

A Statistical Selection Approach to Binomial Models

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### Introduction

A common problem that arises in practice is the comparison of several Bernoulli processes (or populations) with unknown parameters  $p_1, \dots, p_k$ , respectively, where the  $p_i$ 's denote the success probabilities. For example, in examining the output at manufacturing processes it may be of interest to rank (or select) processes based on the estimated proportion of conforming output. In a different context, utilizing customer survey data expressing satisfaction with service at a car dealership it may be of interest to rank (or select) dealers based on the estimated proportion of satisfied customers. In all such examples the decision will be based on sample data and, consequently, statistical criteria should be specified as part of the investigation. We will deal with one such specification here.

Another particular realization of this problem is the critical issue of vendor selection. Deming (1982) notes the importance of vendor selection in a company's efforts to achieve high quality and productivity. In his 14 points, Deming's point 4 suggests the reduction of the number of suppliers to a subset of vendors who can furnish statistical evidence of dependable quality.

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## A Statistical Selection Approach to Binomial Models

### ABSTRACT

A subset selection rule is considered for selecting the best of  $k$  binomial populations (as determined by the binomial probability parameter). Let  $X_i$  denote the number of conforming items in a sample of size  $n$  from the  $i$ -th population (success probability  $p_i$ ),  $i = 1, \dots, k$ . The rule selects the  $i$ -th population if and only if  $X_i \geq \max_{1 \leq j \leq k} X_j - d$ , where  $d$  is a nonnegative integer. Operating characteristics are studied for slippage and equi-spaced parametric configurations. Tables and graphs relating to selection probabilities and expected subset size are presented as well as examples for illustrating use of these. Also, a new rule is discussed for selecting populations when bounds on the probability parameters are available.

Key Words: Binomial model, subset selection rules, operating characteristics, comparison with a control, application to vendor selection.

Vendor selection involves a consideration of many aspects -- cost, service, reliability, and quality. Pettit (1984) described the approach that 3M Corporation uses in the evaluation of prospective suppliers. It consists of evaluating potential vendors in four areas: quality, price, performance, and facility capabilities. While quality is explicitly considered in this approach, it is not evaluated in a statistical sense.

This article illustrates how statistical ranking and selection methodology can be utilized as one objective evaluation tool in this important decision setting.

To formalize the above problems consider  $k$  Bernoulli processes, and let  $p_i$  denote the "success" probability (i.e., sampled item conforms to specifications) of the  $i$ th population. The  $i$ th population we will denote simply by  $\pi_i$ . Let  $p_{[1]} \leq \dots \leq p_{[k]}$  denote the ordered parameters. It is assumed that there is no prior knowledge regarding the correct pairings of the ordered and the unordered  $p_i$ 's. The populations are ranked according to the values of  $p_i$ 's, and that associated with  $p_{[k]}$ , the largest  $p_i$ , is called the best.

Let  $X_1, X_2, \dots, X_k$  denote the number of conforming items from these populations based on a random sample of  $n$  items from each. Our interest is to define a statistical procedure based on  $X_1, \dots, X_k$  to select a nonempty subset of the  $k$  populations with a guarantee of minimum probability  $P^*$  that the best is included in the selected subset. Selection of any subset which includes the best is called a correct selection (CS). Thus the probability of a correct selection using a rule  $R$ ,  $P(\text{CS}|R)$ , should satisfy the condition that

$$P(\text{CS}|R) \geq P^* \quad (1)$$

whatever be the unknown values of the  $p_i$ 's. This condition is generally referred to as the  $P^*$ -condition. Obviously, for a meaningful problem,  $1/k < P^* < 1$ .

Any procedure  $R$  that satisfies (1) is a valid procedure. To distinguish between valid procedures we need to evaluate criteria that characterize effectively procedure performance. One such criterion is the expected value of  $S$ , the number of populations included in the selected subset.  $S$  is known as the subset size and it is a positive integer-valued random variable. One may also consider the related quantity  $E(S')$ , where  $S'$  denotes the number of non-best populations included in the selected subset. It should be pointed out that one would like to have  $E(S)$  (or  $E(S')$ ) as small as possible subject to (1). Let  $\alpha_i$  denote the probability of selecting the process associated with  $p_{[i]}$ ,  $i = 1, \dots, k$ . Obviously,  $\alpha_k = \text{PCS}$ . It is also easy to see that

$$E(S) = \alpha_1 + \dots + \alpha_k \quad (2)$$

$$E(S') = \alpha_1 + \dots + \alpha_{k-1}.$$

The  $\alpha_i$ 's are called the individual selection probabilities. One may also consider a criterion which combines  $E(S)$  and  $\text{PCS}$ . Such a criterion, namely,  $E(S)/\text{PCS}$  has been considered in the literature. All these criteria that are used to evaluate a valid procedure are called operating characteristics of the procedure. In our present study, we use the expected subset size and the individual selection probabilities.

A related approach for selecting a single best vendor (or a manufacturing process) having the highest percentage of conforming product under the indifference zone formulation has been studied by Sobel and Huyett (1957). Some adaptive and sequential versions of the same problem have been discussed by Bechhofer and Turnbull (1977) and by Bechhofer (1984). Gibbons (1982) provided a general introduction to selection procedures including the subset approach. Her examples focused primarily on selection with respect to means of normal probability models.

### The Gupta-Sobel Rule

Gupta and Sobel (1960) proposed and studied a rule  $R_B$  defined as follows.

$R_B$ : Select  $\pi_i$  if and only if  $X_i \geq \max_{1 \leq j \leq k} X_j - d$ ,

where  $d = d(k, n, P^*)$  is the smallest nonnegative integer satisfying

$$\inf_{\Omega} P(\text{CS} | R_B) \geq P^*, \quad (3)$$

where  $\Omega = \{p | p = (p_1, \dots, p_k), 0 \leq p_i \leq 1, i = 1, \dots, k\}$  is the parameter space. Gupta and Sobel (1960) have shown that the infimum on the left-hand side of (3) is attained when  $p_1 = \dots = p_k$ . Thus, we evaluate  $P(\text{CS} | R_B)$  for  $p_1 = \dots = p_k = p$  (say) and rewrite (3) as

$$\inf_{0 \leq p \leq 1} \sum_{j=0}^n \binom{n}{j} p^j (1-p)^{n-j} \left\{ \sum_{y=0}^u \binom{n}{y} p^y (1-p)^{n-y} \right\}^{k-1} \geq P^* \quad (4)$$

where  $u = \min(d+j, n)$ . There is no known analytical result regarding the value of  $p$  for which the infimum in (4) is attained except in the special case of  $k = 2$ . When  $k = 2$ , the infimum is attained for  $p = 0.5$ .

We have evaluated the algebraic expression in (4) as a function of  $p$ ,  $p = 0(.01)1.0$ , to further investigate where the infimum occurs and to compare the corresponding infimum of the probability of a correct selection with that computed at  $p = 1/2$ . Calculations were made for  $k = 2(1)5(5)15$ ;  $n = 2(1)5(5)35$ ;  $d = 0(1)n$ . The results are summarized in Figures 1 and 2. The value of  $p$  at which the infimum in (4) is attained is given in Figure 1 as a function of  $n$  and ranges between 0.26 and 0.50. However, as shown in Figure 2, for these calculations the infimum of the probability of a correct selection is very close to that value calculated assuming that the common value of  $p$  is 0.5 -- especially at the higher probability values.

When PCS calculated at the common  $p$  value of 0.5 exceeds 0.75 the actual infimum value of PCS is within 0.14% of this nominal calculation for  $k=3$  and all  $n$ ; within 0.3% for  $k=5$  and  $n \geq 3$ ; within 1.4% for  $k=10$  and  $n \geq 3$ ; and within

3.4% for  $k=15$  and  $n \geq 3$ . Thus, using a rule where  $d$  is determined assuming a common  $p$  value of 0.5 would result in an actual lower bound on the PCS only somewhat smaller than the nominal value. For large  $n$ , (4) can be approximated by

$$\inf_{0 \leq p \leq 1} \int_{-\infty}^{\infty} \Phi^{k-1} [x + (d + .5)/(npq)^{\frac{1}{2}}] \varphi(x) dx = P^*,$$

where  $q = 1-p$ . The infimum of the expression on the left hand side above occurs at  $p = \frac{1}{2}$  which gives the approximation for  $d$  as the solution to

$$\int_{-\infty}^{\infty} \Phi^{k-1} [x + (2d+1)/(n^{\frac{1}{2}})] \varphi(x) dx = P^*, \quad (5)$$

where  $\Phi$  and  $\varphi$  denote the cdf and density of a standard normal variable. Since  $d$  as obtained from (5) is not necessarily an integer, to implement the procedure we simply replace  $d$  by the smallest integer greater than or equal to  $d$ .

These values have been tabulated by Gupta and Sobel (1960), for  $k = 2(1)20(5)50$  and  $n = 1(1)20(5)50(10)100(25)200(50)500$ . Tables 1 and 2, extracted from Gupta and Sobel (1960), provide the values of  $d$  for  $P^* = .90$  and  $.95$ , respectively, for  $k = 2, 5(5)30(10)50$ , and  $n = 5(5)50(10)100, 250, 500$ .

### Operating Characteristics

Let us assume without loss of generality that  $p_1 \leq \dots \leq p_k$ . As we pointed out earlier, we consider the rule: Select  $\pi_j$  if and only if  $X_j \geq \max_{1 \leq j \leq k} X_j - d$ , where  $0 \leq d \leq n$ . The operating characteristics studied are the expected subset size and the individual selection probabilities. We consider two types of parametric configurations, namely, (1) the slippage configuration defined by  $p = p_1 = \dots = p_{k-1} = p_k - \delta$ ,  $0 < \delta < 1-p$ , and (2) the equi-spaced parametric configuration defined by  $p_{i+1} - p_i = \delta$ ,  $i = 1, \dots, k-1$ ,  $0 < \delta < (1-p)/(k-1)$ . For convenience, let

$$b(x; n, p) = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, 1, \dots, n$$

$$B(t; n, p) = \sum_{x=0}^t b(x; n, p), \quad t = 0, 1, \dots, n. \quad (6)$$

### Slippage Configurations

For the configuration  $(p, p, \dots, p, p+\delta)$ ,  $0 < \delta < 1-p$ , we get

$$\text{PCS} \equiv \alpha_k = \sum_{x=0}^n b(x; n, p+\delta) [B(x+d; n, p)]^{k-1},$$

$$\alpha_i = \sum_{x=0}^n b(x; n, p) B(x+d; n, p+\delta) [B(x+d; n, p)]^{k-2},$$
(7)

$$i = 1, \dots, k-1.$$

Any specified non-best population has the same probability of being selected and we denote this by  $P(\text{NCS})$ . Also,  $E(S) = (k-1)\alpha_1 + \text{PCS}$ .

We present tables and graphs for the operating characteristics in the case of three slippage configurations. These are given by the following pairs of  $p$  and  $\delta$  values:

$$(I) \quad p = .50, \delta = .10, \quad (II) \quad p = .75, \delta = .05, \quad (III) \quad p = .90, \delta = .03.$$

Tables 3 through 5 give the values of PCS,  $P(\text{NCS})$ , and  $E(S)$  for  $k = 3, 5, 10, 15$ ;  $d = 2, 3, 4, 5$ ; and  $n = 5(5)50(10)100, 250, 500$  in the case of the three configurations I - III. Figure 3 shows the graph of  $E(S)$  as a function of  $n$  for the rule with  $d = 2$  for  $k = 3, 5, 10$  when the slippage configuration is given by  $p = .90$  and  $\delta = .03$ . This figure also shows for  $n = 10(10)50$ , the value of PCS when  $\delta = 0$ , that is, when all the parameters are equal to .90. It should be noted that the PCS values, for a fixed value of  $k$ , decrease in this figure because the value of  $d$  is fixed at 2 throughout these calculations. This behavior is clearly understood (for large values of  $n$ ) by referring to equation (5). The limit of this expression is  $1/k$ . Figures 4 and 5 are graphs of  $E(S)$  as a function of  $n$  for  $d = 2, 3, 4, 5$ , and for  $k = 3, 5$ , and 10. Figure 4 is for the slippage configuration with  $p = .75$  and  $\delta = .05$  and Figure 5 is for the configuration with  $p = .90$  and  $\delta = .03$ . These results and examples are discussed in the next section.



For sufficiently large  $n$ , one can use the normal approximation and obtain

$$\begin{aligned} \text{PCS} &\approx \int_{-\infty}^{\infty} \phi^{k-1} \left[ x \sqrt{\frac{(p+\delta)(q-\delta)}{pq}} + \frac{d+n\delta+1/2}{\sqrt{npq}} \right] \varphi(x) dx, \\ \alpha_i &\approx \int_{-\infty}^{\infty} \phi^{k-2} \left[ x + \frac{d+1/2}{\sqrt{npq}} \right] \phi \left[ \sqrt{\frac{pq}{(p+\delta)(q-\delta)}} x + \frac{1/2+d-n\delta}{\sqrt{n(p+\delta)(q-\delta)}} \right] \varphi(x) dx, \\ & \quad i = 1, \dots, k-1. \end{aligned} \quad (8)$$

Note that in equation (8),  $1/2$  is due to the continuity correction in approximating the binomial by the normal.

#### Equi-spaced Parametric Configuration

For the configuration  $(p, p+\delta, \dots, p+(k-1)\delta)$ ,  $0 < \delta < (1-p)/(k-1)$ , we have

$$\alpha_i = \sum_{x=0}^n b(x; n, p+(i-1)\delta) \prod_{j \neq i} B(x+d; n, p+(j-1)\delta), \quad i = 1, \dots, k. \quad (9)$$

We note that  $\alpha_i$  is the probability of including the non-best population with parameter  $p + (i-1)\delta$ ,  $i = 1, \dots, k-1$ , and  $\alpha_k$  is the PCS. For large  $n$ , the normal approximation yields

$$\alpha_i \approx \int_{-\infty}^{\infty} \prod_{j \neq i} \phi \left[ \sqrt{\frac{\theta_j(1-\theta_j)}{\theta_j(1-\theta_j)}} x + \frac{d+1/2+(i-j)n\delta}{\sqrt{n\theta_j(1-\theta_j)}} \right] \varphi(x) dx, \quad i = 1, \dots, k, \quad (10)$$

where  $\theta_j = p+(j-1)\delta$ ,  $j = 1, \dots, k$ .

#### A Modified Procedure $R'_B$

Suppose, the experimenter has the a priori information that for all populations the unknown probabilities  $p_i$ 's are at least as large as  $p_0$  where  $p_0$  is some specified number and which in many situations can be assumed to be greater than  $\frac{1}{2}$ . Then, intuitively speaking, one should be able to use this information to reduce the  $d$ -value, for fixed values of  $P^*$  and  $n$ . This can be shown as follows:

In the least favorable case, i.e. when  $p_1 = p_2 = \dots = p_k = p$ , and  $n$  is large, we have

$$PCS = \int_{-\infty}^{\infty} \phi^{k-1} \left( x + \frac{d+1/2}{\sqrt{npq}} \right) \varphi(x) dx,$$

so that for  $n \rightarrow \infty$ , the infimum of the PCS takes place at  $p = \frac{1}{2}$ . Since the PCS given above decreases with  $p$  for values of  $p > \frac{1}{2}$ , it follows that for  $p_0 > \frac{1}{2}$ ,

$$\begin{aligned} \inf_{0 \leq p \leq 1} PCS &= \int_{-\infty}^{\infty} \phi^{k-1} \left( x + \frac{2d+1}{\sqrt{n}} \right) \varphi(x) dx \\ &< \int_{-\infty}^{\infty} \phi^{k-1} \left( x + \frac{d+1/2}{\sqrt{np_0q_0}} \right) \varphi(x) dx, \text{ where } q_0 = 1-p_0. \end{aligned}$$

Equating the two integrals above to  $P^*$  and relabelling the  $d$ -value in the second integral as  $d^*$ , we have

$$d^* = (2d+1)\sqrt{p_0q_0} - 1/2 < d.$$

Thus, for fixed  $n$  and  $P^*$ , the a priori constraint on  $p_i$ 's leads one to use the following modified procedure,

$R_B^1$ : Select the  $i$ th population if and only if

$$X_i \geq \max_{1 \leq j \leq k} X_j - d^*.$$

The modified procedure  $R_B^1$  will result in a smaller value of the expected size,  $E(S)$ , keeping  $n$  and  $P^*$  fixed. If one is willing to give up the saving in the value of  $E(S)$ , one can, for a fixed  $P^*$ , find a smaller  $n$  corresponding to this smaller value  $d^*$  of  $d$ . This can be done by interpolation in Tables 1 and 2.

### Comparison with a Control

In some situations, one may want to compare several populations with a control. The goal is to select all populations which are better than (that is, having higher p-value) the control. Based on random samples of  $n$  items, let  $X_1, \dots, X_m$  denote the numbers of conforming items from  $m$  populations and let  $X_0$  denote the number for the control. This problem was studied separately by Gupta and Sobel (1958). Their rule is

$R_{BC}$ : Select the population with  $X_i$  conforming items if and only if

$$X_i \geq X_0 - D,$$

where  $D = D(m, n, P^*)$  is the smallest nonnegative integer such that with specified probability  $P^*$  the selected subset will include all populations better than the control. For selected values of  $m$ ,  $n$ , and  $P^*$ , the value of  $D$  can be obtained from Tables 1 and 2 by setting  $m = k-1$ .

### Examples

For the purpose of illustrating our rule and the use of the tables, let us assume that we have five processes for an item. Our goal is to identify a subset of these in such a manner that the best is contained in the subset with a high probability. Having identified this subset, we will then proceed to investigate other nonstatistical criteria (such as price) upon which to base a final decision on process selection. We note that this approach is applicable only if test samples of the item can be obtained. For the five processes, let  $X_i$  denote the number of conforming items based on random samples of size  $n=30$ . (We will say more about the sample size choice later). Suppose that

$$X_1 = 27, X_2 = 25, X_3 = 24, X_4 = 22, \text{ and } X_5 = 28.$$

In simple terms, process 1 generated 30 test items (chosen at random) and 27 of the 30 items conformed satisfactorily to all specifications.

Now we use the statistical selection procedure  $R_B$  with  $d = 2$  to select a subset. (We will say more about the choice of  $d$  later.)

The rule can now be simply stated as: choose all processes for which  $X_i \geq \max X_j - d = 28 - 2 = 26$ . This results in the selection of processes 1 and 5. How good is this procedure? What probabilistic guarantees do we have with its use? That is where our tables and figures are helpful as we will now illustrate.

In the event that four of the processes could produce 90% conforming items (i.e.,  $p = .90$ ) and one could produce 93% conforming items, the selection rule  $R_B$  as we used it ( $n = 30$ ,  $d = 2$ ,  $k = 5$ ) would select the best process with probability 0.86 and would retain a nonbest process with probability 0.66 (see Table 5). The expected size of the selected subset can be read from either Table 5 or Figure 3 and is  $4(.66) + 0.86 = 3.5$ . Also from Figure 3 we find the probability of making a correct selection (i.e., choosing the best process to be in the selected subset) decreases to 0.702 as the best process decreases to 90% conformance -- the same as the other four.

If these operating characteristics are not satisfactory from the decision-maker's perspective then alternative choices for  $n$  and/or  $d$  should be made. Note, however, that all of the probabilities given in the preceding paragraph were obtainable before any data was obtained. The operating characteristics of the selection procedure are determined prior to the actual data analysis. Let us look at how alternative choices of  $n$  and  $d$  can be generated so as to meet a decision-maker's requirements or

preferences. This search and specification is usually conditioned on some statement about the parameter configuration over which the probabilistic statements should be applicable.

For example, if we now focus our concern on parameters in a slippage configuration with  $p = .75$  and  $\delta = .05$  we can look for a pair  $(n,d)$  for which PCS is at least a specified number -- say 0.90. Since this criterion will yield more than one  $(n,d)$  choice we might then choose the pair which has the smallest  $E(S)$ . Consulting Table 4 we generate the options listed below:

n	d	PCS	E(S)
5	2	.96	4.65
10	3	.96	4.56
15	3	.92	4.13
20	4	.95	4.32
25	4	.93	4.06
30	4	.91	3.82
35	4	.90	3.62
40	5	.93	3.91
45	5	.93	3.74
50	5	.92	3.60

It should be noted that because of the discrete nature of the distribution involved, an increase in  $n$  does not produce necessarily a better option. In this illustration the best option would be  $n = 50$  and  $d = 5$ . That is, ask for a random sample of 50 items from each process and select those for which

$$X_i \geq \max_{1 \leq j \leq k} X_j - 5.$$

Alternatively, one may want to set an upper bound for  $E(S)/k$ , the expected proportion of populations selected. If we set this bound as .80, then we look for pairs  $(n,d)$  for which  $E(S) \leq 5 \times .80 = 4$ . If there are more

than one such pair with same  $n$ , we take the pair for which the PCS is maximum. Consulting Table 4 again, we have the following options -- the best being  $n = 45$  and  $d = 5$ .

$n$	$d$	$E(S)$	PCS
10	2	3.89	.87
15	2	3.37	.81
20	3	3.77	.88
25	3	3.48	.86
30	4	3.82	.91
35	4	3.62	.90
40	5	3.91	.93
45	5	3.74	.93
50	5	3.60	.92

It is possible to use other criteria for choosing the pair  $(n,d)$ . If we feel that the true parametric configuration can in some sense be described by one of two possible slippage configurations given by, say,  $p = .75$ ,  $\delta = .05$  and  $p = .90$ ,  $\delta = .03$ , then we can choose the pair  $(n,d)$  that controls the PCS or  $E(S)$  at given levels for both configurations.

#### Summary and Concluding Remarks

In this paper we have presented two statistical selection rules applicable to the binomial model. The first rule is appropriate for the selection of a subset to contain the best population with a preassigned probabilistic guarantee. The second rule is directed towards a selection of a subset to contain all populations better than a standard -- again with a specified probabilistic guarantee. Additionally we have indicated how prior knowledge can be explicitly incorporated in the form of inequality constraints on the binomial probability parameters. Such incorporation, where applicable, can reduce substantially the expected subset size while preserving the stated minimum probability of making a correct selection.

Implementation of such procedures requires several choices by the analyst. That is, in a sense, similar to consideration involved in statistical hypothesis testing. In the latter case the analyst determines a critical region (or rejection region) and sample size by examining operating characteristics (e.g., Type I and Type II errors) and choosing combinations appropriate for the application. With respect to the selection procedures herein discussed the analyst must choose the constant  $d$  to be used with the rule  $R_B$  and the sample size for each process.

Once the number of populations ( $k$ ) is specified the choice of  $d$  and  $n$  depends in turn on operating characteristics of the selection procedure. (For rule  $R_{BC}$  the choice is  $D$  and  $n$ ). We recommend the analyst first specify a  $P^*$  value which is the minimum probability of a correct selection (the analog of Type I error). This specification can generate many  $(d,n)$  combinations. At this point the analyst should specify an upper bound on the expected subset size for a parametric configuration meaningful for the application (the analog of Type II error). Then referring to the figures and tables given here, determine a  $(d,n)$  choice which achieves the requirements on both the probability of a correct selection and the expected subset size. In situations where these tables and figures are not sufficient to represent an application, the reader is referred to the additional references. New calculations may be required using the formulae given.

Once the  $d$  value and sample size  $n$  have been determined the analyst proceeds by random sampling and testing of  $n$  items from each population and then selecting a subset according to the rule  $R_B$  with  $d$  as the constant. The resultant subset, chosen on the basis of a statistical comparison, can then be examined further on other important aspects (such as price, facilities, delivery, etc. for vendor selection).

Statistical methods can play a significant role in vendor selection. Those described here are applicable only to those situations where vendors

are currently producing the product of interest. Since the rules are data dependent, they would not be applicable for decision situations involving new products currently not being produced.

We have illustrated here techniques applicable to attribute data represented by the binomial model. Similar procedures have been developed for continuous measurement data emanating from a wide variety of statistical distributions such as normal, gamma, and exponential. A good discussion of these many rules can be found in the book by Gupta and Panchapakesan (1979). Distribution-free (nonparametric) rules have also been developed and can be applied when only ordinal information is obtained about the vendors (or processes). A review of such procedures can be found in Gupta and McDonald (1982).

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Table 1. Values of  $d$  for implementing the rule  $R_B$  for selecting the best of  $k$  binomial populations or the rule  $R_{BC}$  for selecting from  $k-1$  binomial populations that are better than an unknown control.

$$p^* = .90$$

$n \backslash k$	2	5	10	15	20	25	30	40	50
5	2	3	3	3	4	4	4	4	4
10	3	4	5	5	5	5	5	6	6
15	4	5	6	6	6	7	7	7	7
20	4	6	7	7	7	8	8	8	8
25	5	6	7	8	8	8	9	9	9
30	5	7	8	9	9	9	9	10	10
35	5	8	9	9	10	10	10	11	11
40	6	8	9	10	10	11	11	11	12
45	6	9	10	11	11	11	12	12	12
50	6	9	11	11	12	12	12	13	13
60	7	10	12	12	13	13	13	14	14
70	8	10	12	13	14	14	14	15	15
80	8	12	13	14	15	15	15	16	16
90	9	12	14	15	16	16	16	17	17
100	9	13	15	16	16	17	17	18	18
250	14	21	24	25	26	27	27	28	29
500	20	29	33	35	37	38	39	40	41

The above values of  $d$  were computed by using the normal approximation as given in equation (5). Exact calculations based on (4) were used for  $n \leq 10$ .

Table 2. Values of  $d$  for implementing the rule  $R_B$  for selecting the best of  $k$  binomial populations or the rule  $R_{BC}$  for selecting from  $k-1$  binomial populations that are better than an unknown control.

$$P^* = .95$$

n \ k	2	5	10	15	20	25	30	40	50
5	3	3	4	4	4	4	4	4	5
10	4	5	5	6	6	6	6	6	6
15	5	6	7	7	7	7	8	8	8
20	5	7	8	8	8	8	9	9	9
25	6	8	8	9	9	9	10	10	10
30	6	8	9	10	10	10	11	11	11
35	7	9	10	11	11	11	11	12	12
40	7	10	11	11	12	12	12	13	13
45	8	10	11	12	12	13	13	13	14
50	8	11	12	13	13	13	14	14	14
60	9	12	13	14	14	15	15	15	16
70	10	13	14	15	16	16	16	17	17
80	10	14	15	16	17	17	17	18	18
90	11	14	16	17	18	18	18	19	19
100	12	15	17	18	19	19	19	20	20
250	18	24	27	28	29	30	31	32	32
500	25	34	38	40	42	43	43	45	46

The above values of  $d$  were computed by using the normal approximation as given in equation (5). Exact calculations based on (4) were carried out for  $n \leq 10$ .

Table 3. PCS (top line), P(NCS) (middle line) and E(S) (bottom line)

Slippage Configuration:  $p = 0.50$   $\delta = 0.10$

n	k= 3				5				10				15			
	d=2	3	4	5	2	3	4	5	2	3	4	5	2	3	4	5
5	0.95	0.99	1.00	1.00	0.92	0.99	1.00	1.00	0.87	0.97	1.00	1.00	0.83	0.96	1.00	1.00
	0.87	0.97	1.00	1.00	0.82	0.95	0.99	1.00	0.74	0.93	0.99	1.00	0.70	0.91	0.99	1.00
	2.69	2.93	2.99	3.00	4.19	4.80	4.98	5.00	7.56	9.33	9.92	10.00	10.60	13.69	14.83	15.00
10	0.90	0.96	0.99	1.00	0.84	0.93	0.98	0.99	0.75	0.89	0.96	0.99	0.69	0.86	0.95	0.99
	0.69	0.84	0.93	0.98	0.62	0.79	0.91	0.97	0.51	0.71	0.87	0.95	0.45	0.67	0.84	0.94
	2.29	2.64	2.85	2.95	3.30	4.10	4.61	4.87	5.33	7.32	8.77	9.56	6.98	10.17	12.68	14.14
15	0.88	0.94	0.98	0.99	0.81	0.90	0.96	0.98	0.70	0.84	0.92	0.97	0.64	0.79	0.90	0.96
	0.58	0.73	0.84	0.92	0.50	0.66	0.80	0.90	0.40	0.57	0.73	0.86	0.34	0.51	0.69	0.83
	2.04	2.39	2.66	2.83	2.81	3.56	4.16	4.57	4.26	5.97	7.51	8.66	5.37	7.99	10.51	12.51
20	0.87	0.93	0.97	0.98	0.80	0.88	0.94	0.97	0.69	0.81	0.89	0.95	0.62	0.76	0.86	0.93
	0.50	0.64	0.75	0.85	0.43	0.57	0.70	0.81	0.33	0.47	0.62	0.76	0.27	0.42	0.57	0.72
	1.88	2.20	2.47	2.68	2.50	3.16	3.76	4.23	3.63	5.07	6.51	7.76	4.47	6.62	8.88	10.96
25	0.87	0.92	0.96	0.98	0.79	0.87	0.93	0.96	0.68	0.79	0.88	0.93	0.62	0.74	0.84	0.91
	0.44	0.56	0.68	0.78	0.37	0.50	0.62	0.74	0.28	0.41	0.54	0.67	0.23	0.35	0.49	0.62
	1.75	2.05	2.32	2.54	2.28	2.87	3.43	3.92	3.21	4.45	5.74	6.96	3.88	5.69	7.68	9.66
30	0.87	0.92	0.95	0.98	0.79	0.87	0.92	0.96	0.68	0.79	0.86	0.92	0.62	0.73	0.83	0.90
	0.40	0.50	0.61	0.71	0.33	0.44	0.56	0.67	0.25	0.36	0.48	0.60	0.20	0.31	0.42	0.55
	1.66	1.93	2.18	2.40	2.12	2.64	3.16	3.63	2.91	3.98	5.14	6.30	3.47	5.01	6.77	8.60
35	0.87	0.92	0.95	0.97	0.80	0.87	0.92	0.95	0.69	0.78	0.86	0.91	0.62	0.73	0.82	0.88
	0.36	0.46	0.56	0.65	0.30	0.40	0.50	0.61	0.22	0.32	0.42	0.54	0.18	0.27	0.37	0.49
	1.59	1.83	2.07	2.28	1.99	2.45	2.93	3.39	2.68	3.62	4.66	5.73	3.16	4.50	6.06	7.73
40	0.88	0.92	0.95	0.97	0.80	0.87	0.91	0.95	0.70	0.78	0.85	0.91	0.63	0.73	0.81	0.88
	0.32	0.41	0.51	0.60	0.27	0.36	0.46	0.56	0.20	0.28	0.38	0.48	0.16	0.24	0.33	0.44
	1.52	1.75	1.97	2.17	1.88	2.30	2.74	3.17	2.49	3.33	4.27	5.26	2.92	4.10	5.49	7.01
45	0.88	0.92	0.95	0.97	0.81	0.87	0.91	0.95	0.71	0.79	0.85	0.90	0.64	0.73	0.81	0.87
	0.30	0.38	0.46	0.55	0.25	0.33	0.42	0.51	0.18	0.26	0.34	0.44	0.15	0.22	0.30	0.39
	1.47	1.68	1.88	2.08	1.79	2.18	2.58	2.98	2.34	3.09	3.94	4.85	2.72	3.77	5.02	6.40
50	0.89	0.92	0.95	0.97	0.82	0.87	0.91	0.94	0.72	0.79	0.85	0.90	0.65	0.74	0.81	0.87
	0.27	0.35	0.43	0.51	0.23	0.30	0.38	0.47	0.17	0.23	0.31	0.40	0.14	0.20	0.27	0.36
	1.43	1.61	1.80	1.99	1.72	2.07	2.44	2.81	2.22	2.89	3.67	4.51	2.56	3.50	4.63	5.89

Table 3 (Continued). PCS (top line), P(NCS) (middle line) and E(S) (bottom line)

Slippage Configuration:  $p = 0.50$   $\delta = 0.10$ 

	k= 3				5				10				15			
	d=2	3	4	5	2	3	4	5	2	3	4	5	2	3	4	5
n																
60	0.90	0.93	0.95	0.97	0.83	0.88	0.92	0.94	0.73	0.80	0.86	0.90	0.67	0.75	0.82	0.87
	0.23	0.29	0.36	0.44	0.19	0.25	0.32	0.40	0.14	0.20	0.26	0.34	0.12	0.17	0.23	0.30
	1.36	1.51	1.68	1.84	1.60	1.89	2.21	2.54	2.01	2.57	3.22	3.93	2.30	3.07	3.99	5.05
70	0.91	0.93	0.95	0.97	0.85	0.89	0.92	0.95	0.76	0.81	0.86	0.90	0.70	0.76	0.82	0.87
	0.20	0.25	0.31	0.38	0.17	0.22	0.28	0.34	0.12	0.17	0.22	0.29	0.10	0.14	0.19	0.25
	1.30	1.44	1.58	1.72	1.51	1.76	2.02	2.31	1.86	2.33	2.88	3.49	2.11	2.75	3.53	4.43
80	0.92	0.94	0.96	0.97	0.86	0.90	0.93	0.95	0.78	0.83	0.87	0.91	0.72	0.78	0.83	0.88
	0.17	0.22	0.27	0.32	0.14	0.19	0.24	0.29	0.11	0.15	0.19	0.25	0.09	0.12	0.17	0.22
	1.26	1.37	1.49	0.62	1.44	1.65	1.88	2.13	1.74	2.14	2.61	3.13	1.95	2.50	3.16	3.93
90	0.93	0.95	0.96	0.97	0.88	0.91	0.93	0.95	0.79	0.84	0.88	0.91	0.74	0.80	0.85	0.88
	0.15	0.19	0.23	0.28	0.12	0.16	0.21	0.26	0.09	0.13	0.17	0.21	0.08	0.11	0.14	0.19
	1.22	1.32	1.42	1.54	1.38	1.56	1.76	1.97	1.64	1.99	2.39	2.84	1.83	2.30	2.87	3.53
100	0.93	0.95	0.96	0.98	0.89	0.92	0.94	0.96	0.81	0.86	0.89	0.92	0.76	0.81	0.86	0.89
	0.13	0.16	0.20	0.25	0.11	0.14	0.18	0.22	0.08	0.11	0.15	0.19	0.07	0.09	0.13	0.16
	1.19	1.28	1.37	1.47	1.33	1.48	1.65	1.85	1.56	1.86	2.21	2.60	1.73	2.14	2.63	3.20
250	0.99	0.99	0.99	0.99	0.98	0.98	0.99	0.99	0.96	0.97	0.97	0.98	0.94	0.95	0.96	0.97
	0.02	0.02	0.03	0.04	0.02	0.02	0.03	0.04	0.02	0.02	0.03	0.03	0.01	0.02	0.02	0.03
	1.03	1.04	1.05	1.07	1.05	1.07	1.10	1.13	1.10	1.14	1.20	1.26	1.13	1.20	1.28	1.37
500	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00
	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
	1.00	1.00	1.00	1.00	1.00	1.00	1.01	1.01	1.01	1.01	1.01	1.02	1.01	1.01	1.02	1.02

For values of  $n \geq 60$ , the values in the above table were computed by using the normal approximations given in (8).

Table 4. PCS (top line), P(NCS) (middle line) and E(S) (bottom line)

Slippage Configuration:  $p = 0.75$   $\delta = 0.05$ 

n	k= 3				5				10				15			
	d=2	3	4	5	2	3	4	5	2	3	4	5	2	3	4	5
5	0.97	1.00	1.00	1.00	0.96	1.00	1.00	1.00	0.95	0.99	1.00	1.00	0.94	0.99	1.00	1.00
	0.94	0.99	1.00	1.00	0.92	0.99	1.00	1.00	0.90	0.99	1.00	1.00	0.90	0.98	1.00	1.00
	2.86	2.98	3.00	3.00	4.65	4.95	5.00	5.00	9.08	9.86	9.99	10.00	13.52	14.78	14.99	15.00
10	0.91	0.97	0.99	1.00	0.87	0.96	0.99	1.00	0.80	0.93	0.98	1.00	0.77	0.92	0.98	1.00
	0.81	0.93	0.98	1.00	0.75	0.90	0.97	0.99	0.68	0.86	0.96	0.99	0.63	0.84	0.95	0.99
	2.54	2.83	2.95	2.99	3.89	4.56	4.87	4.97	6.89	8.69	9.59	9.90	9.63	12.65	14.23	14.82
15	0.87	0.95	0.98	0.99	0.81	0.92	0.97	0.99	0.72	0.87	0.95	0.98	0.67	0.84	0.93	0.98
	0.72	0.85	0.93	0.98	0.64	0.80	0.91	0.97	0.54	0.73	0.87	0.95	0.48	0.69	0.85	0.94
	2.31	2.65	2.85	2.95	3.37	4.13	4.61	4.85	5.57	7.47	8.79	9.52	7.45	10.52	12.79	14.10
20	0.85	0.93	0.97	0.99	0.77	0.88	0.95	0.98	0.67	0.82	0.91	0.96	0.61	0.77	0.89	0.95
	0.65	0.78	0.88	0.94	0.56	0.72	0.84	0.92	0.45	0.63	0.79	0.89	0.40	0.58	0.75	0.87
	2.15	2.49	2.73	2.88	3.02	3.77	4.32	4.68	4.75	6.52	7.98	8.98	6.16	8.93	11.36	13.10
25	0.83	0.91	0.96	0.98	0.75	0.86	0.93	0.97	0.63	0.78	0.88	0.94	0.57	0.73	0.85	0.93
	0.59	0.73	0.83	0.91	0.51	0.66	0.78	0.88	0.40	0.56	0.71	0.83	0.34	0.50	0.66	0.80
	2.02	2.36	2.62	2.79	2.77	3.48	4.06	4.47	4.19	5.80	7.26	8.39	5.31	7.77	10.14	12.07
30	0.82	0.90	0.95	0.97	0.73	0.84	0.91	0.95	0.61	0.75	0.85	0.92	0.55	0.70	0.82	0.90
	0.55	0.68	0.78	0.87	0.46	0.60	0.73	0.83	0.35	0.50	0.64	0.77	0.30	0.44	0.59	0.73
	1.92	2.25	2.51	2.71	2.58	3.25	3.82	4.27	3.79	5.25	6.65	7.84	4.72	6.91	9.13	11.11
35	0.82	0.89	0.94	0.97	0.72	0.82	0.90	0.94	0.60	0.73	0.83	0.90	0.53	0.67	0.79	0.88
	0.51	0.63	0.74	0.83	0.43	0.56	0.68	0.78	0.32	0.45	0.59	0.71	0.27	0.40	0.54	0.67
	1.85	2.16	2.42	2.63	2.43	3.05	3.62	4.08	3.49	4.82	6.15	7.34	4.28	6.24	8.31	10.26
40	0.81	0.88	0.93	0.96	0.72	0.81	0.89	0.93	0.59	0.71	0.81	0.89	0.52	0.66	0.77	0.86
	0.48	0.60	0.70	0.79	0.40	0.52	0.64	0.74	0.30	0.42	0.54	0.67	0.24	0.36	0.49	0.62
	1.78	2.08	2.33	2.55	2.31	2.89	3.44	3.91	3.25	4.47	5.72	6.89	3.94	5.71	7.63	9.51
45	0.81	0.88	0.92	0.96	0.72	0.81	0.88	0.93	0.59	0.70	0.80	0.88	0.52	0.64	0.75	0.84
	0.46	0.56	0.67	0.76	0.37	0.49	0.60	0.70	0.27	0.39	0.51	0.62	0.23	0.33	0.45	0.57
	1.73	2.01	2.26	2.47	2.21	2.76	3.28	3.74	3.06	4.18	5.35	6.49	3.67	5.28	7.07	8.87
50	0.81	0.87	0.92	0.95	0.71	0.80	0.87	0.92	0.58	0.70	0.79	0.86	0.51	0.63	0.74	0.83
	0.43	0.54	0.64	0.73	0.35	0.46	0.57	0.67	0.26	0.36	0.47	0.59	0.21	0.31	0.42	0.53
	1.68	1.94	2.19	2.40	2.13	2.64	3.14	3.60	2.90	3.93	5.04	6.13	3.45	4.93	6.59	8.30

Table 4 (Continued). PCS (top line), P(NCS) (middle line) and E(S) (bottom line)

Slippage Configuration:  $p = 0.75$   $\delta = 0.05$ 

n	k= 3				5				10				15			
	d=2	3	4	5	2	3	4	5	2	3	4	5	2	3	4	5
60	0.81	0.87	0.91	0.94	0.70	0.78	0.85	0.90	0.56	0.66	0.75	0.83	0.48	0.59	0.69	0.78
	0.39	0.48	0.58	0.66	0.31	0.41	0.50	0.60	0.22	0.30	0.40	0.50	0.18	0.25	0.34	0.44
	1.59	1.83	2.06	2.27	1.96	2.41	2.86	3.30	2.55	3.41	4.35	5.34	2.94	4.11	5.48	6.97
70	0.81	0.87	0.91	0.94	0.71	0.78	0.85	0.89	0.57	0.66	0.75	0.82	0.49	0.59	0.68	0.77
	0.36	0.44	0.53	0.61	0.29	0.37	0.46	0.55	0.20	0.28	0.36	0.45	0.16	0.23	0.31	0.39
	1.53	1.75	1.96	2.16	1.86	2.26	2.68	3.08	2.38	3.14	3.99	4.88	2.72	3.75	4.96	6.30
80	0.82	0.87	0.91	0.94	0.71	0.78	0.84	0.89	0.57	0.66	0.74	0.81	0.49	0.59	0.68	0.76
	0.33	0.41	0.49	0.57	0.27	0.34	0.42	0.50	0.19	0.25	0.33	0.41	0.15	0.21	0.28	0.36
	1.48	1.68	1.88	2.07	1.78	2.15	2.52	2.90	2.25	2.93	3.69	4.51	2.55	3.47	4.55	5.75
90	0.82	0.87	0.90	0.93	0.72	0.79	0.84	0.89	0.58	0.67	0.74	0.81	0.50	0.59	0.68	0.75
	0.31	0.38	0.45	0.53	0.25	0.31	0.39	0.47	0.17	0.23	0.30	0.38	0.14	0.19	0.25	0.33
	1.44	1.63	1.81	1.99	1.71	2.05	2.40	2.75	2.14	2.75	3.45	4.20	2.41	3.24	4.21	5.30
100	0.83	0.87	0.91	0.93	0.73	0.79	0.84	0.88	0.59	0.67	0.74	0.80	0.51	0.60	0.68	0.75
	0.29	0.35	0.42	0.49	0.23	0.29	0.36	0.43	0.16	0.22	0.28	0.35	0.13	0.18	0.23	0.30
	1.41	1.58	1.75	1.92	1.66	1.96	2.28	2.62	2.05	2.61	3.24	3.93	2.30	3.05	3.93	4.93
250	0.90	0.92	0.94	0.95	0.84	0.87	0.89	0.91	0.74	0.78	0.82	0.85	0.68	0.72	0.77	0.81
	0.13	0.15	0.18	0.21	0.11	0.13	0.15	0.18	0.08	0.10	0.12	0.15	0.06	0.08	0.10	0.12
	1.16	1.22	1.29	1.37	1.26	1.38	1.51	1.65	1.44	1.65	1.90	2.16	1.56	1.84	2.17	2.55
500	0.97	0.97	0.98	0.98	0.94	0.95	0.96	0.96	0.89	0.91	0.92	0.93	0.85	0.87	0.89	0.91
	0.04	0.05	0.06	0.07	0.04	0.04	0.05	0.06	0.03	0.04	0.04	0.05	0.02	0.03	0.04	0.04
	1.05	1.07	1.09	1.11	1.08	1.12	1.16	1.20	1.15	1.22	1.30	1.39	1.20	1.30	1.41	1.53

For values of  $n \geq 60$ , the values in the above table were computed by using the normal approximations given in (8).

Table 5. PCS (top line), P(NCS) (middle line) and E(S) (bottom line)

Slippage Configuration:  $p = 0.90$   $\delta = 0.03$

n	k= 3				5				10				15			
	d=2	3	4	5	2	3	4	5	2	3	4	5	2	3	4	5
5	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00
	0.99	1.00	1.00	1.00	0.99	1.00	1.00	1.00	0.99	1.00	1.00	1.00	0.99	1.00	1.00	1.00
	2.98	3.00	3.00	3.00	4.96	5.00	5.00	5.00	9.92	10.00	10.00	10.00	14.88	14.99	15.00	15.00
10	0.98	1.00	1.00	1.00	0.98	1.00	1.00	1.00	0.97	1.00	1.00	1.00	0.97	1.00	1.00	1.00
	0.95	0.99	1.00	1.00	0.94	0.99	1.00	1.00	0.93	0.99	1.00	1.00	0.93	0.99	1.00	1.00
	2.88	2.98	3.00	3.00	4.73	4.95	4.99	5.00	9.35	9.88	9.99	10.00	13.99	14.82	14.98	15.00
15	0.96	0.99	1.00	1.00	0.94	0.99	1.00	1.00	0.93	0.98	1.00	1.00	0.92	0.98	1.00	1.00
	0.89	0.97	0.99	1.00	0.86	0.96	0.99	1.00	0.83	0.95	0.99	1.00	0.82	0.95	0.99	1.00
	2.74	2.93	2.98	3.00	4.38	4.82	4.96	4.99	8.39	9.52	9.89	9.98	12.40	14.23	14.82	14.97
20	0.94	0.98	1.00	1.00	0.91	0.98	0.99	1.00	0.87	0.96	0.99	1.00	0.86	0.96	0.99	1.00
	0.83	0.94	0.98	1.00	0.78	0.92	0.97	0.99	0.73	0.89	0.97	0.99	0.70	0.88	0.96	0.99
	2.60	2.86	2.96	2.99	4.05	4.64	4.89	4.97	7.44	8.99	9.68	9.92	10.71	13.28	14.45	14.86
25	0.92	0.98	0.99	1.00	0.88	0.96	0.99	1.00	0.83	0.94	0.98	1.00	0.80	0.93	0.98	1.00
	0.77	0.90	0.96	0.99	0.72	0.87	0.95	0.98	0.65	0.83	0.93	0.98	0.61	0.81	0.92	0.97
	2.47	2.77	2.92	2.98	3.76	4.44	4.79	4.93	6.66	8.41	9.38	9.80	9.35	12.24	13.90	14.64
30	0.91	0.97	0.99	1.00	0.86	0.95	0.98	1.00	0.80	0.92	0.97	0.99	0.76	0.90	0.97	0.99
	0.73	0.86	0.94	0.98	0.66	0.82	0.92	0.97	0.58	0.77	0.89	0.96	0.54	0.74	0.88	0.95
	2.36	2.69	2.87	2.95	3.52	4.24	4.67	4.88	6.04	7.86	9.02	9.62	8.30	11.28	13.26	14.31
35	0.90	0.96	0.99	1.00	0.85	0.93	0.98	0.99	0.77	0.90	0.96	0.99	0.73	0.88	0.95	0.98
	0.68	0.82	0.91	0.96	0.62	0.78	0.89	0.95	0.53	0.72	0.85	0.94	0.48	0.68	0.83	0.92
	2.27	2.61	2.82	2.92	3.31	4.06	4.54	4.80	5.53	7.36	8.65	9.40	7.47	10.42	12.60	13.92
40	0.89	0.95	0.98	0.99	0.83	0.92	0.97	0.99	0.75	0.88	0.95	0.98	0.70	0.85	0.94	0.98
	0.65	0.79	0.89	0.95	0.58	0.74	0.86	0.93	0.49	0.67	0.81	0.91	0.44	0.63	0.79	0.89
	2.19	2.53	2.76	2.89	3.14	3.88	4.40	4.72	5.12	6.90	8.28	9.16	6.81	9.66	11.95	13.47
45	0.88	0.95	0.98	0.99	0.82	0.91	0.96	0.99	0.73	0.86	0.94	0.98	0.68	0.83	0.92	0.97
	0.62	0.76	0.86	0.93	0.54	0.70	0.83	0.91	0.45	0.63	0.77	0.88	0.40	0.58	0.74	0.86
	2.11	2.46	2.70	2.85	2.99	3.72	4.27	4.63	4.77	6.50	7.91	8.89	6.27	8.99	11.32	13.00
50	0.88	0.94	0.97	0.99	0.81	0.90	0.96	0.98	0.72	0.85	0.93	0.97	0.67	0.81	0.91	0.96
	0.59	0.73	0.83	0.91	0.51	0.67	0.80	0.89	0.42	0.59	0.74	0.85	0.37	0.54	0.70	0.83
	2.05	2.39	2.64	2.81	2.86	3.58	4.14	4.53	4.48	6.14	7.57	8.62	5.82	8.41	10.74	12.53



Table 5. (Continued). PCS (top line), P(NCS) (middle line) and E(S) (bottom line)

Slippage Configuration:  $p = 0.90$   $\delta = 0.03$ 

n	k= 3				5				10				15			
	d=2	3	4	5	2	3	4	5	2	3	4	5	2	3	4	5
60	0.86	0.93	0.96	0.98	0.77	0.87	0.93	0.97	0.64	0.78	0.88	0.94	0.56	0.72	0.84	0.92
	0.53	0.66	0.77	0.86	0.44	0.58	0.71	0.82	0.33	0.47	0.62	0.75	0.27	0.41	0.56	0.70
	1.91	2.24	2.51	2.71	2.54	3.20	3.79	4.25	3.61	5.04	6.46	7.69	4.37	6.47	8.69	10.74
70	0.86	0.92	0.96	0.98	0.77	0.86	0.93	0.96	0.64	0.77	0.86	0.93	0.56	0.70	0.82	0.90
	0.49	0.61	0.73	0.82	0.41	0.54	0.66	0.77	0.30	0.43	0.57	0.69	0.25	0.37	0.51	0.64
	1.83	2.14	2.41	2.62	2.39	3.01	3.58	4.06	3.34	4.63	5.96	7.18	3.99	5.86	7.91	0.90
80	0.86	0.92	0.95	0.98	0.77	0.86	0.92	0.96	0.63	0.76	0.85	0.92	0.55	0.69	0.81	0.89
	0.45	0.57	0.68	0.78	0.38	0.50	0.62	0.73	0.28	0.39	0.52	0.65	0.22	0.34	0.46	0.59
	1.76	2.06	2.32	2.53	2.27	2.84	3.39	3.87	3.12	4.30	5.54	6.73	3.70	5.38	7.26	9.16
90	0.86	0.91	0.95	0.97	0.77	0.85	0.91	0.95	0.63	0.75	0.84	0.91	0.55	0.68	0.79	0.88
	0.42	0.53	0.64	0.74	0.35	0.46	0.58	0.69	0.26	0.36	0.48	0.60	0.21	0.31	0.42	0.55
	1.70	1.98	2.23	2.45	2.17	2.70	3.22	3.70	2.94	4.02	5.18	6.32	3.46	4.99	6.73	8.52
100	0.86	0.91	0.95	0.97	0.77	0.85	0.91	0.95	0.64	0.75	0.84	0.90	0.56	0.68	0.79	0.87
	0.40	0.50	0.61	0.70	0.33	0.43	0.54	0.65	0.24	0.34	0.45	0.56	0.19	0.28	0.39	0.51
	1.65	1.91	2.16	2.38	2.08	2.58	3.08	3.54	2.79	3.78	4.87	5.96	3.27	4.66	6.27	7.97
250	0.90	0.93	0.95	0.97	0.84	0.88	0.91	0.94	0.73	0.79	0.84	0.89	0.66	0.73	0.79	0.85
	0.19	0.24	0.29	0.35	0.16	0.21	0.26	0.32	0.12	0.16	0.21	0.26	0.10	0.13	0.18	0.23
	1.28	1.40	1.53	1.67	1.47	1.70	1.94	2.20	1.80	2.22	2.70	3.24	2.02	2.59	3.26	4.04
500	0.96	0.97	0.97	0.98	0.92	0.94	0.95	0.97	0.86	0.89	0.91	0.93	0.81	0.85	0.88	0.91
	0.07	0.09	0.11	0.13	0.06	0.08	0.10	0.12	0.05	0.07	0.08	0.10	0.04	0.06	0.07	0.09
	1.10	1.15	1.19	1.25	1.18	1.26	1.35	1.45	1.32	1.48	1.66	1.87	1.43	1.65	1.90	2.20

For values of  $n \geq 60$ , the values in the above table were computed by using the normal approximations given in (8).

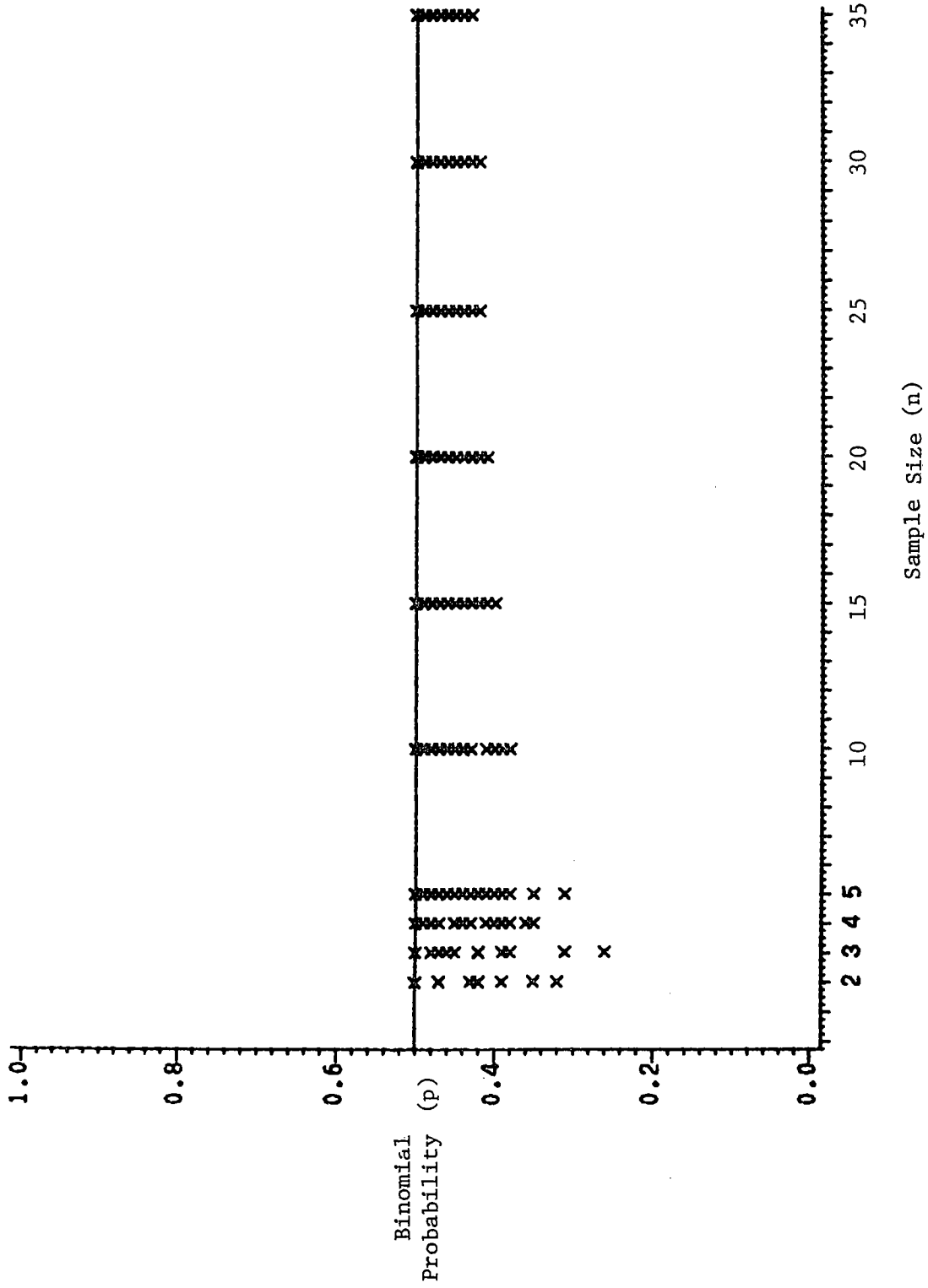


Figure 1. Value of p minimizing the probability of a correct selection:  $k=2,3,5,10,15$ ;  $d=0(1)n$ .

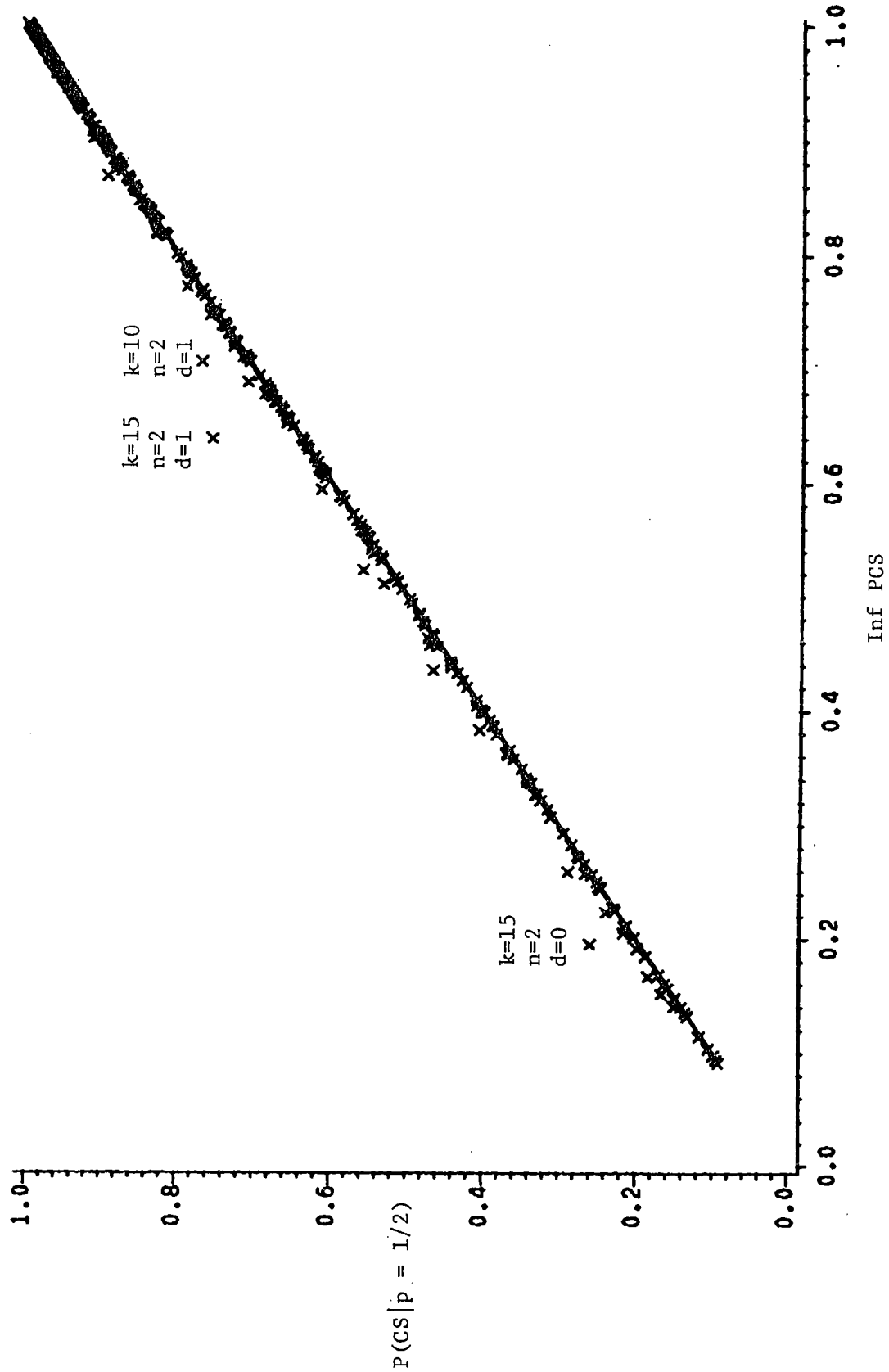


Figure 2. The infimum value of the probability of a correct selection vs. the probability of a correct selection computed at  $p=0.5$  for all populations:  $k=2,3,5,10,15$ ;  $n=2(1)5(5)35$ ;  $d=0(1)n$ .

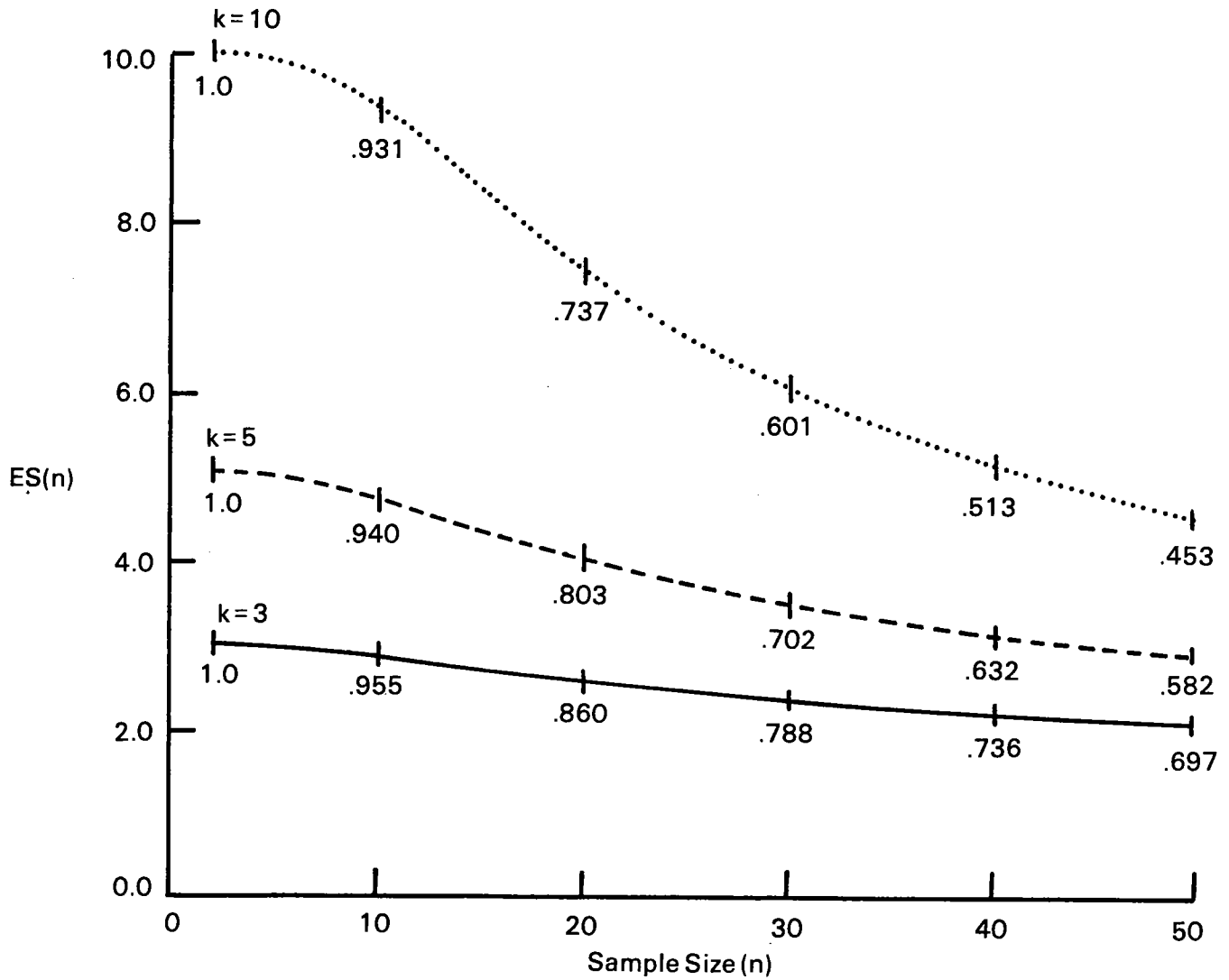


Figure 3. Expected size of selected subset for  $p = .90$ ,  $\delta = .03$ ,  $d = 2$ , and  $k = 3, 5, 10$ . Inserted numbers are probability of a correct selection with  $\delta = 0$  and  $p = .90$ .

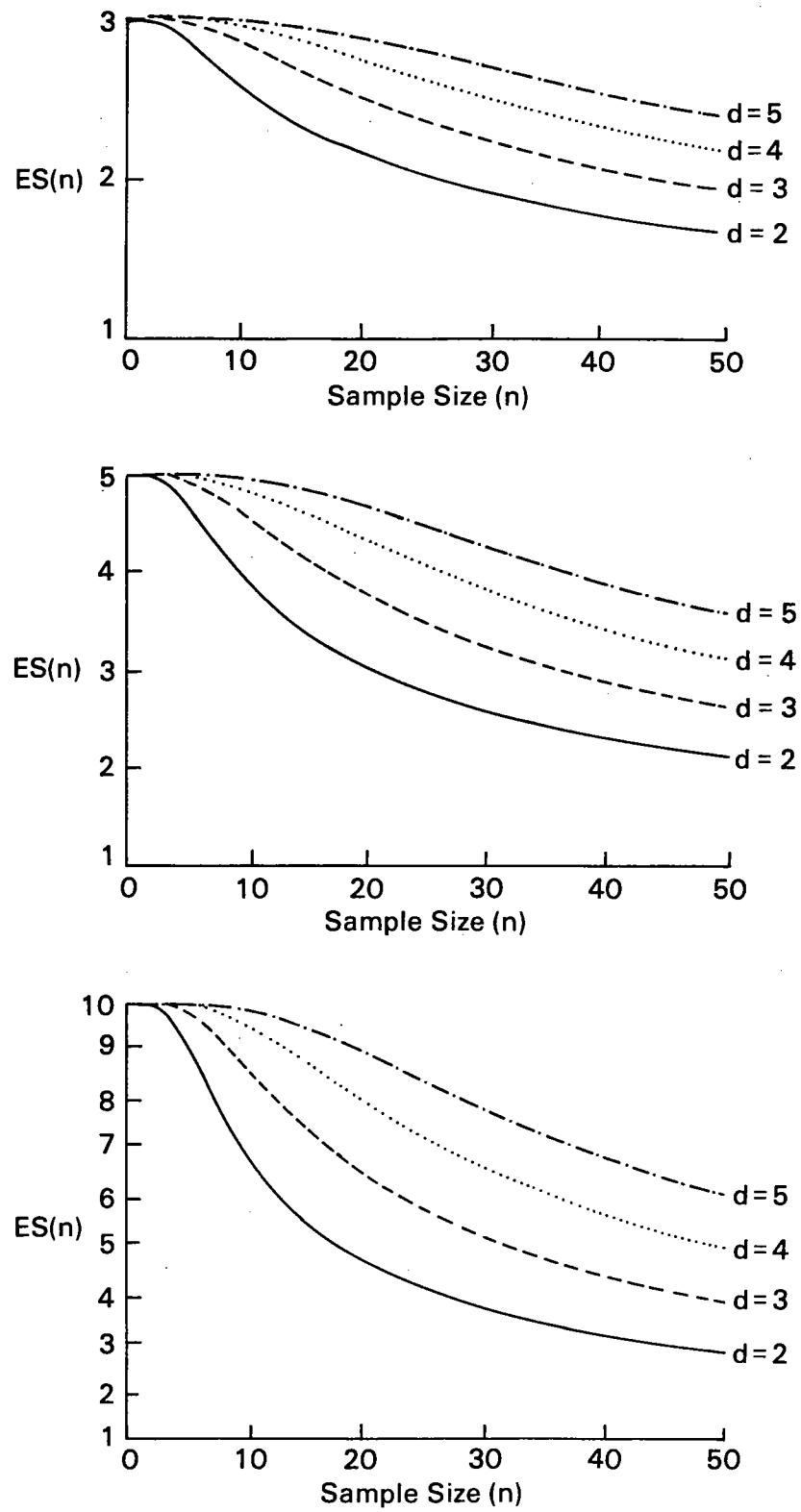


Figure 4. Expected size of selected subset for  $p = .75$ ,  $\delta = .05$  and  $k = 3$  (top),  $k = 5$  (middle) and  $k = 10$  (bottom).

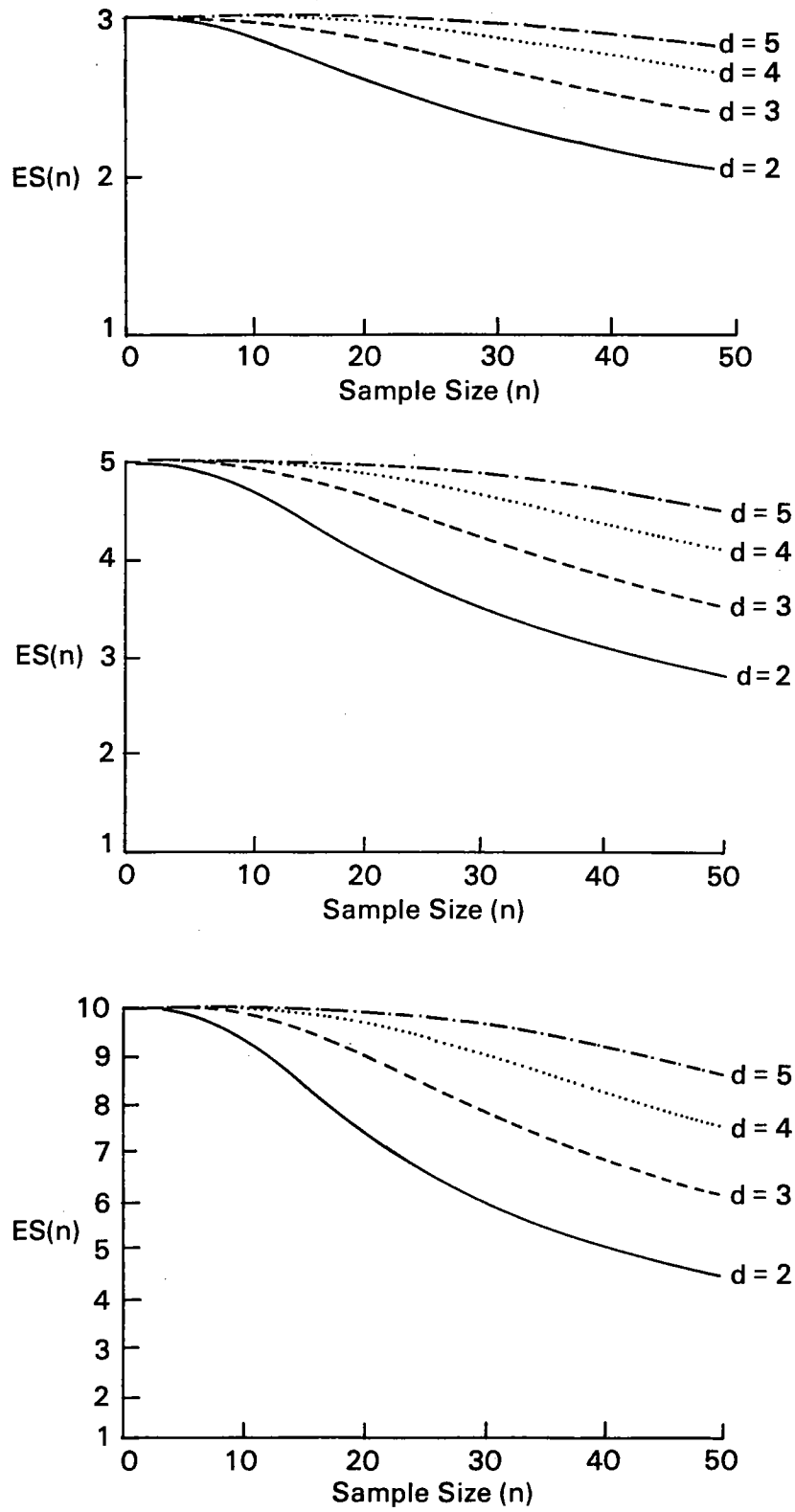


Figure 5. Expected size of selected subset for  $p = .90$ ,  $\delta = .03$  and  $k = 3$  (top),  $k = 5$  (middle, and  $k = 10$  (bottom).

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20. ABSTRACT (Continue on reverse side if necessary and identify by block number) A subset selection procedure $R_B$ for binomial populations is considered for the problem of selecting the best of $k$ vendors whose manufacturing processes have the probabilities $p_1, \dots, p_k$ of turning out an item which conforms to specifications. Let $X_1, \dots, X_k$ denote the number of conforming items from samples of size $n$ from the $k$ processes. Then the rule $R_B$ is of the form: Select $\pi_j$ if and only if $X_j \geq \max_{1 \leq j \leq k} X_j - d$ , where $d$ is a nonnegative integer. The operating characteristics (over)			

(i.e. selection probabilities and expected size of the selected subset) of this rule are related to the underlying  $p_i$ 's, the common sample size  $n$ , and  $d$ . Formulae (both exact and asymptotic) are given for these quantities for slippage as well as equi-spaced parametric configurations. Tables and graphs relating these quantities are presented for three specific slippage configurations. Numerical illustrations are given to show the use of the tables in determining the sample size  $n$  and the constant  $d$  to be used in the rule  $R_B$ . Also, a rule  $R_{BC}$  is mentioned for selecting vendors who are better than (that is, having higher success-probability) a given (control) vendor.

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