, if the criterion is modified by reducing the sample size N to the degrees of freedom $n = N-1$, then Sugiura and Nagao [27] have shown that the test is unbiased. The monotonicity of the power function with respect to each of the p characteristic roots of $\Sigma \Sigma_0^{-1}$ is established by Nagao [22] and Das Gupta [7].

Let

$$(1.2) \quad \lambda_1 = (e/n)^{\frac{np}{2}} \left| \frac{S}{\Sigma} \Sigma_0^{-1} \frac{n}{2} e^{-\frac{1}{2} \text{tr} \Sigma_0^{-1} S} \right|$$

be the modified likelihood ratio statistic. Korin [17] expressed the null distribution of $-2 \log \lambda_1$ in the form of an asymptotic series of central chi-square distributions and computed its percentage points but his tables are incomplete in regard to small values of n for the values of $p = 3(1)10$. Recently Davis [8] expressed the percentage points of $-2 \log \lambda_1$ in terms of chisquared percentiles using a Cornish-Fisher inversion of Box's series but his tables are also incomplete in regard to small values of n for the values of p he has considered i.e. $p = 6$ and 10 .

The object of the present chapter is to develop a method similar to the one used in Chapter III, in order to obtain the exact distribution of $L = \lambda_1^{2/n}$ in a series form and to compute percentage points of L to any degree of accuracy even for small sample sizes. Tables of percentage points for $p = 2(1)10$ for various significance levels are given and comparisons made with the results of Korin [17] and Davis [8].

2. DERIVATION OF THE DISTRIBUTION OF $L = \lambda_1^{2/n}$

The h -th moment of λ_1 under the null hypothesis is given by

$$(2.1) \quad E(\lambda_1^h) = (2e/n)^{\frac{nhp}{2}} \left[\frac{\Gamma_p\{n(1+h)/2\}}{\Gamma_p(n/2)} \right] \cdot (1+h)^{-np(1+h)/2},$$

where $\Gamma_p(x) = \pi^{p(p-1)/4} \prod_{i=1}^p \Gamma\{x - (i-1)/2\}$.

Let

$$L = \lambda_1^{2/n}.$$

Then

$$\begin{aligned} (2.2) \quad E(L^h) &= E\left(\frac{2h/n}{1}\right) \\ &= \frac{(2e/n)^{ph}}{\prod_{\alpha=1}^p \Gamma\left(\frac{n+1-\alpha}{2}\right)} \frac{\prod_{\alpha=1}^p \Gamma\left\{\frac{n}{2} + h + \frac{1-\alpha}{2}\right\}}{(1+2h/n)^{np(1+2h/n)/2}}. \end{aligned}$$

Using inverse Mellins' transform, the density of L is given by

$$(2.3) \quad f(L) = \left[\prod_{\alpha=1}^p \Gamma\left(\frac{n+1-\alpha}{2}\right) \right]^{-1} \cdot \frac{1}{2\pi i} \int_{-i\infty}^{i\infty} \frac{L^{-h-1} (2e/n)^{ph} \prod_{\alpha=1}^p \Gamma\left\{\frac{n}{2} + h + \frac{1-\alpha}{2}\right\}}{(1+2h/n)^{np(1+2h/n)/2}} dh.$$

Putting $\frac{n}{2} + h = t$ in (2.3), we have

$$(2.4) \quad f(L) = K(p,n) \cdot L^{\frac{n}{2} - 1} \cdot \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} L^{-t} C(t) dt$$

where $c = \frac{n}{2}$,

$$(2.5) \quad C(t) = (e/t)^{pt} \prod_{\alpha=1}^p \Gamma\left\{t + \frac{1-\alpha}{2}\right\},$$

and

$$(2.6) \quad K(p,n) = (2e/n)^{-pn/2} \left[\prod_{\alpha=1}^p \Gamma\left(\frac{n+1-\alpha}{2}\right) \right]^{-1}.$$

Using the expansion (2.64) of Chapter III to each gamma function in

(2.5), we have

$$(2.7) \quad \log C(t) = \frac{p}{2} \log(2\pi) - \frac{p(p+1)}{4} \log t + [A_1/t + A_2/t^2 + \dots \\ + A_\gamma/t^\gamma + \dots]$$

where the coefficients A_γ 's are given by

$$(2.8) \quad A_\gamma = (-1)^{\gamma-1} \left[\sum_{\alpha=1}^p B_{\gamma+1} \left(\frac{1-\alpha}{2} \right) \right] / \gamma(\gamma+1)$$

where $B_\gamma(x)$ is the Bernoulli polynomial of degree γ and order unity.

Thus

$$(2.9) \quad C(t) = (2\pi)^{p/2} t^{-p(p+1)/4} [1 + B_1/t + B_2/t^2 + \dots]$$

where the coefficients B_i 's can be obtained from (2.67) of Chapter III.

Using (2.9) in (2.4), we have the density of L as

$$(2.10) \quad f(L) = K(p,n) \cdot L^{\frac{n}{2} - 1} \cdot (2\pi)^{p/2} \cdot \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} L^{-t} t^{-v} \left[1 + \sum_{\gamma=1}^{\infty} B_\gamma/t^\gamma \right] dt$$

where $K(p,n)$ is as given in (2.6) and $v = p(p+1)/4$.

The integral on the right hand side of (2.10) can be easily computed if v is an integer and its value is by Cauchy's theorem of residues, the residue of $L^{-t} t^{-v} \left[1 + \sum_{\gamma=1}^{\infty} B_\gamma/t^\gamma \right]$ at $t = 0$. This is easily seen to be equal to

$$(2.11) \quad \sum_{\gamma=0}^{\infty} [(-\log L)^{v+\gamma-1} B_\gamma / \Gamma(v+\gamma)], \quad B_0 = 1$$

and thus from (2.10), the density of L is

$$(2.12) \quad f(L) = K(p,n) (2\pi)^{p/2} L^{\frac{n}{2} - 1} \sum_{\gamma=0}^{\infty} (B_\gamma) (-\log L)^{v+\gamma-1} / \Gamma(v+\gamma).$$

The probability that L is less than any value, say x_0 , is

$$(2.13) \quad P(L \leq x_0) = K(p, n) (2\pi)^{p/2} \sum_{\gamma=0}^{\infty} (B_{\gamma}) \int_0^{x_0} L^{\frac{n}{2}-1} (-\log L)^{v+\gamma-1} dL / \Gamma(v+\gamma) .$$

For computational purposes, we let

$$(2.14) \quad I_{v+\gamma-1, u} = \int_0^{x_0} L^u (-\log L)^{v+\gamma-1} dL / \Gamma(v+\gamma)$$

where $u = \frac{n}{2} - 1$. Then integrating by parts the right hand side of (2.14), we have the following recurrence relation:

$$(2.15) \quad I_{v+\gamma-1, u}(x_0) = [x_0^{u+1} (-\log x_0)^{v+\gamma-1} / \Gamma(v+\gamma) + I_{v+\gamma-2, u}] / (u+1)$$

and

$$(2.16) \quad I_{0, u}(x_0) = x_0^{(u+1)} / (u+1) .$$

With this notation (2.14) can be written as

$$(2.17) \quad P(L \leq x_0) = K(p, n) (2\pi)^{p/2} \sum_{\gamma=0}^{\infty} (B_{\gamma}) I_{v+\gamma-1, u}(x_0)$$

where $I_{v+\gamma-1, u}(x_0)$ satisfies the recurrence relations (2.15)

and (2.16). It is to be noted that (2.17) holds only if

$v = p(p+1)/4$ is an integer. Otherwise we can appeal to Theorem 2.1 of Chapter III, since in this case

$$(2.18) \quad \phi(t) = t^{-v} [1 + \sum_{\gamma=1}^{\infty} B_{\gamma} / t^{\gamma}] = O(t^{-v}) .$$

Thus according to the theorem, we can expand $\phi(t)$ in the factorial series as

$$\begin{aligned}
 (2.19) \quad \phi(t) &= t^{-v} \left[1 + \sum_{\gamma=1}^{\infty} B_{\gamma} / t^{\gamma} \right] \\
 &= \sum_{i=0}^{\infty} R_i \Gamma(t+\lambda) / \Gamma(t+v+i+\lambda)
 \end{aligned}$$

where λ is an arbitrary positive constant and can be chosen to govern the rate of convergence of the resulting series. The coefficients R_i 's can be determined explicitly as is done below.

Using the formula (2.6) of Chapter III to each gamma function on the right hand side of (2.19), we have

$$(2.20) \quad \log \{ \Gamma(t+\lambda) / \Gamma(t+v+i+\lambda) \} = [-(v+i) \log t + \sum_{j=1}^{\infty} A_{ij} / t_j]$$

where the first few coefficients A_{ij} 's which are needed in our computations are listed below.

$$\begin{aligned}
 A_{i1} &= -[t_i^2 - \lambda^2 - v] / 2 \\
 A_{i2} &= [t_i^3 - \lambda^3 - 3(t_i^2 - \lambda^2) / 2 + (t_i - \lambda) / 2] / 6 \\
 A_{i3} &= -[t_i^4 - \lambda^4 - 2(t_i^3 - \lambda^3) + (t_i^2 - \lambda^2)] / 12 \\
 A_{i4} &= [t_i^5 - \lambda^5 - 5(t_i^4 - \lambda^4) / 2 + 5(t_i^3 - \lambda^3) / 3 - (t_i - \lambda) / 6] / 20 \\
 (2.21) \quad A_{i5} &= -[t_i^6 - \lambda^6 - 3(t_i^5 - \lambda^5) / 2 + 5(t_i^4 - \lambda^4) / 2 - (t_i^2 - \lambda^2) / 2] / 30 \\
 A_{i6} &= [t_i^7 - \lambda^7 - 7(t_i^6 - \lambda^6) / 2 + 7(t_i^5 - \lambda^5) / 2 \\
 &\quad - 7(t_i^3 - \lambda^3) / 6 + (t_i - \lambda) / 6] / 42 \\
 A_{i7} &= -[t_i^8 - \lambda^8 - 4(t_i^7 - \lambda^7) + 14(t_i^6 - \lambda^6) / 3 \\
 &\quad - 7(t_i^4 - \lambda^4) / 3 + 2(t_i^2 - \lambda^2) / 3] / 56
 \end{aligned}$$

$$A_{i8} = [t_i^{9-\lambda^9} - 9(t_i^{8-\lambda^8})/2 + 6(t_i^{7-\lambda^7}) - 21(t_i^{5-\lambda^5})/5 \\ + 2(t_i^{3-\lambda^3}) - 3(t_i^{-\lambda})/10]/72$$

$$A_{i9} = -[t_i^{10-\lambda^{10}} - 5(t_i^{9-\lambda^9}) + 15(t_i^{8-\lambda^8})/2 \\ - 7(t_i^{6-\lambda^6}) + 5(t_i^{4-\lambda^4}) - 3(t_i^{2-\lambda^2})/2]/90$$

$$A_{i10} = [t_i^{11-\lambda^{11}} - 11(t_i^{10-\lambda^{10}})/2 + 55(t_i^{9-\lambda^9})/6 \\ - 11(t_i^{7-\lambda^7}) + 11(t_i^{5-\lambda^5}) - 11(t_i^{3-\lambda^3})/2]/110$$

where $t_i = \lambda + v + i$. The remaining coefficients are listed in Appendix B. Thus from (2.20) we have

$$(2.22) \quad \Gamma(t+\lambda)/\Gamma(t+\lambda+v+i) = t^{-(v+i)} \left[1 + \sum_{j=1}^{\infty} (Q_{ij}/t^j) \right],$$

where the coefficients Q_{ij} 's can be obtained in terms of A_{ij} as in Chapter III. Using (2.22) on the right hand side of (2.19) we have

$$(2.23) \quad t^{-v} \left[1 + \sum_{\gamma=1}^{\infty} (B_{\gamma}/t^{\gamma}) \right] = \sum_{i=0}^{\infty} R_i t^{-(v+i)} \left[1 + \sum_{j=1}^{\infty} (Q_{ij}/t^j) \right].$$

Equating the coefficients of t on both sides of (2.23), it can be seen that R_i 's can be determined using (2.90) of Chapter III. Now using (2.19) in (2.10) and noting that the term by term integration is valid since a factorial series is uniformly convergent in a half-plane (see Doetch [9]) we have the density of L in the case that v is not an integer in the form

$$\begin{aligned}
 (2.24) \quad f(L) &= K(p,n) L^{\frac{n}{2} - 1} (2\pi)^{p/2} \sum_{i=0}^{\infty} R_i \frac{1}{2^{\pi i}} \\
 &\quad \cdot \int_{c-i\infty}^{c+i\infty} L^{-t} [\Gamma(t+\lambda)/\Gamma(t+\lambda+v+i)] dt \\
 &= K(p,n) (2\pi)^{p/2} \sum_{i=0}^{\infty} R_i L^{\frac{n}{2} + \lambda - 1} (1-L)^{v+i-1} / \Gamma(v+i) .
 \end{aligned}$$

The distribution of L is then given by

$$(2.25) \quad P(L \leq x_0) = K(p,n) (2\pi)^{p/2} \sum_{i=0}^{\infty} R_i I_{x_0} \left(\frac{n}{2} + \lambda, v+i \right) / \Gamma(v+i)$$

where $I_{x_0}(p,q)$ is the incomplete beta function $\int_0^{x_0} x^{p-1} (1-x)^{q-1} dx$.

3. COMPUTATIONS OF PERCENTAGE POINTS

.005, .01, .025, .05, .1 and .25 significance points of $L = \lambda_1^{2/n}$ were computed for $p = 2(1)10$ and various n using (2.17) and (2.25) and these are presented in table 6. The computation was carried out on CDC 6500 using double precision arithmetic. Table 7 compares the exact values with those obtained by Korin [17] and Davis [8].

Table 6. Percentage Points of $L = \lambda_1^2/n$ for Testing $\Sigma = \Sigma_0$

p = 2

$n \backslash \alpha$.005	.01	.025	.05	.1	.25
2	.0 ⁴ 11569	.0 ⁴ 46327	.0 ³ 29074	.0 ² 11716	.0 ² 047596	.031333
3	.0 ² 27788	.0 ² 55954	.014187	.028861	.059260	.15763
4	.018257	.029250	.054811	.088639	.14444	.28133
5	.047798	.068263	.10982	.15809	.22903	.38018
6	.085961	.11453	.16797	.22533	.30386	.45748
7	.12773	.16242	.22387	.28636	.36795	.51863
8	.16994	.20897	.27544	.34047	.42251	.56786
9	.21086	.25281	.32217	.38807	.46909	.60820
10	.24963	.29344	.36422	.42995	.50913	.64180
11	.28591	.33080	.40199	.46689	.54380	.67017
12	.31962	.36503	.43594	.49961	.57407	.69443
13	.35085	.39636	.46653	.52874	.60067	.71541
14	.37974	.42505	.49417	.55479	.62422	.73371
15	.40645	.45136	.51923	.57821	.64521	.74982
16	.43118	.47554	.54203	.59935	.66401	.76410
17	.45410	.49779	.56285	.61852	.68094	.77685
18	.47536	.51834	.58191	.63599	.69627	.78829
19	.49513	.53734	.59943	.65195	.71022	.79863
20	.51355	.55496	.61557	.66659	.72295	.80800
22	.54678	.58658	.64433	.69252	.74535	.82437
24	.57592	.61411	.66914	.71474	.76442	.83817
26	.60165	.63828	.69077	.73400	.78085	.84996
28	.62450	.65966	.70978	.75084	.79515	.86016
30	.64492	.67869	.72660	.76569	.80770	.86906
32	.66328	.69573	.74160	.77888	.81881	.87690
34	.67985	.71107	.75505	.79067	.82871	.88385
36	.69490	.72495	.76718	.80127	.83759	.89006
38	.70861	.73758	.77817	.81086	.84559	.89564
40	.72115	.74910	.78817	.81956	.85285	.90068
45	.74828	.77395	.80965	.83819	.86832	.91138
50	.77062	.79434	.82719	.85334	.88085	.92001
55	.78935	.81137	.84177	.86590	.89121	.92710
60	.80526	.82580	.85409	.87648	.89991	.93304
65	.81894	.83819	.86464	.88552	.90732	.93808
70	.83083	.84894	.87376	.89332	.913712	.94242
75	.84127	.84835	.88174	.90013	.91928	.94619
80	.85049	.86666	.88876	.90613	.92417	.94950
85	.85870	.87405	.89500	.91144	.92851	.95242
90	.86606	.88067	.90058	.91619	.93237	.95502
95	.87269	.88662	.90560	.92045	.93584	.95736
100	.87869	.89201	.91013	.92430	.93897	.95946

Table 6 (continued)

p = 3

$n \backslash \alpha$.005	.01	.025	.05	.1	.25
3	.0 ⁵ ₃ 2506	.0 ⁴ ₃ 10160	.0 ⁴ ₂ 64805	.0 ³ ₂ 26498	.0 ² ₁ 11027	.0 ² ₁ 78035
4	.0 ³ ₂ 81873	.0 ³ ₁ 66678	.0 ² ₁ 43296	.0 ² ₁ 90548	.019385	.056755
5	.0 ² ₁ 65792	.010719	.020707	.034585	.058931	.12591
6	.019979	.029067	.048280	.071783	.10854	.19607
7	.040201	.054573	.082592	.11426	.16034	.26062
8	.065194	.084426	.11991	.15788	.21047	.31801
9	.093006	.11638	.15782	.20042	.25732	.36846
10	.12215	.14891	.19492	.24079	.30037	.41275
11	.15160	.18105	.23046	.27854	.33964	.45172
12	.18067	.21220	.26407	.31356	.37534	.48614
13	.20894	.24205	.29564	.34594	.40781	.51671
14	.23617	.27044	.32517	.37583	.43736	.54398
15	.26223	.29732	.35272	.40341	.46431	.56844
16	.28705	.32270	.37842	.42888	.48896	.59048
17	.31063	.34663	.40239	.45243	.51155	.61044
18	.33300	.36916	.42475	.47426	.53231	.62858
19	.35419	.39039	.44564	.49450	.55145	.64513
20	.37427	.41038	.46517	.51333	.56914	.66030
22	.41130	.44701	.50062	.54724	.60075	.68710
24	.44458	.47965	.53187	.57688	.62813	.71002
26	.47455	.50887	.55958	.60298	.65207	.72984
28	.50164	.53512	.58430	.62612	.67315	.74714
30	.52661	.55882	.60647	.64676	.69186	.76238
32	.54857	.58029	.62644	.66528	.70857	.77589
34	.56899	.59983	.64452	.68198	.72357	.78796
36	.58769	.61768	.66096	.69711	.73712	.79880
38	.60488	.63403	.67597	.71089	.74941	.80859
40	.62073	.64907	.68973	.72348	.76061	.81747
45	.65540	.68184	.71954	.75065	.78470	.83645
50	.68434	.70907	.74418	.77300	.80440	.85187
55	.70886	.73206	.76486	.79169	.82081	.86464
60	.72987	.75170	.78247	.80754	.83468	.87538
65	.74808	.76868	.79763	.82117	.84657	.88454
70	.76401	.78350	.81083	.83299	.85685	.89245
75	.77805	.79654	.82242	.84335	.86587	.89935
80	.79052	.80811	.83267	.85251	.87381	.90541
85	.80168	.81843	.84181	.86065	.88086	.91079
90	.81171	.82771	.85000	.86795	.88717	.91559
95	.82077	.83609	.85739	.87452	.89284	.91990
100	.82901	.84369	.86408	.88047	.89797	.92379

Table 6 (Continued)

p = 4

α						
n	.005	.01	.025	.05	.1	.25
4	.0 ⁶ 68258	.0 ⁵ 27562	.0 ⁴ 17377	.0 ⁴ 71482	.0 ³ 30192	.0 ² 22422
5	.0 ³ 26081	.0 ³ 53590	.0 ² 14157	.0 ² 30195	.0 ² 66607	.020829
6	.0 ² 24414	.0 ² 40310	.0 ² 79751	.013653	.024052	.054981
7	.0 ² 83707	.012366	.021080	.032154	.050263	.096845
8	.018557	.025595	.039760	.056400	.081716	.14106
9	.032548	.042820	.062385	.084139	.11560	.18472
10	.049524	.062930	.087449	.11362	.15008	.22637
11	.068633	.084918	.11378	.14363	.18402	.26541
12	.089139	.10798	.14054	.17338	.21678	.30166
13	.11045	.13149	.16715	.20236	.24803	.33516
14	.13210	.15503	.19322	.23030	.27762	.36606
15	.15377	.17829	.21852	.25702	.30550	.39454
16	.17521	.20104	.24291	.28248	.33172	.42082
17	.19625	.22315	.26631	.30664	.35634	.44510
18	.21677	.24455	.28869	.32955	.37945	.46756
19	.23670	.26517	.31004	.35123	.40114	.48838
20	.25599	.28500	.33040	.37175	.42151	.50772
22	.29256	.32229	.36823	.40954	.45867	.54249
24	.32645	.35652	.40252	.44344	.49164	.57283
26	.35777	.38791	.43362	.47392	.52101	.59951
28	.38669	.41670	.46188	.50143	.54731	.62312
30	.41339	.44314	.48764	.52634	.57097	.64416
32	.43808	.46746	.51117	.54898	.59235	.66301
34	.46092	.48987	.53273	.56962	.61175	.67998
36	.48210	.51057	.55255	.58851	.62943	.69536
38	.50178	.52973	.57081	.60586	.64560	.70933
40	.52009	.54751	.58768	.62184	.66044	.72210
45	.56068	.58677	.62470	.65673	.69269	.74962
50	.59514	.61992	.65574	.68582	.71940	.77222
55	.62471	.64824	.68211	.71042	.74188	.79109
60	.65034	.67270	.70478	.73147	.76105	.80708
65	.67274	.69403	.72446	.74970	.77758	.82081
70	.69249	.71277	.74170	.76562	.79198	.83272
75	.71001	.72937	.75692	.77965	.80464	.84315
80	.72566	.74418	.77046	.79211	.81585	.85236
85	.73972	.75746	.78258	.80323	.82585	.86055
90	.75243	.76943	.79349	.81324	.83482	.86788
95	.76396	.78029	.80336	.82227	.84292	.87447
100	.77446	.79017	.81234	.83048	.85026	.88045

Table 6 (Continued)

p = 5

$n \backslash \alpha$.005	.01	.025	.05	.1	.25
5	.0 ⁶ ₄ 209	.0 ⁶ ₃ 813	.0 ⁵ ₅ 571	.0 ⁴ ₂ 211	.0 ⁴ ₂ 883	.0 ³ ₂ 678
6	.0 ⁴ ₃ 8631	.0 ³ ₂ 1787	.0 ³ ₂ 4787	.0 ² ₁ 1038	.0 ² ₁ 2341	.0 ² ₁ 7707
7	.0 ³ ₂ 91558	.0 ² ₁ 5289	.0 ² ₁ 30854	.0 ² ₁ 53894	.0 ² ₁ 97536	.023522
8	.0 ² ₁ 34846	.0 ² ₁ 52164	.0 ² ₁ 90895	.014165	.022760	.046217
9	.0 ² ₁ 894223	.011779	.018714	.027119	.040364	.073262
10	.015873	.021176	.031545	.043440	.061249	.10262
11	.025655	.033050	.046932	.062212	.084226	.13284
12	.037426	.04693	.064206	.082619	.10836	.16301
13	.050797	.062331	.082772	.10400	.13296	.19252
14	.065391	.078826	.10214	.12584	.15754	.22102
15	.080870	.096049	.12193	.14777	.18176	.24833
16	.096950	.11370	.14185	.16952	.20539	.27435
17	.11340	.13155	.16167	.19088	.22830	.29907
18	.13001	.14942	.18124	.21175	.25040	.32249
19	.14666	.16715	.20043	.23201	.27165	.34465
20	.16321	.18465	.21916	.25163	.29203	.36562
22	.19568	.21865	.25508	.28883	.33022	.40422
24	.22696	.25104	.28877	.32329	.36513	.43880
26	.25679	.28165	.32020	.35511	.39702	.46987
28	.28506	.31043	.34945	.38447	.42617	.49787
30	.31174	.33742	.37663	.41156	.45285	.52321
32	.33687	.36270	.40190	.43654	.47733	.54621
34	.36052	.38637	.42539	.45971	.49983	.56717
36	.38276	.40854	.44726	.48115	.52057	.58634
38	.40369	.42932	.46764	.50104	.53974	.60392
40	.42337	.44880	.48667	.51954	.55748	.62011
45	.46775	.49249	.52906	.56052	.59655	.65542
50	.50618	.53010	.56523	.59525	.62943	.68481
55	.53967	.56272	.59640	.62503	.65745	.70964
60	.56907	.59125	.62350	.65080	.68158	.73087
65	.59505	.61636	.64727	.67331	.70258	.74924
70	.61815	.63864	.66825	.69313	.72101	.76527
75	.63881	.65851	.68691	.71071	.73731	.77939
80	.65738	.67635	.70361	.72641	.75182	.79191
85	.67417	.69244	.71864	.74050	.76482	.80309
90	.68942	.70702	.73223	.75322	.77654	.81314
95	.70331	.72030	.74458	.76476	.78715	.82222
100	.71603	.73243	.75585	.77528	.79680	.83045

Table 6 (Continued)

p = 6

$n \backslash \alpha$.005	.01	.025	.05	.1	.25
6	.0 ⁷ ₄ 679	.0 ⁶ ₄ 256	.0 ⁵ ₃ 156	.0 ⁵ ₃ 637	.0 ⁴ ₂ 272	.0 ³ ₂ 212
7	.0 ⁴ ₃ 2922	.0 ⁴ ₃ 6086	.0 ³ ₂ 1650	.0 ³ ₂ 3626	.0 ³ ₂ 8340	.0 ² ₂ 2857
8	.0 ³ ₂ 34475	.0 ³ ₂ 58147	.0 ² ₁ 1936	.0 ² ₁ 21210	.0 ² ₁ 39263	.0 ² ₁ 98987
9	.0 ² ₁ 4392	.0 ² ₁ 21796	.0 ² ₁ 38712	.0 ² ₁ 61448	.010108	.021457
10	.0 ² ₁ 37618	.0 ² ₁ 53261	.0 ² ₁ 86316	.012744	.019416	.036790
11	.0 ² ₁ 75778	.010237	.015555	.021817	.031466	.054938
12	.012966	.016912	.024487	.033043	.045723	.075008
13	.019867	.025215	.035157	.04026	.061643	.096263
14	.028141	.034941	.047256	.060366	.078740	.11814
15	.037607	.045856	.060477	.07570	.096606	.14020
16	.048072	.057730	.074544	.091731	.11491	.16215
17	.059348	.070352	.089215	.10819	.13341	.18377
18	.071262	.083532	.10429	.12488	.15189	.20491
19	.083659	.097107	.11959	.14163	.17021	.22546
20	.096406	.11094	.13499	.15831	.18826	.24538
22	.12250	.13894	.16566	.19111	.22232	.28315
24	.14884	.16682	.19567	.22274	.25644	.31816
26	.17492	.19416	.22465	.25292	.28772	.35048
28	.20043	.22065	.25239	.28153	.31705	.38029
30	.22515	.24614	.27881	.30854	.34448	.40778
32	.24896	.27053	.30387	.33397	.37011	.43316
34	.27181	.29381	.32759	.35790	.39405	.45661
36	.29366	.31597	.35062	.38039	.41643	.47833
38	.31453	.33703	.37121	.40154	.43736	.49848
40	.33443	.35704	.39124	.42144	.45695	.51721
45	.38017	.40277	.43664	.46626	.50078	.55866
50	.42070	.44301	.47621	.50503	.53838	.59375
55	.45668	.47854	.51089	.53880	.57090	.62381
60	.48873	.51006	.54146	.56842	.59927	.64980
65	.51741	.53816	.56857	.59457	.62421	.67249
70	.54318	.56332	.59275	.61781	.64628	.69245
75	.56643	.58597	.61443	.63859	.66595	.71015
80	.58751	.60645	.63396	.65726	.68358	.72594
85	.60668	.62504	.65165	.67413	.69946	.74011
90	.62420	.64199	.66774	.68943	.71384	.75290
95	.64025	.65750	.68242	.70338	.72692	.76450
100	.65501	.67175	.69588	.71615	.73886	.77507

Table 6 (Continued)

p = 7

$n \backslash \alpha$.005	.01	.025	.05	.1	.25
7	.07330	.06105	.06555	.05214	.05891	.04696
8	.04101	.04210	.04576	.03128	.03299	.02106
9	.031301	.032213	.034611	.03832	.021570	.024110
10	.025897	.029025	.021630	.022630	.024416	.029742
11	.0216566	.0223716	.0239120	.0258722	.0291321	.017970
12	.0235507	.0248512	.0275041	.010700	.015747	.028516
13	.0264117	.0284584	.012465	.017094	.02412	.040984
14	.010298	.013217	.018749	.024932	.034030	.054961
15	.015203	.019083	.026246	.034038	.045216	.07006
16	.021073	.025970	.034812	.04422	.057427	.08595
17	.027829	.033766	.044293	.055279	.070428	.10236
18	.035373	.042352	.054532	.067037	.084013	.11906
19	.043603	.051608	.065385	.079330	.098006	.13586
20	.05242	.061417	.076720	.092015	.11226	.15264
22	.071424	.082284	.10038	.11809	.14107	.18569
24	.091711	.10423	.12476	.14450	.16972	.21763
26	.11273	.12670	.14929	.17072	.19773	.24813
28	.13406	.14927	.17359	.19640	.22482	.27707
30	.15540	.17165	.19740	.22131	.25082	.30440
32	.17650	.19363	.22055	.24533	.27566	.33013
34	.19721	.21506	.24292	.26837	.29931	.35433
36	.21742	.23587	.26447	.29043	.32178	.37708
38	.23705	.25598	.28516	.31148	.34310	.39847
40	.25607	.27537	.30498	.33156	.36333	.41859
45	.30077	.32067	.35089	.37771	.40944	.46388
50	.34145	.36158	.39190	.41858	.44990	.50303
55	.37834	.39845	.42855	.45486	.48552	.53711
60	.41177	.43170	.46137	.48716	.51705	.56698
65	.44211	.46176	.49087	.51605	.54510	.59333
70	.46970	.48900	.51747	.54199	.57019	.61674
75	.48155	.51375	.54154	.56540	.59273	.63766
80	.49912	.53632	.56342	.58660	.61309	.65645
85	.51231	.55698	.58337	.60588	.63155	.67342
90	.52010	.57593	.60162	.62349	.64836	.68882
95	.54119	.59337	.61838	.63963	.66373	.70284
100	.55222	.60948	.63382	.65446	.67783	.71567

Table 6 (Continued)

p = 8

n	α					
	.005	.01	.025	.05	.1	.25
9	.0 ⁵ ₄ 294	.0 ⁵ ₄ 674	.0 ⁴ ₃ 196	.0 ⁴ ₃ 448	.0 ³ ₂ 107	.0 ³ ₂ 393
10	.0 ⁴ ₃ 490	.0 ⁴ ₃ 841	.0 ³ ₂ 177	.0 ³ ₂ 325	.0 ² ₁ 624	.0 ² ₁ 169
11	.0 ³ ₂ 2400	.0 ³ ₂ 3707	.0 ² ₁ 6799	.0 ² ₁ 1113	.0 ¹ ₀ 1903	.0 ¹ ₀ 4342
12	.0 ² ₁ 72041	.0 ² ₁ 10418	.0 ¹ ₀ 17462	.0 ¹ ₀ 26605	.0 ⁰ ₀ 42133	.0 ⁰ ₀ 85765
13	.0 ² ₁ 16368	.0 ² ₁ 22596	.0 ¹ ₀ 35523	.0 ¹ ₀ 51410	.0 ⁰ ₀ 77049	.014418
14	.0 ² ₁ 31108	.0 ² ₁ 41466	.0 ¹ ₀ 62100	.0 ¹ ₀ 86418	.012414	.021771
15	.0 ² ₁ 52269	.0 ² ₁ 67778	.0 ¹ ₀ 97682	.013177	.018299	.030474
16	.0 ² ₁ 80314	.010184	.014225	.018707	.025270	.040327
17	.011537	.014359	.019541	.025159	.033206	.051132
18	.015731	.019272	.025653	.032438	.041978	.062700
19	.020580	.024873	.032485	.040443	.051455	.074859
20	.026039	.031101	.039952	.049070	.061512	.087458
22	.038571	.045177	.056459	.067802	.082923	.11347
24	.052873	.060971	.074544	.087927	.10544	.13993
26	.068506	.077999	.093668	.10887	.12846	.16627
28	.085081	.095847	.11339	.13019	.15157	.19210
30	.10227	.11418	.13338	.15156	.17444	.21719
32	.11980	.13273	.15338	.17273	.19686	.24140
34	.13748	.15129	.17319	.19353	.21870	.26464
36	.15512	.16971	.19267	.21385	.23986	.28689
38	.17261	.18787	.21173	.23360	.26029	.30813
40	.18985	.20568	.23030	.25273	.27996	.32840
45	.23144	.24835	.27434	.29774	.32580	.37496
50	.27046	.28804	.31482	.33870	.36708	.41618
55	.30670	.32466	.35182	.37585	.40420	.45276
60	.34020	.35832	.38557	.40952	.43760	.48532
65	.37109	.38923	.416350	.44006	.46773	.51442
70	.39956	.41760	.44446	.46784	.49500	.54055
75	.42583	.44369	.47018	.49315	.51974	.56412
80	.45008	.46771	.49377	.51629	.54228	.58546
85	.47251	.48987	.51546	.53751	.56287	.60486
90	.49329	.51035	.53545	.55701	.58175	.62257
95	.51259	.52934	.55392	.57499	.59911	.63879
100	.53053	.54696	.57103	.59161	.61513	.65370

Table 6 (Continued)

p = 9

$n \backslash \alpha$.005	.01	.025	.05	.1	.25
10	.0 ₄ ⁵ 124	.0 ₄ ⁵ 260	.0 ₄ ⁵ 722	.0 ₃ ⁴ 163	.0 ₃ ⁴ 392	.0 ₃ ³ 147
11	.0 ₄ ⁴ 185	.0 ₃ ⁴ 320	.0 ₃ ⁴ 684	.0 ₃ ³ 126	.0 ₃ ³ 246	.0 ₂ ³ 687
12	.0 ₃ ⁴ 970	.0 ₃ ³ 151	.0 ₃ ³ 281	.0 ₂ ³ 466	.0 ₂ ³ 810	.0 ₂ ² 191
13	.0 ₃ ³ 3099	.0 ₂ ³ 4523	.0 ₂ ³ 7692	.0 ₂ ² 1188	.0 ₂ ² 1913	.0 ₂ ² 4012
14	.0 ₂ ³ 74384	.0 ₂ ² 10366	.0 ₂ ² 16541	.0 ₂ ² 24265	.0 ₂ ² 36974	.0 ₂ ² 71266
15	.0 ₂ ² 14841	.0 ₂ ² 19973	.0 ₂ ² 30359	.0 ₂ ² 42820	.0 ₂ ² 62518	.011287
16	.0 ₂ ² 26040	.0 ₂ ² 34091	.0 ₂ ² 49857	.0 ₂ ² 68150	.0 ₂ ² 61569	.016469
17	.0 ₂ ² 41592	.0 ₂ ² 53236	.0 ₂ ² 75439	.010050	.013787	.022608
18	.0 ₂ ² 61860	.0 ₂ ² 77705	.010724	.013982	.018734	.029614
19	.0 ₂ ² 87034	.010759	.014529	.018585	.024404	.037386
20	.011714	.014283	.018905	.023817	.030732	.045819
22	.019172	.022853	.029299	.035959	.045084	.064268
24	.028386	.033239	.041555	.049957	.061224	.084226
26	.039112	.045140	.055293	.065365	.078637	.10509
28	.051078	.058248	.070152	.081782	.096887	.12641
30	.064022	.072275	.085811	.098868	.11562	.14781
32	.077706	.086969	.10200	.11635	.13456	.16905
34	.091921	.10211	.11851	.13400	.15348	.18995
36	.10649	.11753	.13514	.15165	.17224	.21037
38	.12127	.13307	.15177	.16916	.19072	.23024
40	.13613	.14861	.16827	.18644	.20882	.24950
45	.17307	.18694	.20852	.22819	.25210	.29481
50	.20894	.22382	.24672	.26739	.29226	.33606
55	.24316	.25875	.28254	.30383	.32923	.37348
60	.27550	.29155	.31589	.33751	.36314	.407381
65	.30587	.32219	.34683	.36858	.39422	.43814
70	.33428	.35076	.37549	.39722	.42271	.46609
75	.36084	.37735	.40204	.42364	.44887	.49157
80	.38563	.40210	.42664	.44803	.47293	.51486
85	.40878	.42515	.44947	.47059	.49510	.53619
90	.43042	.44664	.47068	.49149	.51557	.55581
95	.45066	.46671	.49041	.51089	.53453	.57389
100	.46962	.48545	.50881	.52893	.55211	.59059

Table 6 (Continued)

p = 10

n	α					
	.005	.01	.025	.05	.1	.25
11	.0 ₅ ⁶ 463	.0 ₅ ⁶ 962	.0 ₄ ⁵ 262	.0 ₄ ⁵ 592	.0 ₄ ⁴ 143	.0 ₃ ⁴ 547
12	.0 ₄ ⁵ 6996	.0 ₄ ⁵ 1215	.0 ₃ ⁴ 2625	.0 ₃ ⁴ 4906	.0 ₃ ⁴ 9696	.0 ₂ ³ 2773
13	.0 ₃ ⁴ 3902	.0 ₃ ⁴ 6127	.0 ₃ ³ 1154	.0 ₃ ³ 1936	.0 ₃ ³ 3415	.0 ₂ ³ 8253
14	.0 ₃ ³ 13203	.0 ₃ ³ 19434	.0 ₂ ³ 33497	.0 ₂ ³ 52370	.0 ₂ ³ 85630	.0 ₂ ² 18458
15	.0 ₃ ³ 33375	.0 ₂ ³ 46925	.0 ₂ ² 75906	.0 ₂ ² 11275	.0 ₂ ² 17443	.0 ₂ ² 34545
16	.0 ₂ ³ 69756	.0 ₂ ² 94719	.0 ₂ ² 14596	.0 ₂ ² 20844	.0 ₂ ² 30892	.0 ₂ ² 57271
17	.0 ₂ ² 12760	.0 ₂ ² 16854	.0 ₂ ² 24986	.0 ₂ ² 34574	.0 ₂ ² 49504	.0 ₂ ² 87007
18	.0 ₂ ² 21158	.0 ₂ ² 27320	.0 ₂ ² 39237	.0 ₂ ² 52903	.0 ₂ ² 73624	.0 ₂ ² 12380
19	.0 ₂ ² 32552	.0 ₂ ² 41244	.0 ₂ ² 57677	.0 ₂ ² 76085	.0 ₂ ² 10338	.0 ₂ ² 16745
20	.0 ₂ ² 47226	.0 ₂ ² 58879	.0 ₂ ² 80485	.0 ₂ ² 10421	.0 ₂ ² 13870	.0 ₂ ² 21758
22	.0 ₂ ² 87070	.0 ₂ ² 10575	.0 ₂ ² 13931	.0 ₂ ² 17497	.0 ₂ ² 22523	.0 ₂ ² 33526
24	.0 ₂ ² 14116	.0 ₂ ² 16798	.0 ₂ ² 21501	.0 ₂ ² 26372	.0 ₂ ² 33067	.0 ₂ ² 47238
26	.0 ₂ ² 20893	.0 ₂ ² 24459	.0 ₂ ² 30588	.0 ₂ ² 36806	.0 ₂ ² 45184	.0 ₂ ² 62439
28	.0 ₂ ² 28915	.0 ₂ ² 33395	.0 ₂ ² 40974	.0 ₂ ² 48531	.0 ₂ ² 58548	.0 ₂ ² 78717
30	.0 ₂ ² 38025	.0 ₂ ² 43423	.0 ₂ ² 52430	.0 ₂ ² 61282	.0 ₂ ² 72856	.0 ₂ ² 95719
32	.0 ₂ ² 48057	.0 ₂ ² 54353	.0 ₂ ² 64736	.0 ₂ ² 7482	.0 ₂ ² 87841	.0 ₂ ² 11316
34	.0 ₂ ² 58846	.0 ₂ ² 66004	.0 ₂ ² 77692	.0 ₂ ² 88920	.0 ₂ ² 10328	.0 ₂ ² 13080
36	.0 ₂ ² 70239	.0 ₂ ² 78215	.0 ₂ ² 91124	.0 ₂ ² 10341	.0 ₂ ² 11898	.0 ₂ ² 14847
38	.0 ₂ ² 82098	.0 ₂ ² 90840	.0 ₂ ² 10488	.0 ₂ ² 11813	.0 ₂ ² 13480	.0 ₂ ² 16603
40	.0 ₂ ² 94302	.0 ₂ ² 10376	.0 ₂ ² 11883	.0 ₂ ² 13297	.0 ₂ ² 15061	.0 ₂ ² 18336
45	.0 ₂ ² 12567	.0 ₂ ² 13666	.0 ₂ ² 15394	.0 ₂ ² 16990	.0 ₂ ² 18955	.0 ₂ ² 22529
50	.0 ₂ ² 15732	.0 ₂ ² 16952	.0 ₂ ² 18850	.0 ₂ ² 20583	.0 ₂ ² 22691	.0 ₂ ² 26469
55	.0 ₂ ² 18845	.0 ₂ ² 20159	.0 ₂ ² 22185	.0 ₂ ² 24016	.0 ₂ ² 26224	.0 ₂ ² 30133
60	.0 ₂ ² 21860	.0 ₂ ² 23243	.0 ₂ ² 25361	.0 ₂ ² 27261	.0 ₂ ² 29536	.0 ₂ ² 33521
65	.0 ₂ ² 24748	.0 ₂ ² 26183	.0 ₂ ² 28365	.0 ₂ ² 30311	.0 ₂ ² 32625	.0 ₂ ² 36646
70	.0 ₂ ² 27497	.0 ₂ ² 28967	.0 ₂ ² 31193	.0 ₂ ² 33165	.0 ₂ ² 35500	.0 ₂ ² 39526
75	.0 ₂ ² 30102	.0 ₂ ² 31596	.0 ₂ ² 33847	.0 ₂ ² 35833	.0 ₂ ² 38173	.0 ₂ ² 42183
80	.0 ₂ ² 32565	.0 ₂ ² 34073	.0 ₂ ² 36336	.0 ₂ ² 38325	.0 ₂ ² 40658	.0 ₂ ² 44636
85	.0 ₂ ² 34891	.0 ₂ ² 36405	.0 ₂ ² 38669	.0 ₂ ² 40652	.0 ₂ ² 42970	.0 ₂ ² 46904
90	.0 ₂ ² 37084	.0 ₂ ² 38598	.0 ₂ ² 40856	.0 ₂ ² 42826	.0 ₂ ² 45124	.0 ₂ ² 49005
95	.0 ₂ ² 39153	.0 ₂ ² 40662	.0 ₂ ² 42907	.0 ₂ ² 44860	.0 ₂ ² 47132	.0 ₂ ² 50955
100	.0 ₂ ² 41105	.0 ₂ ² 42606	.0 ₂ ² 44833	.0 ₂ ² 46765	.0 ₂ ² 49007	.0 ₂ ² 52769

Table 7. Upper Percentage Points of $-2 \log \lambda_1$ for testing $\xi = \xi_0$
1%

n	p = 4		p = 6		p = 10	
	Korin	Series (2.17)	Korin	Davis	Korin	Davis
4	-	51.2067	-	-	-	-
5	-	37.65785	-	-	-	-
6	-	33.08249	-	-	-	-
7	30.8	30.74936	-	-	-	-
8	29.33	29.32298	-	52.3	-	114.9
9	28.36	28.35673	-	48.96	-	-
10	27.66	27.65726	-	46.234	-	-
11	27.13	27.12675	43.99	43.9754	-	102.70
12	26.71	26.71020	42.80	42.7890	-	97.234
14	26.10	26.09754	42.07	42.0559	-	94.1027
					Series (2.25)	Series (2.25)
					91.068	152.510
					67.9490	135.813
					59.5996	126.1024
					55.1577	119.6427
					52.35128	114.9656
					48.95695	111.39217
					46.23382	106.24872
					43.97543	102.69722
					42.78900	97.23448
					42.05590	94.10278

5%

n	p = 4		p = 6		p = 10	
	Korin	Series (2.17)	Korin	Davis	Korin	Davis
4	-	38.18424	-	-	-	-
5	-	29.01339	-	-	-	-
6	25.8	25.76297	-	-	-	-
7	24.06	24.06050	-	-	-	-
8	23.00	23.00227	-	-	-	-
9	22.28	22.27753	-	43.6	-	101.8
10	21.75	21.74883	-	40.919	-	-
11	21.35	21.34555	36.87	36.8638	-	91.28
12	-	21.02752	35.89	35.8847	-	86.516
13	20.77	20.77014	35.28	35.2774	80.7	80.7067
					Series (2.25)	Series (2.25)
					71.781	132.403
					55.4554	119.069
					49.24692	111.1458
					45.8294	105.7642
					43.6268	101.8159
					40.91920	98.77239
					38.71385	94.35388
					36.86380	91.27911
					35.88471	86.51565
					35.27737	80.70671

CHAPTER V

DISTRIBUTION OF THE LIKELIHOOD RATIO CRITERION
FOR TESTING $\Sigma = \Sigma_0, \mu = \mu_0$ 1. INTRODUCTION

Let a p -variate random sample of size N from the normal distribution with mean μ and covariance matrix Σ be denoted by x_1, x_2, \dots, x_N . The likelihood ratio criterion for testing the hypothesis $H_0: \Sigma = \Sigma_0$ and $\mu = \mu_0$ against alternatives $H_1: \Sigma \neq \Sigma_0$ or $\mu \neq \mu_0$, where Σ_0 is a given positive definite matrix and μ_0 a given vector, is expressed as (Anderson [1])

$$(1.1) \quad L = (e/N)^{\frac{Np}{2}} \frac{|\Sigma_0^{-1}|^{\frac{N}{2}} e^{-\frac{1}{2} \text{tr} \Sigma_0^{-1} \{ \Sigma + N(\bar{x} - \mu_0)(\bar{x} - \mu_0)' \}}}{|\Sigma \Sigma_0^{-1}|^{\frac{N}{2}}}$$

where $\Sigma = \sum_{\alpha=1}^N (x_{\alpha} - \bar{x})(x_{\alpha} - \bar{x})'$ and $\bar{x} = \sum_{\alpha=1}^N x_{\alpha} / N$. Although the exact distribution of L is unknown, it has been shown that the asymptotic distribution of $-2 \log L$ is a chisquare with $\frac{1}{2} p(p+1) + p$ degrees of freedom. No further information on the distribution of L appears to be available.

In this chapter, the null distribution of L is obtained first using the derivations for the chisquare distribution and then using

the derivations for the beta distribution by methods similar to those used in Chapter III. From these percentage points of L can be computed to any degree of accuracy even for small sample sizes. Tabulations of percentage points for $p = 2(1)6$ for various significance levels are given.

2. DISTRIBUTION OF L AS A CHISQUARE SERIES

The h -th moment of L under the null hypothesis is known to be (Anderson [1])

$$(2.1) \quad E(L^h) = (2e/N)^{Nhp/2} \left[\frac{\Gamma_p\{(n+Nh)/2\}}{\Gamma_p(n/2)} \right] \\ \cdot (1+h)^{-\frac{Np}{2}} (1+h)$$

where $n = N-1$. Let

$$(2.2) \quad \lambda = -2 \log L.$$

If $\phi(t)$ is the characteristic function of λ , then

$$(2.3) \quad \phi(t) = \frac{(2e/N)^{-Npit}}{(1-2it)^{-Np(1-2it)/2}} \frac{\prod_{\alpha=1}^p \Gamma\{\frac{N}{2} (1-2it)^{-\frac{\alpha}{2}}\}}{\prod_{\alpha=1}^p \Gamma(\frac{N-\alpha}{2})},$$

and therefore

$$(2.4) \quad \log \phi(t) = -Npit \log(2e/N) + \sum_{\alpha=1}^p \log \Gamma\{\frac{N}{2} (1-2it)^{-\frac{\alpha}{2}}\} \\ - \sum_{\alpha=1}^p \log \Gamma(\frac{N-\alpha}{2}) - \frac{Np}{2} (1-2it) \log(1-2it).$$

Using the expansion (2.64) of Chapter III to each gamma function in (2.4) we have

$$(2.5) \quad \log \phi(t) = \frac{p}{2} \log(2\pi) + \sum_{\alpha=1}^p \log \Gamma\left(\frac{N-\alpha}{2}\right) + \frac{pN}{2} \log(N/2e) \\ - \left(\frac{p^3+3p}{4}\right) \log \Gamma\left\{\frac{N}{2}(1-2it)\right\} + \sum_{\gamma=1}^m (Q_\gamma/N^\gamma) \left(\frac{1-2it}{2}\right)^{-\gamma} \\ + R'_{m+1}(N,t),$$

where the coefficients Q_r 's are given by

$$(2.6) \quad Q_r = (-1)^{r-1} \sum_{\alpha=1}^p B_{r+1}(-\alpha/2)/r(r+1).$$

The characteristic function of L can then be obtained from (2.5) as

$$(2.7) \quad \phi(t) = K(p,N) [N(1-2it)/2]^{-v} \left(\sum_{j=0}^{\infty} (B_j/N^j) \left(\frac{1-2it}{2}\right)^{-j} \right) \\ + R''_{m+1}(N,t),$$

where

$$(2.8) \quad K(p,N) = (2\pi)^{p/2} (N/2e)^{\frac{pN}{2}} \left[\prod_{\alpha=1}^p \Gamma\left(\frac{N-\alpha}{2}\right) \right]^{-1},$$

and

$$(2.9) \quad v = (p^2 + 3p)/4.$$

The coefficients B_j 's can be expressed in terms of Q_j 's using (2.67) of Chapter III.

Since $(1-2it)^{-\frac{n}{2}}$ is the characteristic function of a chisquare density with n degrees of freedom, say $g_n(x^2)$, the density of λ can be derived from (2.7) in the form

$$(2.10) \quad f(\lambda) = K(p,N) \sum_{j=0}^{\infty} (B_j) (2/N)^{j+v} g_{2(j+v)}(x^2) + R_{m+1}'(N).$$

The probability that λ is larger than any value, say λ_0 , is

$$(2.11) \quad P(\lambda \geq \lambda_0) = K(p,N) \sum_{j=0}^{\infty} (B_j) (2/N)^{j+v} G_{2(j+v)}(x^2) + R_{m+1}(N),$$

where $G_{2(j+v)}(x^2) = \int_{\lambda_0}^{\infty} g_{2(j+v)}(x^2) dx^2$

and
$$R_{m+1}(N) = \frac{1}{2\pi} k(p,N) \int_{\lambda_0}^{\infty} \int_{-\infty}^{\infty} e^{-it \lambda} \sum_{j=0}^{\infty} (B_j) (2/N)^{j+v} (1-2it)^{-(j+v)} \cdot [\exp\{R_{m+1}(N,t)\} - 1] dt d\lambda .$$

Thus from (2.11) we have that the distribution of λ may be obtained from a series of chisquare distributions.

3. DISTRIBUTION OF L AS A BETA SERIES

Let

$$(3.1) \quad \lambda = L^{2/N} .$$

Then from (2.1), we have

$$(3.2) \quad E(\lambda^h) = \frac{(2e/N)^{ph}}{\prod_{\alpha=1}^p \Gamma(\frac{N-\alpha}{2})} \frac{\prod_{\alpha=1}^p \Gamma\{\frac{N}{2} + h - \frac{\alpha}{2}\}}{(1+2h/N)^{Np(1+2h/N)/2}} .$$

Using inverse Mellin's transform, the density of λ is given by

$$(3.3) \quad f(\lambda) = \frac{1}{\prod_{\alpha=1}^p \Gamma(\frac{N-\alpha}{2})} \cdot \frac{1}{2\pi i} \int_{-i\infty}^{i\infty} \frac{\lambda^{-h-1} (2e/N)^{ph} \prod_{\alpha=1}^p \Gamma\{\frac{N}{2} + h - \frac{\alpha}{2}\}}{(1+2h/N)^{Np} (1+2h/N)^{/2}} dh.$$

Putting $\frac{N}{2} + h = t$ in (3.3), we have

$$(3.4) \quad f(\lambda) = K_1(p, N) \lambda^{\frac{N}{2}-1} \frac{1}{2\pi i} \int_{c-i\infty}^{c+i\infty} \lambda^{-t} C(t) dt$$

where $c = N/2$,

$$(3.5) \quad C(t) = (e/t)^{pt} \prod_{\alpha=1}^p \Gamma(t - \frac{\alpha}{2})$$

and

$$(3.6) \quad K_1(p, N) = (2e/N)^{-pN/2} \left[\prod_{\alpha=1}^p \Gamma(\frac{N-\alpha}{2}) \right]^{-1}.$$

Using the expansion (2.64) of Chapter III to each gamma function in (3.6), we have

$$(3.7) \quad \log C(t) = \frac{p}{2} \log(2\pi) - \frac{(p^2+3p)}{4} \log t \\ + [A_1/t + A_2/t^2 + \dots + A_r/t^r + \dots]$$

where the coefficients A_r 's are given by

$$(3.8) \quad A_r = (-1)^{r-1} \left[\sum_{\alpha=1}^p B_{r+1}(-\alpha/2) \right] / r(r+1).$$

Thus from (3.7), we have

$$(3.9) \quad C(t) = (2\pi)^{p/2} t^{-(p^2+3p)/4} [1+B_1/t+B_2/t^2+\dots+B_r/t^r+\dots]$$

where the coefficients B_r 's can be computed using (2.10) of Chapter IV.

Proceeding as in Chapter IV, it can be seen that when $v = (p^2+3p)/4$ is an integer, the probability that λ is less than any given value, say λ_0 , is

$$(3.10) \quad P(\lambda \leq \lambda_0) = K_1(p,N) (2\pi)^{p/2} \sum_{r=0}^{\infty} (B_r) I_{v+r-1,u}(\lambda_0)$$

where $u = \frac{N}{2} - 1$ and $I_{v+r-1,u}(\lambda_0)$ satisfies the recurrence relation (2.15) of Chapter IV. If $v = (p^2+3p)/4$ is not an integer. The distribution of λ is given by

$$(3.11) \quad P(\lambda \leq \lambda_0) = K_1(p,N) (2\pi)^{p/2} \sum_{i=0}^{\infty} R_i I_{\lambda_0} \left(\frac{N}{2} + \lambda, v+i \right) / \Gamma(v+i)$$

where $I_{\lambda_0}(p,q)$ is the incomplete beta function $\int_0^{\lambda_0} x^{p-1} (1-x)^{q-1} dx$.

4. COMPUTATIONS OF PERCENTAGE POINTS

.005, .01, .025, .05, .1 and .25 significance points of $\lambda = L^{2/N}$ were computed for $p = 2(1)8$ and various n using (2.11), (3.10) and (3.11) and these are presented in table 7. The computation was carried out on CDC 6500 using double precision arithmetic.

Table 8. Percentage Points of $\lambda = L^{2/N}$ for testing $\bar{x} = \bar{x}_0, \mu = \mu_0$

		p = 2					
$\alpha \backslash N$.005	.01	.025	.05	.1	.25	
4	.02 ² 11950	.02 ² 24255	.02 ² 62516	.012967	.027422	.078104	
5	.02 ² 90438	.014658	.028055	.046404	.078032	.16230	
6	.026253	.037975	.062461	.091968	.13730	.24199	
7	.051029	.068870	.10323	.14149	.19621	.31188	
8	.080540	.10371	.14595	.19050	.25119	.37189	
9	.11243	.13993	.18810	.23694	.30114	.42330	
10	.14505	.17593	.22840	.28001	.34605	.46750	
11	.17738	.21080	.26628	.31956	.38628	.50575	
12	.20877	.24408	.30155	.35570	.42233	.53910	
13	.23889	.27553	.33424	.38867	.45471	.56837	
14	.26754	.30509	.36447	.41878	.48387	.59425	
15	.29468	.33281	.39240	.44630	.51024	.61726	
16	.32031	.35874	.41824	.47151	.53415	.63785	
17	.34446	.38299	.44215	.49466	.55593	.65638	
18	.36720	.40568	.46431	.51597	.57582	.67313	
19	.38862	.42692	.48489	.53563	.59406	.68834	
20	.40879	.44682	.50404	.55381	.61083	.70221	
22	.44572	.48300	.53853	.58634	.64061	.72659	
24	.47862	.51499	.56871	.61457	.66622	.74730	
26	.50804	.54341	.59529	.63926	.68848	.76512	
28	.53448	.56881	.61886	.66104	.70798	.78060	
30	.55832	.59161	.63989	.68037	.72521	.79417	
32	.57992	.61218	.65876	.69764	.74054	.80617	
34	.59956	.63082	.67579	.71316	.75425	.81685	
36	.61749	.64778	.69121	.72718	.76660	.82642	
38	.63392	.66328	.70525	.73990	.77777	.83504	
40	.64902	.67750	.71808	.75150	.78793	.84285	
45	.68191	.70834	.74579	.77645	.80970	.85948	
50	.70923	.73385	.76858	.79687	.82742	.87294	
55	.73227	.75528	.78763	.81388	.84213	.88404	
60	.75195	.77355	.80380	.82827	.85454	.89336	
65	.76896	.78928	.81769	.84060	.86514	.90130	
70	.78379	.80299	.82975	.85129	.87430	.90814	
75	.79684	.81502	.84031	.86063	.88230	.91409	
80	.80842	.82567	.84965	.86887	.88934	.91932	
85	.81875	.83517	.85795	.87619	.89559	.92395	
90	.82803	.84369	.86539	.88274	.90117	.92807	
95	.83640	.85137	.87209	.88863	.90619	.93177	
100	.84400	.85834	.87815	.89396	.91072	.93511	

Table 8 (Continued)

p = 3

α						
N	.005	.01	.025	.05	.1	.25
5	.0 ³ 34244	.0 ³ 70203	.0 ² 18466	.0 ² 39217	.0 ² 85876	.026462
6	.0 ² 30928	.0 ² 50905	.010015	.017047	.029804	.067094
7	.010307	.015172	.025706	.038979	.060467	.11482
8	.02233	.030688	.047382	.066822	.096110	.16366
9	.038451	.050397	.072989	.097894	.13358	.21074
10	.057614	.072947	.10079	.13027	.17095	.25483
11	.078828	.097194	.12952	.16269	.20716	.29553
12	.10127	.12227	.15832	.19439	.24166	.33285
13	.12432	.14754	.18662	.22493	.27421	.36699
14	.14750	.17258	.21408	.25408	.30473	.39819
15	.17049	.19710	.24051	.28174	.33327	.42675
16	.19306	.22090	.26578	.30789	.35991	.45292
17	.21506	.24389	.28987	.33256	.38478	.47695
18	.23638	.26598	.31278	.35580	.40798	.49908
19	.25697	.28717	.33452	.37770	.42967	.51950
20	.27680	.30745	.35515	.39833	.44994	.53839
22	.31415	.34533	.39326	.43609	.48671	.57217
24	.34852	.37986	.42755	.46974	.51912	.60148
26	.38008	.41132	.45847	.49982	.54783	.62711
28	.40906	.44003	.48644	.52684	.57343	.64970
30	.43570	.46627	.51180	.55119	.59636	.66975
32	.46022	.49031	.53489	.57325	.61701	.68766
34	.48284	.51240	.55598	.59330	.63569	.70375
36	.50374	.53273	.57530	.61160	.65267	.71828
38	.52310	.55150	.59306	.62836	.66816	.73146
40	.54107	.56888	.60943	.64375	.68234	.74347
45	.58077	.60710	.64523	.67727	.71307	.76930
50	.61432	.63923	.67511	.70510	.73842	.79043
55	.64300	.66659	.70041	.72855	.75969	.80802
60	.66778	.69014	.72209	.74857	.77777	.82290
65	.68939	.71062	.74087	.76586	.79334	.83565
70	.70839	.72859	.75729	.78094	.80688	.84668
75	.72523	.74448	.77177	.79420	.81876	.85634
80	.74024	.75862	.78463	.80596	.82926	.86485
85	.75371	.77129	.79612	.81645	.83862	.87241
90	.76585	.78270	.80645	.82487	.84701	.87917
95	.77687	.79303	.81579	.83437	.85457	.88525
100	.78690	.80243	.82427	.84208	.86142	.89074

Table 8 (Continued)

p = 4

$N \backslash \alpha$.005	.01	.025	.05	.1	.25
6	.0 ³ 10678	.0 ³ 220703	.0 ³ 58981	.0 ² 12748	.0 ² 28695	.0 ² 93374
7	.0 ² 11061	.0 ² 18433	.0 ² 37061	.0 ² 64488	.011612	.027730
8	.0 ² 41278	.0 ² 61637	.010696	.016600	.026537	.053367
9	.0 ² 98136	.013688	.021655	.031256	.046289	.083241
10	.018242	.024272	.036006	.049390	.069303	.11510
11	.029144	.037447	.052960	.069939	.094254	.14743
12	.042098	.052653	.071755	.092003	.12014	.17933
13	.056655	.069348	.091745	.11488	.14626	.21023
14	.072395	.087064	.11241	.13804	.17213	.23983
15	.088956	.10541	.13336	.16112	.19743	.26798
16	.10604	.12409	.15430	.18385	.22196	.29465
17	.12340	.14286	.17501	.20605	.24560	.31983
18	.14084	.16154	.19534	.22761	.26829	.34358
19	.15823	.17999	.21519	.24846	.29004	.36596
20	.17544	.19812	.23449	.26856	.31075	.38705
22	.20901	.23314	.27127	.30645	.34941	.42568
24	.24112	.26627	.30555	.34134	.38454	.46010
26	.27157	.29741	.33736	.37339	.41647	.49087
28	.30029	.32655	.36682	.40282	.44552	.51849
30	.32728	.35376	.39409	.42987	.47201	.54339
32	.35261	.37916	.41934	.45476	.49624	.56594
34	.37637	.40287	.44275	.47772	.51845	.58642
36	.39864	.42501	.46449	.49893	.53887	.60511
38	.41955	.44570	.48470	.51858	.55769	.62222
40	.43917	.46506	.50353	.53680	.57508	.63794
45	.48326	.50835	.54532	.57705	.61326	.67213
50	.52127	.54544	.58085	.61103	.64528	.70050
55	.55429	.57752	.61137	.64007	.67248	.72440
60	.58320	.60549	.63784	.66515	.69585	.74479
65	.60868	.63007	.66100	.68701	.71615	.76240
70	.63129	.65182	.68141	.70622	.73393	.77774
75	.65148	.67119	.69953	.72323	.74963	.79123
80	.66961	.68855	.71573	.73839	.76360	.80319
85	.68597	.70419	.73028	.75200	.77609	.81385
90	.70080	.71835	.74343	.76426	.78734	.82342
95	.71432	.73123	.75537	.77538	.79752	.83206
100	.72668	.74300	.76625	.78550	.80677	.83990

Table 8 (Continued)

p = 5

$N \backslash \alpha$.005	.01	.025	.05	.1	.25
7	.0 ⁴ 3483	.0 ⁴ 7246	.0 ³ 1959	.0 ³ 4301	.0 ³ 9868	.0 ² 3360
8	.0 ³ 40379	.0 ³ 67996	.0 ² 13921	.0 ² 24668	.0 ² 45511	.011394
9	.0 ² 16608	.0 ² 25107	.0 ² 44458	.0 ² 70365	.011532	.024312
10	.0 ² 42873	.0 ² 60583	.0 ² 97873	.014408	.021870	.041164
11	.0 ² 85445	.011519	.017449	.024403	.035072	.060841
12	.014486	.018858	.027220	.036629	.050512	.082357
13	.022021	.027896	.038779	.050629	.067587	.10493
14	.030978	.038392	.051773	.065965	.085773	.12798
15	.041149	.050083	.065870	.082248	.10464	.15107
16	.052319	.062720	.080771	.099155	.12386	.17390
17	.064284	.076075	.096224	.11642	.14316	.19627
18	.076861	.089948	.11202	.13384	.16236	.21804
19	.089887	.10417	.12798	.15125	.18131	.23912
20	.10322	.11861	.14398	.16852	.19991	.25947
22	.13098	.14766	.17568	.20229	.23577	.29789
24	.15761	.17642	.20651	.23467	.26961	.33329
26	.18444	.20446	.23613	.26541	.30133	.36584
28	.21055	.23152	.26436	.29442	.33096	.39573
30	.23575	.25745	.29114	.32171	.35856	.42321
32	.25996	.28218	.31646	.34732	.38428	.44849
34	.28310	.30571	.34035	.37135	.40824	.47181
36	.30519	.32805	.36289	.39389	.43057	.49335
38	.32622	.34924	.38414	.41503	.45142	.51329
40	.34624	.36932	.40417	.43488	.47090	.53179
45	.39209	.41507	.44945	.47946	.51435	.57263
50	.43257	.45519	.48878	.51789	.55148	.60710
55	.46839	.49049	.52315	.55126	.58353	.63653
60	.50021	.52173	.55336	.58046	.61141	.66194
65	.52863	.54952	.58010	.60619	.63587	.68408
70	.55411	.57436	.60390	.62901	.65749	.70352
75	.57707	.59669	.62522	.64939	.67672	.72074
80	.59785	.61684	.64439	.66768	.69393	.73608
85	.61674	.63512	.66174	.68418	.70943	.74984
90	.63396	.65177	.67750	.69914	.72344	.76224

Table 8 (Continued)

p = 6

$N \backslash \alpha$.005	.01	.025	.05	.1	.25
8	.0 ⁴ ₂ 1175	.0 ⁴ ₂ 2449	.0 ⁴ ₂ 6667	.0 ³ ₁ 1493	.0 ³ ₂ 3467	.0 ² ₁ 1228
9	.0 ³ ₁ 14899	.0 ³ ₂ 25325	.0 ³ ₂ 52651	.0 ³ ₂ 94699	.0 ² ₁ 17839	.0 ² ₁ 46618
10	.0 ³ ₂ 66790	.0 ² ₁ 10206	.0 ² ₁ 18395	.0 ² ₂ 29611	.0 ² ₂ 49579	.010881
11	.0 ² ₁ 18575	.0 ² ₂ 26552	.0 ² ₂ 43693	.0 ² ₂ 65440	.010148	.019869
12	.0 ² ₂ 39475	.0 ² ₂ 53849	.0 ² ₂ 83094	.011822	.017350	.031265
13	.0 ² ₂ 70751	.0 ² ₂ 93190	.013700	.018747	.026384	.044617
14	.011289	.014467	.020474	.027168	.036988	.059468
15	.016571	.020769	.028500	.036886	.048878	.075407
16	.022853	.028122	.037615	.047684	.061779	.092085
17	.030043	.036401	.047647	.059352	.075444	.10922
18	.038032	.045474	.058431	.071697	.089657	.12657
19	.046710	.055212	.069812	.084550	.10424	.14397
20	.055967	.065492	.081652	.097763	.11903	.16128
22	.075822	.087246	.10624	.12479	.14879	.19522
24	.096888	.10999	.13143	.15201	.17821	.22783
26	.11861	.13315	.15665	.17889	.20685	.25885
28	.14055	.15633	.18154	.205119	.23443	.28816
30	.16241	.17923	.20583	.23047	.26082	.31575
32	.18397	.20164	.22937	.25484	.28595	.34165
34	.20507	.22344	.25206	.27816	.30982	.36596
36	.22560	.24454	.27386	.30042	.33244	.38875
38	.24550	.26489	.29475	.32163	.35387	.41014
40	.26474	.28448	.31473	.34182	.37415	.43022
45	.30982	.33010	.36084	.38809	.42027	.47532
50	.35069	.37114	.40190	.42892	.46059	.51418
55	.38764	.40803	.43848	.46506	.49600	.54791
60	.42105	.44121	.47117	.49717	.52727	.57742
65	.45131	.47114	.50049	.52583	.55504	.60341
70	.47877	.49822	.52688	.55153	.57983	.62647
75	.50377	.52279	.55073	.57468	.60209	.64704
80	.52660	.54517	.57238	.59563	.62215	.66550

CHAPTER VI

SUMMARY AND CONCLUSION

The study of the exact and asymptotic distributions of some statistics including some well-known likelihood ratio test criteria in multivariate analysis has been carried out in this dissertation. The main objective has been to present in convenient forms the distributions of these statistics in order to facilitate further work including numerical computations. This research was carried out for the reason that earlier authors who attempted to study these distributions either gave expressions which were practically not suitable for further use or merely restricted themselves to asymptotic derivations.

In the first chapter, the non-central distributions of statistics of the form $Y = \prod_{i=1}^p \theta_i^a (1-\theta_i)^b$, where a and b are real numbers have been obtained in the form of H-functions as a result of employing inverse Mellin transform and then asymptotic expansions of the distribution of Y have been obtained for suitable values of (a, b) . It may be possible to express the H-functions in a simple computational form by using the methods of contour integration as done in Chapter III at least in some cases and thus the power study of Y for some sets of values (a, b) may be made.

In Chapter II the moments of the sphericity criteria W were obtained in the null case and then use was made of the Mellin transform and Meijer's G-functions to find its non-central distribution in a closed form. Here the methods of contour integration may be employed to obtain the non-central distribution of W in a more suitable form for further work including numerical computations and thus power study of W can be facilitated.

In Chapter III the exact distribution of the sphericity criterion W was obtained in a form from which the percentage points of W were computed for $p = 2(1)10$ and various degrees of freedom n . The methods employed are quite general and could be used to obtain the exact distribution of other likelihood ratio criteria. In particular the methods can be applied to obtain the exact distribution of the likelihood ratio test criterion for testing the hypothesis that all off diagonal elements of Σ are zero, while the diagonal elements are equal in sets. Obviously the sphericity test considered in Chapter III is a special case of this hypothesis.

In Chapter IV, the density of the likelihood ratio criterion $L = \lambda_1^{2/n}$ for testing $\Sigma = \Sigma_0$ was obtained in a form which could be used to compute percentage points to any degree of accuracy even for small sized samples. The methods employed in this chapter along with the methods in Chapter I may yield an expression for the asymptotic non-central distribution of λ_1 in terms of central or non-central chisquare distributions.

In Chapter V, the methods of Chapter III and IV were further used to obtain the distribution of $\lambda = L^{2/N}$ for testing $\xi = \xi_0$, $\mu = \mu_0$ as a chisquare as well as a Beta series from which percentage points for $p = 2(1)6$ were computed. It may be possible that the methods used in this chapter together with the methods of Chapter I may yield a suitable form for the asymptotic distribution of λ .

In conclusion, the dissertation embodies the theoretical work solving once for all the distribution problems at least in the null case of some well-known likelihood ratio test criteria and makes available much needed tabulations which are fairly complete. The methods obtained to solve these distribution problems (See Chapter III) are of wide application.

LIST OF REFERENCES

REFERENCES

- [1] Anderson, T. W. (1958). Introduction to Multivariate Analysis, John Wiley, New York.
- [2] Box, G.E.P. (1949). A general distribution theory for a class of likelihood criteria. Biometrika, 36, 317-346.
- [3] Braaksma, B.L.J. (1964). Asymptotic expansions and analytic continuations for a class of Barnes-integrals. Composite Math. 15, 239-341.
- [4] Constantine, A. G. (1963). Some Non-Central Distribution problems in Multivariate Analysis. Ann. Math. Statist. 34, 1270-1285.
- [5] Consul, P. C. (1967). On the exact distribution of the criterion W for Testing Sphericity in a p-variate Normal Distribution. Ann. Math. Statist. 38, 1170-1174.
- [6] Consul, P. C. (1969). The Exact Distributions of Likelihood Criteria for Different Hypotheses. Multivariate Analysis, Vol. 2, Academic Press, New York.
- [7] Das Gupta, S. (1969). Properties of power functions of some tests concerning dispersion matrices of multivariate normal distributions. Ann. Math. Statist. 40, 697-701.
- [8] Davis, A. W. (1971). Percentile approximations for a class of likelihood ratio criteria, Biometrika, 58, 349-356.
- [9] Doetsch, G. (1971). Guide to the applications of the Laplace and Z-transformations. Van Nostrand-Reinhold, New York.
- [10] Erdélyi, A. et al (1953). Higher Transcendental Functions, Vol. I. McGraw-Hill, New York.
- [11] Goodman, N. R. (1963). Statistical Analysis Based on a certain Multivariate Complex Gaussian Distribution. Ann. Math. Statist. 34, 152-176.

- [12] James, A. T. (1964). Distribution of Matrix Variates and Latent Roots Derived from Normal Samples. Ann. Math. Statist. 35, 475-501.
- [13] John, S. (1972). The distribution of a statistic used for testing sphericity of normal distributions. Biometrika, 59, 169-173.
- [14] Khatri, C. G. (1965). Classical Statistical Analysis Based on Certain Multivariate Complex Gaussian Distributions. Ann. Math. Statist. 36, 98-114.
- [15] Khatri, C. G. (1966). On Certain Distribution Problems Based on Positive Definite Quadratic Functions in Normal Vectors. Ann. Math. Statist. 37, 468-479.
- [16] Khatri, C. G. (1967). Some distribution problems connected with the characteristic roots of $S_1 S_2^{-1}$. Ann. Math. Statist. 39, 944-948.
- [17] Korin, B. P. (1968). On the distribution of a statistic used for testing a covariance matrix. Biometrika 55, 171-178.
- [18] Luke, Y. L. (1969). The Special Functions and Their Approximations. Academic Press, New York.
- [19] Mathai, A. M. and Rathie, P. N. (1970). The exact distributions for the sphericity test. Jour. Statist. Res. 4, 140-159.
- ✓ [20] Mauchly, J. W. (1940). Significance Test for Sphericity of n-variate Normal Population. Ann. Math. Statist. 11, 204-209.
- [21] Meijer, G. (1946). Nederl Akad. Wetensch. Proc. 49, 344-456.
- [22] Nagao, H. (1967). Monotonicity of the modified likelihood ratio test for a covariance matrix. J. Sci. Hiroshima Univ. Ser. A-I 31, 147-150.
- [23] Nair, U. S. (1938). The Application of the Moment Function in the Study of Distribution Laws in Statistics. Biometrika, 30, 274-294.
- [24] Nair, U. S. (1940). Application of Factorial Series in the Study of Distribution Laws in Statistics. Sankhyā 5, 175.
- [25] Pillai, K.C.S., Al-Ani, S. and Jouris, G. M. (1969). On the distributions of the ratios of the roots of a covariance matrix and Wilks' criterion for tests of three hypotheses. Ann. Math. Statist. 40, 2033-2040.

- [26] Sugiura, N. (1969). Asymptotic expansions of the non-null distributions of the likelihood ratio criteria for multivariate linear hypothesis and independence. Ann. Math. Statist. 40, 942-952.
- [27] Sugiura, N. and Nagao, H. (1968). Unbiasedness of some test criteria for the equality of one or two covariance matrices Ann. Math. Statist. 39, 1686-1692.
- [28] Sugiura, N. and Fujikoshi, Y. (1969). Asymptotic expansions of the non-null distributions of the likelihood ratio criteria for multivariate linear hypothesis and independence. Ann. Math. Statist. 40, 942-952.
- [29] Titchmarsh, E. C. (1948). Introduction to the Theory of Fourier Integrals. Oxford University Press, London.
- [30] Tukey, J. W. and Wilks, S. S. (1946). Approximation of the Distribution of the Product of Beta Variables by a Single Beta Variable. Ann. Math. Statist. 17, 318-324.

APPENDICES

Appendix A

Coefficients B_i 's ($i = 11, \dots, 14$) are:

$$B_{11} = Q_{11} + 8Q_1Q_2Q_8/11 + 2Q_2B_9/11 + 8Q_3Q_8/11 \\ + 9Q_9B_2/11 + 4Q_8Q_1^3/33 + Q_1B_{10}/11 + 7Q_2Q_1^2Q_7/22 \\ + 10Q_{10}Q_1/11 + 7Q_1^4Q_7/264 + 3Q_3B_8/11 + 7Q_4Q_7/11 \\ + 6Q_6B_5/11 + 7Q_3Q_1Q_7/11 + 5Q_5B_6/11 + 7Q_2^2Q_7/22$$

$$B_{12} = Q_{12} + Q_1B_{11}/12 + Q_2B_{10}/6 + Q_3B_9/4 + Q_4B_8/3 + 5Q_5B_7/12 + Q_6B_6/2 \\ + 7Q_7B_5/12 + 2Q_8B_4/3 + 3Q_9B_3/4 + 5Q_{10}Q_1^2/12 + 5Q_{10}Q_2/6 + 11Q_{11}B_1/12$$

$$B_{13} = Q_{13} + Q_1B_{12}/13 + 2Q_2B_{11}/13 + 3Q_3B_{10}/13 + 4Q_4B_9/13 + 5Q_5B_8/13 + 6Q_6B_7/13 \\ + 7Q_7B_6/13 + 8Q_8B_5/13 + 9Q_9B_4/13 + 10Q_{10}B_3/13 + 11Q_{11}Q_1^2/26 \\ + 11Q_{11}Q_2/13 + 12Q_{12}B_1/13$$

$$B_{14} = Q_{14} + Q_1B_{13}/14 + 3Q_3B_{11}/14 + Q_2B_{12}/7 + 2Q_4B_{10}/7 + 5Q_5B_9/14 + 3Q_6B_8/7 \\ + 3Q_6B_8/7 + Q_7B_7/2 + 4Q_8B_6/7 + 9Q_9B_5/14 + 5Q_{10}B_4/7 + 11Q_{11}B_3/14 \\ + 3Q_{12}Q_1^2/7 + 6Q_{12}Q_2/7 + 13Q_{13}B_1/14$$

Coefficients C_{ij} 's ($j = 11, \dots, 14$) are:

$$C_{i11} = -(t_i^{12} - 6t_i^{11} + 11t_i^{10} - 33t_i^8/2 + 22t_i^6 - 33t_i^4/2 + 5t_i^2)$$

$$C_{i12} = t_i^{13} - 13t_i^{12}/2 + 13t_i^{11} - 143t_i^9/6 + 288t_i^7/7 - 429t_i^5/10 + 65t_i^3/3 - 691t_i/210$$

$$C_{i13} = -(t_i^{14} - 7t_i^{13} + 9t_i^{12}/6 - 100t_i^{10}/30 + 143t_i^8/2 - 100t_i^6/10 + 455t_i^4/6 \\ - 691t_i^2/30)$$

Coefficients R_{is} ($i = 11, \dots, 14$) are:

$$R_{11} + R_{10}^d{}_{10,1} + R_9^d{}_{92} + R_8^d{}_{83} + R_7^d{}_{74} + R_6^d{}_{65} + R_5^d{}_{56} + R_4^d{}_{47} + R_3^d{}_{38} \\ + R_2^d{}_{29} + R_1^d{}_{1,10} + R_0^d{}_{0,11} = B_{11}$$

$$R_{12} + R_{11}^d{}_{11,1} + R_{10}^d{}_{10,2} + R_9^d{}_{93} + R_8^d{}_{84} + R_7^d{}_{75} + R_6^d{}_{66} + R_5^d{}_{57} + R_4^d{}_{48} \\ + R_3^d{}_{39} + R_2^d{}_{210} + R_1^d{}_{1,11} + R_0^d{}_{0,12} = B_{12}$$

$$R_{13} + R_{12}^d{}_{12,1} + R_{11}^d{}_{11,2} + R_{10}^d{}_{10,3} + R_9^d{}_{94} + R_8^d{}_{85} + R_7^d{}_{76} + R_6^d{}_{67} \\ + R_5^d{}_{58} + R_4^d{}_{49} + R_3^d{}_{3,10} + R_2^d{}_{2,11} + R_1^d{}_{1,12} + R_0^d{}_{0,13} = B_{13}$$

$$R_{14} + R_{13}^d{}_{13,1} + R_{12}^d{}_{12,2} + R_{11}^d{}_{11,3} + R_{10}^d{}_{10,4} + R_9^d{}_{95} + R_8^d{}_{86} + R_7^d{}_{77} \\ + R_6^d{}_{68} + R_5^d{}_{59} + R_4^d{}_{4,10} + R_3^d{}_{3,11} + R_2^d{}_{2,12} + R_1^d{}_{1,13} \\ + R_0^d{}_{0,14} = B_{14}$$

Appendix B

Coefficients A_{ij} 's ($j = 11, \dots, 15$) are:

$$A_{i11} = -[(t_i^{12} - \lambda^{12}) - 6(t_i^{11} - \lambda^{11}) + 11(t_i^{10} - \lambda^{10}) - 33(t_i^8 - \lambda^8)/2 + 22(t_i^6 - \lambda^6) - 33(t_i^4 - \lambda^4)/2 + 5(t_i^2 - \lambda^2)]/132$$

$$A_{i12} = [(t_i^{13} - \lambda^{13}) - 13(t_i^{12} - \lambda^{12})/2 + 13(t_i^{11} - \lambda^{11}) - 143(t_i^9 - \lambda^9)/6 + 286(t_i^7 - \lambda^7)/7 - 429(t_i^5 - \lambda^5)/10 + 65(t_i^3 - \lambda^3)/3 - 691(t_i - \lambda)/210]/156$$

$$A_{i13} = -[(t_i^{14} - \lambda^{14}) - 7(t_i^{13} - \lambda^{13}) + 91(t_i^{12} - \lambda^{12})/6 - 1001(t_i^{10} - \lambda^{10})/30 + 143(t_i^8 - \lambda^8)/2 - 1001(t_i^6 - \lambda^6)/10 + 455(t_i^4 - \lambda^4)/6 - 691(t_i^2 - \lambda^2)/30]/182$$

$$A_{i14} = [(t_i^{15} - \lambda^{15}) - 15(t_i^{14} - \lambda^{14})/2 + 35(t_i^{13} - \lambda^{13})/2 - 91(t_i^{11} - \lambda^{11})/2 + 715(t_i^9 - \lambda^9)/6 - 429(t_i^7 - \lambda^7)/2 + 455(t_i^5 - \lambda^5)/2 - 691(t_i^3 - \lambda^3)/6 + 35(t_i - \lambda)/2]/210$$

$$A_{i15} = -[(t_i^{16} - \lambda^{16}) - 8(t_i^{15} - \lambda^{15}) + 20(t_i^{14} - \lambda^{14}) - 182(t_i^{12} - \lambda^{12})/3 + 4004(t_i^{10} - \lambda^{10})/21 - 1287(t_i^8 - \lambda^8)/3 + 20020(t_i^6 - \lambda^6)/33 - 1382(t_i^4 - \lambda^4)/3 + 140(t_i^2 - \lambda^2)]/240$$